# Confidence Interval of Parameters for Gaussian Statistical Structures Z-Criteria's Application

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#### **ABSTRACT**

In this paper is proven 100% confidence interval of parameters for Gaussian statistical structures in Banach space of measures. **Key words:** Gaussian statistical structure, consistent estimators of parameters, Z-criteria, orthogonal structure, strongly separable structure, confidence interval of parameters.

#### Introduction

Recall that a statistical criterion is any measurable mapping from the set all possible samples values to the set of hypothesis. It is said that an error of h-th kind of the  $\delta$  criterion occurs, if the criterion ejects the main hypothesis of  $H_h$ . The following  $\alpha_h(\delta) = \mu_h(\{x: \delta(x) \neq h\})$  is called the probability of an error of the h-th kind for a given criterion  $\delta$ .

The notion and corresponding construction of Z-criteria (same "Generalization criterion of Neiman-Pearson, consistent criterion") for hypothesis testing were introduced and studied by Z. Zerakidze (see [2-13]).

We recall some definitions from the works [1-14].

Let (E, S) be a measurable space. The density of Gaussian law is determined by the equality

$$f(x) = \frac{1}{\sqrt{2\pi\sigma}} e^{-\frac{(x-a)^2}{2\sigma^2}}$$

Let  $\mu$  be the probability measure given on  $([-\infty, +\infty), L[-\infty, +\infty))$  by the formula  $\mu(A) = \int_A f(x) dx, A \in L[-\infty; +\infty)$ , where  $L([-\infty, +\infty))$  is Lebesgue  $\sigma$ -algebra. Let  $\{\mu, i \in I\}$  be Gaussian measures.

Definition 1. An object  $\{E, S, \mu, i \in I\}$  is called an Gaussian statistical structure. Definition 2. An Gaussian statistical structure  $\{E, S, \mu, i \in I\}$  is called orthogonal if  $\mu_i$  and  $\mu_j$  are orthogonal for each  $\forall i \neq j, i \in I, j \in I$ .

Definition 3. An Gaussian statistical structure  $\{E, S, \mu, i \in I\}$  is called weakly separable if there exists a family of S-measurable sets  $\{X, i \in I\}$  such that the relations are fulfilled:

$$(\forall i)(\forall j)(i \in I \& j \in I) \Rightarrow \mu_i(X_j) = \begin{cases} 1, & \text{if } i = j \\ 0, & \text{if } i \neq j \end{cases}$$

Let  $\{\mu_{i}, i \in I\}$  be Gaussian measures defined on the measurable space (E, S). For each  $i \in I$  we denote by  $\overline{\mu}_{i}$  the completion of the measure  $\mu_{i}$ , and by dom  $(\overline{\mu}_{i})$  – the  $\sigma$  – algebra of all  $\mu_{i}$  – measurable subsets of E. We denote  $S_{1} = \bigcap_{i \in I} \text{dom}(\overline{\mu}_{i})$ .

Definition 4. The Gaussian statistical structure  $\{E, S, \mu_l, i \in I\}$  is called strongly separable Gaussian statistical structure if there exists a family of *S*-measurable sets  $\{Z_i, i \in I\}$  such that the relations are fulfilled

1  $\mu_i(Z_i) = 1, \forall i \in I;$ 

2 
$$Z_i \cap Z_j = \emptyset \forall i \neq j; i, j \in I$$

3  $\bigcup_{i\in I} Z_i = E.$ 

Let I be set of parameters and B(I) be  $\sigma$ -algebra if subsets of I which contains all finite subsets of I.

Definition 5. We will say that the Gaussian statistical structure  $\{E, S_1, \overline{\mu}_i, i \in I\}$  admits a consistent estimators of parameters if there exists at least one measurable mapping  $f: (E, S_1) \rightarrow (I, B(I))$ , such that  $\overline{\mu}_i(\{x; f(x) = i\}) = 1, \forall i \in I$ .

