

e-ISSN 2667-9973

p-ISSN 1512-1127

**საქართველოს გეოფიზიკური საზოგადოების
ჟურნალი**

**მყარი დედამიწის, ატმოსფეროს, ოკეანისა და კოსმოსური პლაზმის
ფიზიკა**

ტომი 24, № 1

**JOURNAL
OF THE GEORGIAN GEOPHYSICAL SOCIETY**

Physics of Solid Earth, Atmosphere, Ocean and Space Plasma

Vol. 24, № 1

Tbilisi

2021

e-ISSN 2667-9973
p-ISSN 1512-1127

**საქართველოს გეოფიზიკური საზოგადოების
ჟურნალი**

**მყარი დედამიწის, ატმოსფეროს, ოკეანისა და კოსმოსური პლაზმის
ფიზიკა**

ტომი 24 , № 1

**JOURNAL
OF THE GEORGIAN GEOPHYSICAL SOCIETY**

Physics of Solid Earth, Atmosphere, Ocean and Space Plasma

Vol. 24 , № 1

**Tbilisi
2021**

საქართველოს გეოფიზიკური საზოგადოების ჟურნალი
მყარი დედამიწის, ატმოსფეროს, ოკეანისა და კოსმოსური პლაზმის ფიზიკა
მთავარი რედაქტორი: თ. ჭელიძე

სარედაქციო კოლეგია

ა. ამირანაშვილი (მდივანი), თ. ბიბილაშვილი (აშშ), ე. ბოლოპოულოსი (საბერძნეთი), გ. ჩაგელიშვილი, თ. ჭელიძე, ლ. დარახველიძე, დ. დემეტრაშვილი, კ. ევტაქსიასი (საბერძნეთი), ვ. ერემეევი (უკრაინა), ნ. ლლონტი, ა. გოგიჩაიშვილი (მექსიკა), ი. გეგენი (საფრანგეთი), თ. გვენცაძე, ზ. კერესელიძე, ო. ხარშილაძე, ზ. ხვედელიძე, ჯ. ქირია (მთ. რედაქტორის მოადგილე), თ. ქირია, გ. კოროტაევი (უკრაინა), თ. მაჭარაშვილი, გ. მეტრეველი, ი. მურუსიძე, თ. ოგუზი (თურქეთი), ვ. სტაროსტენკო (უკრაინა), რ. ტამსალუ (ესტონეთი), კ. თავართქილაძე, ნ. ვარამაშვილი, ვ. ზალესნი (რუსეთი), ი. ჩშაუ (გერმანია).

ჟურნალის შინაარსი:

ჟურნალი მოიცავს მყარი დედამიწის, ატმოსფეროს, ოკეანისა და კოსმოსური პლაზმის ფიზიკის ყველა მიმართულებას. ჟურნალში ქვეყნდება: კვლევითი წერილები, მიმოხილვები, მოკლე ინფორმაციები, დისკუსიები, წიგნების მიმოხილვები, განცხადებები, კონფერენციების მოხსენებები.

JOURNAL OF THE GEORGIAN GEOPHYSICAL SOCIETY

Physics of Solid Earth, Atmosphere, Ocean and Space Plasma

Editor-in-chief: T. Chelidze

Editorial board:

A. Amiranashvili (secretary), T. Bibilashvili (USA), E. Bolopoulos (Greece), G. Chagelishvili, T. Chelidze, L. Darakhvelidze, D. Demetrashvili, K. Eftaxias (Greece), V. N. Eremeev (Ukraine), N. Ghlonti, A. Gogichaishvili (Mexico), Y. Gueguen (France), T. Gventsadze, Z. Kereselidze, O. Kharshiladze, Z. Khvedelidze, J. Kiria (Vice-Editor), T. Kiria, G. K. Korotaev (Ukraine), T. Matcharashvili, G. Metreveli, I. Murusidze, T. Oguz (Turkey), V. Starostenko (Ukraine), R. Tamsalu (Estonia), K. Tavartkiladze, N. Varamashvili, V. B. Zalesny (Russia), J. Zschau (Germany).