Let *H* be set of hypotheses and B(H) be  $\sigma$ -algebra of subsets of *H* which contains all finite subsets of H.

Definition 6. We will say that the Gaussian statistical structure  $\{E, S_1, \overline{\mu}_h, h \in H\}$  admits Z-criterion (same "Generalization Neimana-Pearson, consistent criterion") for hypothesis testing if there exists at least one measurable mapping  $\delta: (E, S_1) \rightarrow (H, B(H))$ , such that

 $\bar{\mu_h}(\{x:\delta(x)=h\})=1, \forall h\in H.$ 

Definition 7. The probability  $\alpha_h(\delta) = \overline{\mu}_h(\{x: \delta(x) \neq h\})$  is called the probability of error of hth kind for the given criterion  $\delta$ .

Theorem 1. The Gaussian statistical structure  $\{E, S_1, \overline{\mu}_h h \in H\}$  admits a Z-criterion (same "Generalization Neimana-Pearson, consistent criterion") for hypothesis testing if and only if this probability of error of kind is equal to zero for the criterion  $\delta$ .

Proof. Necessity. Since the statistical structure  $\{E, S_1, \bar{\mu}_h h \in H\}$  admits a Z-criterion countable Gaussian statistical structure  $\{E, S, \mu_h, h \in H\}$  admits a Z-criterion for hypothesis testing, there exists a measurable mapping  $\delta: (E, S_1) \to (H, B(H), \text{ such that } \bar{\mu}_h(\{x: \delta(x) = h\}) = 1, \forall h \in H$ . Therefore,  $\alpha_h(\delta) = \bar{\mu}_h(\{x: \delta(x) \neq h\}) = 0, \forall h \in H$ .

Sufficiency. Since the probability of any kind is equal to zero, have  $\alpha_h(\delta) = \overline{\mu}_h(\{x:\delta(x) \neq h\}) = 0, \forall h \in H.$ 

On other hand,  $\mu\{x: [(\delta(x) = h) \cup (\delta(x) \neq h)]\} = \overline{\mu}_h(\{x: \delta(x) = h\}) = 1, \forall h \in H.$ 

2. Confidence interval for of parameters Gaussian statistical structures in Banach space of measures

Let  $M^{\sigma}$  be a real linear space of all alternating finite measures on S.

Definition 8. A linear subset  $M_B \subset M^{\sigma}$  is called a Banach space of measures if:

1 The norm on  $M_B$  can be defined so that  $M_B$  it is Banach space with respect to this norm, and the inequality  $\| \mu + \lambda v \| \ge \| \mu \|$  holds for any orthogonal measures  $\mu, v \in M_B$  and real number  $\lambda \neq 0$ ;

2 If  $\mu \in M_B$  and  $|f(x)| \le 1$ , then  $v_f(A) = \int_A f(x)\mu(dx) \in B_B$  and  $||v_f|| \le ||\mu||$ ;

3 If  $v_n \in M_B$ ,  $v_n > 0$ ,  $v_n(E) < \infty$ , n = 1, 2, ... and  $v_n \downarrow 0$ , then for any linear functional  $l^* \in M_B^*$ :  $\lim_{n \to \infty} l^*(v_n) = 0$ , where  $M_B^*$  conjugate to linear space  $M_B$ .

Remark 1. The definition and construction of a Banach space of measures were given by Z. Zerakidze (see [14]).

Definition 8. Let I be a set of indexes and  $M_{B_i}$  is a Banach space for all  $i \in I$ . The Banach space  $M_B = \{X_i\}_{i \in I} : X_i \in M_{B_i}, \forall i \in I, \sum_{i \in I} ||X_i|| \le 0\}$  with the norm  $||X_i||_{i \in I} = \sum_{i \in I} ||X_i||_{M_{B_i}}$  is called the direct sum of Banach space  $M_{B_i}$  and is denoted by  $M_B = \bigoplus M_{B_i}$ .

Remark 2. Obviously, any Banach space of measures is a Banach space the elements of which are alternating measures, but not vice versa. The following theorem was proved in [14].