Scope of the Journal:

The Journal is devoted to all branches of the Physics of Solid Earth, Atmosphere, Ocean and Space Plasma. Types of contributions are: research papers, reviews, short communications, discussions, book reviews, announcements, conference reports.

ЖУРНАЛ ГРУЗИНСКОГО ГЕОФИЗИЧЕСКОГО ОБЩЕСТВА

Физика Твердой Земли, Атмосферы, Океана и Космической Плазмы

Главный редактор: Т. Челидзе

Редакционная коллегия:

А. Амиранашвили (секретарь), Т. Бибилашвили (США), Е. Болополоус (Греция), Г. Чагелишвили, Т.Л. Челидзе, Л. Дарахвелидзе, Д. Деметрашвили, К. Эфтаксиас (Греция), В. Н. Еремеев (Украина), Н. Глонти, А. Гогичайшвили (Мексика), И. Геген (Франция), Т. Гвенцадзе, З. Кереселидзе, О. Харшиладзе, З. Хведелидзе, Дж. Кирия (зам. гл. редактора), Т. Кирия, Г. К. Коротаев (Украина), Т. Мачарашвили, Г. Метревели, И. Мурусидзе, Т. Огуз (Турция), В. Старостенко (Украина), Р. Тамсалу (Эстония), К. Таварткиладзе, Н. Варамашвили, В. Б. Залесный (Россия), И. Чшау (Германия).

Содержание журнала:

Журнал Грузинского геофизического общества охватывает все направления физики твердой Земли, Атмосферы, Океана и Космической Плазмы. В журнале публикуются научные статьи, обзоры, краткие информации, дискуссии, обзоры книг, объявления, доклады конференций.

მისამართი:

საქართველო, 0160, თბილისი, ალექსიძის ქ. 1, მ. ნოდიას სახ. გეოფიზიკის ინსტიტუტი
ტელ.: 233-28-67; ფაქსი; (995 32 2332867); ელ. ფოტა: tamaz.chelidze@gmail.com;
avtandilamiranashvili@gmail.com;
geophysics.journal@tsu.ge

გამოქვეყნების განრიგი და ხელმოწერა:

გამოიცემა წელიწადში ორჯერ. მყარი ვერსიის წლიური ხელმოწერის ფასია: უცხოელი ხელმომწერისათვის - 30 დოლარი, საქართველოში - 10 ლარი. ხელმოწერის მოთხოვნა უნდა გაიგზავნოს რედაქციის მისამართით. შესაძლებელია უფასო ონლაინ წვდომა:

<http://openjournals.gela.org.ge/index.php/GGS>

ჟურნალი ინდექსირებულია Google Scholar-ში:

<https://scholar.google.com/citations?hl=en&user=pdG-bMAAAAAAJ>

Address:

M. Nodia Institute of Geophysics, 1 Alexidze Str., 0160 Tbilisi, Georgia
Tel.: 233-28-67; Fax: (99532) 2332867; e-mail: tamaz.chelidze@gmail.com;
avtandilamiranashvili@gmail.com;
geophysics.journal@tsu.ge

Publication schedule and subscription information:

The journal is issued twice a year. The subscription price for print version is 30 \$ in year. Subscription orders should be sent to editor's address. Free online access is possible:

<http://openjournals.gela.org.ge/index.php/GGS>

The journal is indexed in the Google Scholar:

<https://scholar.google.com/citations?hl=en&user=pdG-bMAAAAAAJ>

Адрес:

Грузия, 0160, Тбилиси, ул. Алексидзе, 1. Институт геофизики им. М. З. Нодиа
Тел: 233-28-67; 294-35-91; Fax: (99532)2332867; e-mail: tamaz.chelidze@gmail.com;
avtandilamiranashvili@gmail.com;
geophysics.journal@tsu.ge

Порядок издания и условия подписки:

Журнал издается дважды в год. Годовая подписная цена для печатной версии - 30 долларов США. Заявка о подписке высылается в адрес редакции. Имеется бесплатный онлайн доступ

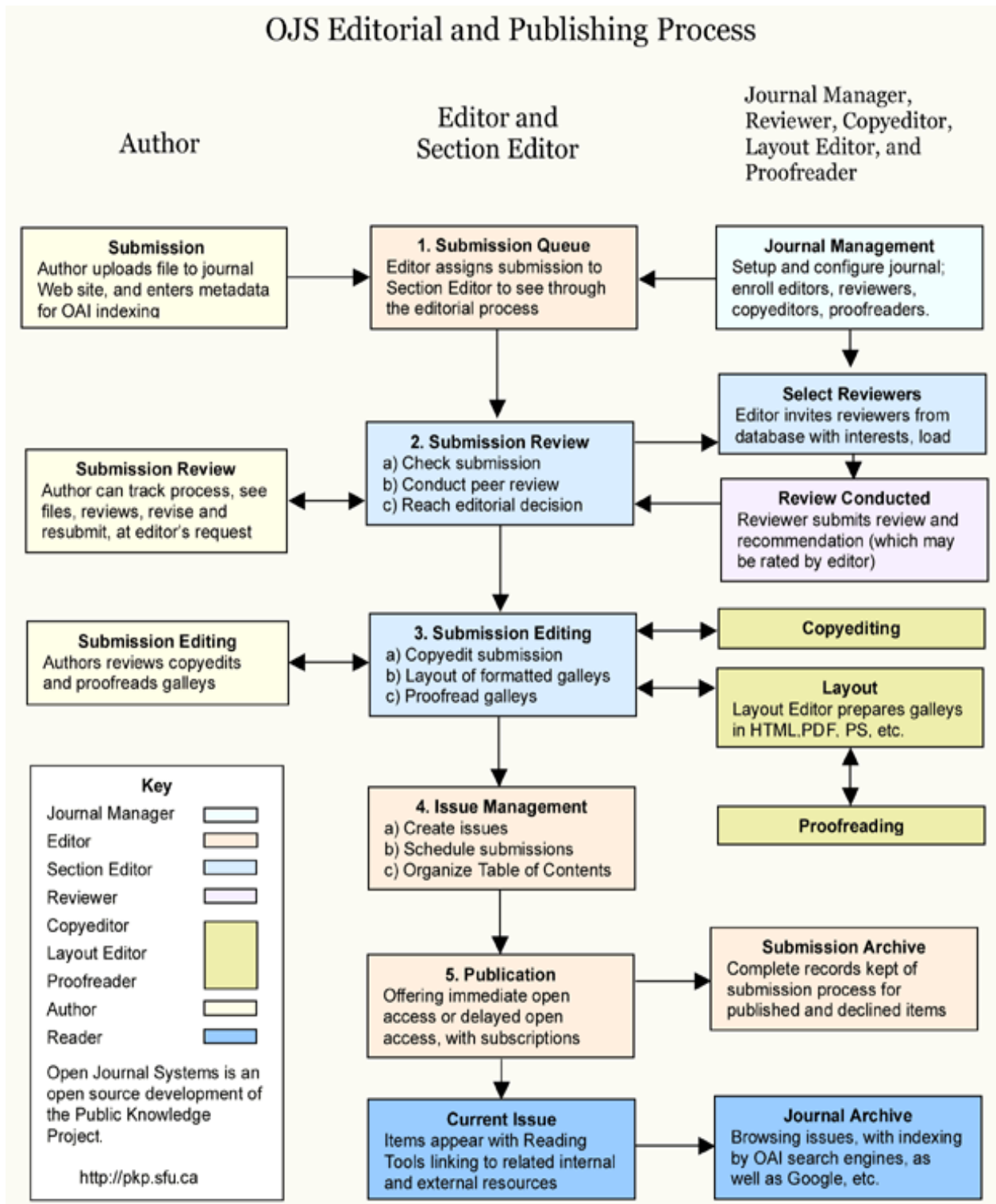
<http://openjournals.gela.org.ge/index.php/GGS>

Журнал индексируется в Google Scholar:

<https://scholar.google.com/citations?hl=en&user=pdG-bMAAAAAAJ>

This journal uses Open Journal Systems 2.4.8.3, which is open source journal management and publishing software developed, supported, and freely distributed by the Public Knowledge Project under the GNU General Public License.