Theorem 2. Let  $M_B$  be a Banach space of measures, then there exists the funnily of pairwise orthogonal probability measures  $\{\mu_{h_i}, i \in I\}$ , Card  $I = 2^{2^c}$ , such that  $M_B = \bigoplus M_{B_i}(\mu_{h_i})$  is Banach space of elements  $\nu$  of the from

$$v(B) = \int f(x)\mu_{h_i}(dx), B \in S, \int |f(x)|\mu_{h_i}(dx) < +\infty, \text{ with the norm}$$
$$\| v \|_{M_{B_i}(\mu_{B_i})} = \int |f(x)|\mu_{h_i}(dx).$$

We define by  $\mathbf{F} = \mathbf{F}(M_B)$  the set of real function f such that  $\int f(x)\bar{\mu}_h dx$  is defined all  $\bar{\mu}_h = M_B$ . Theorem 3. Let  $M_B = \bigoplus M_{B_i}(\bar{\mu}_h)$ , Card  $H \le c$  be the Banach space of measures, E be a complete separable metric space,  $S_1 = \bigcap_{h \in H} \operatorname{dom}(\bar{\mu}_h)$  is a Borel  $\sigma$ -algebra on E. In order for the Borel orthogonal Gaussian statistical structure  $\{E, S_1, \bar{\mu}_h h \in H\}$ , Card H = c to admit Z-criterion (same "Generalization Neimana; Pearson consistent criterion") for hypothesis testing in the theory (ZFC)&(MA) it is necessary and sufficient the correspondence  $f \leftrightarrow h_f$  defined by the

equality  $\int f(x)\bar{\mu}_h(dx) = l_f(\bar{\mu}_h), \ \bar{\mu}_h \in M_B$  was one-to-one (here  $l_f$  is a linear continuous functional on  $M_{B_f} f \in F(M_B)$ .

Proof. Necessity. The existence of Z-criterion for hypothesis testing  $\delta: (E, S_1) \to (H, B(H))$ , implies that  $\overline{\mu}_h(\{x: \delta(x) = h\}) = 1, \forall h \in H$ . Setting  $X_h = (\{x: \delta(x) = h\}) = 1, \forall h \in H$  we get:

1  $\bar{\mu}_h(X_h) = 1, \forall h \in H;$ 

2  $X_{h'} \cap X_{h''} = \emptyset$  for all different h' and h'' from H;

3  $\bigcup_{n \in H} X_h = \{x: \delta(x) \in H\} = E.$ 

Therefore the Gaussian statistical structure  $\{E, S_1, \overline{\mu}_h h \in H\}$  is strongly separable, hence, there exists

 $S_1$  - measurable sets  $\{X_h, h \in H\}$  such that  $\bar{\mu}_h(X_{h'}) = \begin{cases} 1, & \text{if } h = h'\\ 0, & \text{if } h \neq h' \end{cases}$ 

We put the linear continuous functional  $l_{c_h}$  into correspondence to function by the formula  $\int l_{c_h}(x)\bar{\mu}_h(dx) = l_{c_h}(\bar{\mu}_h) = \|\bar{\mu}_h\|_{M_B(\bar{\mu}_h)}.$ 

Let  $l_{X_h}$  be a linear continuous functional that correspondence to the function  $\bar{f_1}(x) = f_1(x)I_{X_h}(x)$ . Then for any  $\bar{\mu}_{h_1} \in M_B(\bar{\mu}_h)$  we have

$$\int \bar{f_1}(x)\bar{\mu}_{h_1}(dx) = \int f_1(x)f(x)I_{X_h}(x)\bar{\mu}_h(dx) = l_{\bar{f_1}}(\bar{\mu}_{h_1}) = \|\bar{\mu}_{h_1}\|_{M_B(\bar{\mu}_h)}.$$

Let  $\sum$  be the set of extensions of a functional that satisfy the condition  $l_f \leq p(x)$  in those subspace where they are defined. Lets introduce a partial ordering into, assuming  $l_{f_1} < l_{f_1}$  if  $f_2$  is defined on a large set than  $l_{f_1}$  and  $l_{f_1} = l_{f_2}$  where both of them are defined.

Let  $\{l_{fh}\}_{h\in H}$  be a linear ordered subsid in  $\sum M_B(\bar{\mu}_h)$  the subspace on which  $l_{fh}$  is defined. We define  $l_f \in \bigcup M_B(\bar{\mu}_h)$  setting  $l_f(\mu) = l_{fh}(\mu)$  if  $\mu \in M_B(\bar{\mu}_h)$ . It is obvious that  $l_{fh} < l_f$ . Since any lineally ordered subset in  $\sum$  has an upper bound due to the Chorn lemma  $\sum$  contains the maximal element  $\lambda$  defined on some set X' satisfying the condition  $\lambda \leq p(x)$  for  $x \in X'$ . But X' must coincide with the entire space  $M_B$  because otherwise we could extended  $\lambda$  to a wider space by adding as above one more dimension. This contradicts the maximality of  $\lambda$  and, hence  $X' = M_B$ . Therefore, the extension of the functional is defined everywhere.

Let  $l_f$  be a linear functional that corresponds to the function  $f(x) = \sum g_h(x) I_{X_h}(x) \in F(M_B)$ . Then we have  $\int f(x) \mu(dx) = ||\mu|| = \sum ||\bar{\mu}_h||_{M_B(\bar{\mu}_h)}, M_B(\bar{\mu}_h)$  where

$$\mu(B) = \sum \int g_h(x) \bar{\mu}_h(dx), B \in S_1$$

Sufficiency. If for each  $f \in F(M_B)$  the integral  $\int f(x)\bar{\mu}_h(dx), \forall \bar{\mu}_h \in M_B$ , is defined them there exist a countable subsets  $I_f$  in H for which  $\int f(x)\bar{\mu}_h(dx) = 0$ , if  $h\bar{\epsilon}I_f$ ,  $\sum \int |f(x)|\bar{\mu}_h(dx) < \infty$  and for any countable subset  $\bar{I} \subset H$  and for the measure

$$v(c) = \int_{h \in i} \int_c g_h(x) \bar{\mu}_h(dx) \text{ we have } \int_E f(x) v(dx) = \sum_{h \in I_f \cap \overline{I}} \int_E f(x) g_h(x) \bar{\mu}_h(dx).$$

Let the correspondence  $f \to l_f$  be calefied the equality  $\int_E f(x)\bar{\mu}_h(dx) = l_f(\bar{\mu}_h)$ , then for  $\bar{\mu}_{h_1}$ ,  $\bar{\mu}_{h_2} \in M_B(\bar{\mu}_h)$  we have  $\int_E f_{h_1}(x)\bar{\mu}_{h_2}(dx) = l_{fh_1}(\bar{\mu}_{h_2}) = \int_E f_1(x)f_2(x)\bar{\mu}_{h_1}(dx) = \int_E f_{h_1}(x)f_{h_2}(x)\bar{\mu}_{h_1}(dx).$ 

Therefore  $f_{h_1}(x) = f_1(x)$  almost everywhere with respect to the measure  $\bar{\mu}_{h_1}$ . Let  $f_{\bar{\mu}_{h_1}}(x) > 0$ almost everywhere with respect to  $\bar{\mu}_{h_i}$  and  $\int_E f_{\bar{\mu}_h}(x)\bar{\mu}_h(dx) < \infty$ . If we denote now  $\bar{\mu}_h(c) = \int_c f_{\bar{\mu}_h}(x)\bar{\mu}_h(dx)$ , the we obtain  $\int_E f_{\bar{\mu}_h}(x)\bar{\mu}_{h'}(dx) = l_{f\bar{\mu}_h}(\bar{\mu}_{h'}) = 0$ ,  $\forall h \neq h' \forall \bar{\mu}_h \in M_B(\bar{\mu}_h)$ .

Denote by  $C_h = \{x: f_{\bar{\mu}h}(x) > 0\}$ . Then  $\mu_{h'}(C_h) = 0 \forall h \neq h'$ . Therefore, there exist  $S_1$  – measurable sets  $(h \in H)$  such that that  $\mu_h(X_{h'}) = \begin{cases} 1, & \text{if } h = h' \\ 0, & \text{if } h \neq h' \end{cases}$  and hence the Gaussian statistical structure

 $\{E, S_1, \overline{\mu}_h h \in H, \text{card} H = c\}$  is weakly separable. We represent as an inductive sequence  $\{\overline{\mu}_h < w_1\}$  where  $w_1$  denotes the first ordinal number of the power of the set H.

We define  $w_1$  sequence  $Z_h$  of parts of the E such that the following relations hold: 1)  $Z_i$  is Borel subset of E,  $\forall h < w_1$ ; 2)  $Z_h \subset X_h, \forall h < w_1$ ; 3)  $Z_h \cap Z_{h'} = \emptyset$  for all  $h' < w_1, h = h'$ ; 4)  $\bar{\mu}_h(Z_h) = 1, \forall h < w_1$ .

Suppose that  $Z_{h_0} = X_{h'_0}$ . Suppose that the partial sequence  $\{Z_{h'}\}_{h' < h}$  is already defined for  $h < w_1$ . It is clear that  $\mu^*(\bigcup_{h' < h} Z_{h'}) = 0$ . Thus there exists a Borel subset  $y_h$  of the space E such that the following relations are valid  $\bigcup_{h' < n} y_h$  and  $\mu^*(y_h) = 0$ . Assuming that  $Z_h = X_h y_h$ , we construct the  $w_1$  sequence  $\{Z_h\}_{h < w_1}$  of disjunctive measurable subsets of the space E. Therefore,  $\mu_h(Z_h) = 1, \forall h < w_1$  and the Gaussian statistical structure  $\{E, S_1, \overline{\mu}_h h \in H, \text{card} H = c\}$  is strongly separable because that exists a family of elements of the  $\sigma$ -algebra  $S_1 = \bigcap_{h \in H} \text{dom}(\overline{\mu}_h)$  such that 1)  $\overline{\mu}_h(Z_h) = 1, \forall h \in H; 2)Z_{h'} \cap Z_h = \emptyset, \forall h' \neq h; 3) \bigcup_{h \in H} Z_h = E.$ 

For  $\mathbf{x} \in E$ , we put  $\delta(\mathbf{x}) = h$ , where *h* is the unique hypothesis from the set H for which  $\mathbf{x} \in Z_h$ . The existence of such a unique hipotez from H can be proved using conditions 2), 3).

Let now  $y \in B(H)$ . Then  $\{x: \delta(x) \in y\} = U_{h \in H}Z_h$ .

It  $h_0 \in H$ , then  $\{x: \delta(x) \in y\} = U_{h \in H} Z_h = Z_{h_0} \cup (U_{h \in H} Z_h)$ . On the other hand the validity of the condition  $\bigcup_{h \in H} Z_h \subseteq E - Z_{h_0}$  implies that  $\bar{\mu}_{h_0} (\bigcup_{h \in y - h_0} Z_h) = 0$ . The last equality yields  $\bigcup_{h \in y - h_0} Z_h \in \text{dom}(\bar{\mu}_{h_0})$ . Since dom  $(\bar{\mu}_{h_0})$  is a  $\sigma$ -algebra, we deduce that  $\{x: \delta(x) \in y\} \in \text{dom}(\bar{\mu}_h)$ .

If  $h_0 \notin y$ , then  $\{x: \delta(x) \in y\} = U_{h \in H} Z_h \le (E - Z_{h_0})$  and we conclude that  $\bar{\mu}_{h_0}(\{x: \delta(x) \in y\}) = 0$ . The last relation implies that  $\{x: \delta(x) \in y\} \in \text{dom}(\bar{\mu}_{h_0})$ .

We have shown that the map  $\delta: (E, S_1) \to (H, B(H))$  is a measurable map. Since B(H) contains all singletons of H we as certain that  $\bar{\mu}_h(\{x:\delta(x)=h\}) = \bar{\mu}_h(Z_h) = 1, \forall h \in H$ .

The following Theorem is proven to Theorem 2.

Theorem 3. Let  $M_B = \bigoplus M_B(\bar{\mu}_i)$ , Card  $I \leq c$  be the Banach space of measures, E be a complete metric space,  $S_1 = \bigcap_{i \in I} \operatorname{dom}(\bar{\mu}_i)$  is a Borel  $\sigma$ -algebra on E. In order for the Borel orthogonal Gaussian statistical structure  $\{E, S_1, \bar{\mu}_i, i \in I\}$ , Card  $I \leq c$  to admit consistent estimator of parameters it is necessary and sufficient that correspondence  $f \leftrightarrow l_f$  defined by the equality  $\int f(x)\bar{\mu}_i(dx) = l_f(\bar{\mu}_i), \bar{\mu}_i \in M_B$  was oneto-one (have  $l_f$  is a linear continuous functional on  $M_B, f \in F(M_B)$ .

The following Theorems 1,2,3,4 follows that exponentials Gaussian structures existence consistent estimator of parameters Z-criterion for hypothesis testing and 100% confidence interval of parameters.

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# პარამეტრების ნდობის ინტერვალი გაუსის სტატისტიკური სტრუქტურებისათვის Z - კრიტერიუმის გამოყენებით

### ზ. ზერაკიძე, ჯ. ქირია, თ. ქირია

### რეზიუმე

ნაშრომში განმარტებულია გაუსის ორთოგონალური, სუსტად განცალებადი, განცალებადი და მლიერად განცალებადი სტატისტიკური სტრუქტურები. ასევე განმარტებულია პარამეტრების მალდებული შეფასებები და პარამეტრების ჰიპოთეზათა შემოწმების სტრუქტურებისათვის Z კრიტერიუმი (იგივეა, რაც "განზოგადოებული ნეიმან-პირსონის კრიტერიუმი", "მალდებული კრიტერიუმი"). აგებულია გაუსის ალბათობების ზომების მიხედვით ბანახის ზომათა სივრცე და დამტკიცებულია ამ სივრცეში აუცილებელი და საკმარისი პირობები პარამეტრების ძალდებული შეფასებების და Z - კრიტერიუმის არსებობის შესახებ.

აგებულია გაუსის სტატისტიკური სტრუქტურების პარამეტრებისათვის 100%-იანი ნდობის ინტერვალი.

საკვანმო სიტყვები: გაუსის სტატისტიკური სტრუქტურა, თანმიმდევრული პარამეტრების შეფასებები, Z-ტესტი, ორთოგონალური სტრუქტურა, მკაცრად განცალკევებული სტრუქტურა, პარამეტრების ნდობის ინტერვალი.

## Доверительный интервал параметров для статистических структур Гаусса с использованием Z-критерия

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### Аннотация

В статье объясняются ортогональные, слабо разделимые, разделимые и сильно разделимые статистические структуры Гаусса. Также даются пояснения о вынужденных оценках параметров и Z-критерии для проверки гипотез о параметрах статистических структур (аналогичен «обобщённому критерию Неймана-Пирсона», «вынужденному критерию»). На основе вероятностных мер Гаусса построено пространство размерностей выборки и доказаны необходимые и достаточные условия существования вынужденных оценок параметров и Z-критерия в этом пространстве.

Для параметров статистических структур Гаусса построен 100%-й доверительный интервал.

Ключевые слова: гауссовская статистическая структура, состоятельные оценки параметров, Z-критерий, ортогональная структура, сильно разделимая структура, доверительный интервал параметров.