(<http://openjournals.gela.org.ge/index.php?journal=GGS&page=about&op=aboutThisPublishingSystem>)



Magnetic Declination Statistics (Dusheti 1935-1989) and Deep Self-Learning Model

Tengiz V. Kiria, Abesalom A. Esakia, Manana M. Nikolaishvili,
Elene D. Lomadze

Mikheil Nodia Institute of Geophysics of Ivane Javakhishvili Tbilisi State University, Tbilisi, Georgia

ABSTARCT

Nowadays, there are many new instruments available for studying the parameters of the Earth's magnetic field with higher precision and more discretization. Moreover, data processing techniques have been developing on strong mathematical basis. The paper presents a rather long-term (1935 - 1950, total 19332 records) data of Dusheti Observatory on the statistics of magnetic declination (1) and considers the possibilities of so called machine learning (ML), a widespread method nowadays. It gives hypotheses to prove certain hidden regularities and periodicity of some geomagnetic parameters and determines so called "storages" of high statistical reliability, which are the etalon samples we use to build attribution function by use of so called Adam Deep Learning network (2).

Key words: Earth's magnetic field, Adam Deep Learning network.

Preamble

It is required to carry out certain statistical works in order to study a kind of long-term statistics as described in this work. We decided it was necessary to build standard variation series for magnetic declinations on the basis of months and years. Namely, it is important to make time period observations on variation series, also on separate series for 2σ and 3σ (3) parts. These two series turn out to be interesting in the annual point of view. It is highly essential to identify the density distribution of the statistical data in these clusters and whether they show any distribution regularities or behavior in dynamics.

Obviously, there is always an objectivity degree problem (equipment defects, etc.) in data. Consequently, it is required to exclude subjective statistics and there are numerous techniques for that. Nowadays, softwares with filtration tasks have been highly developed among computer technologies. Our data have undergone strict filtration processes. Of course, main informative anomalies have been preserved that is definitely significant.

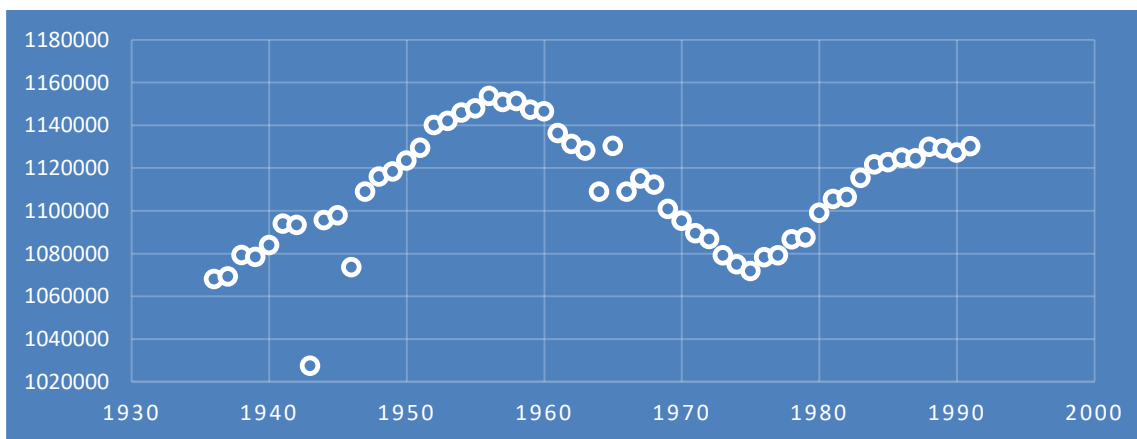


Fig. 1. Dynamics of average annual magnetic declination values (1936 - 1991).

The average annual values in *Figure 1* have periodic and regular structure. There is an exception of the Dusheti Observatory data of 1941-1942 and 1964-1966 in the average annual value point of view. These two episodes are conceivable, though, due to quite objective reasons, they might be a result of strong influence of artificial anomaly nearby the equipment location. We may conclude that, according to the figure, the average annual values have approximately 40-year cycles. Additionally, so called wave behavior has been revealed.

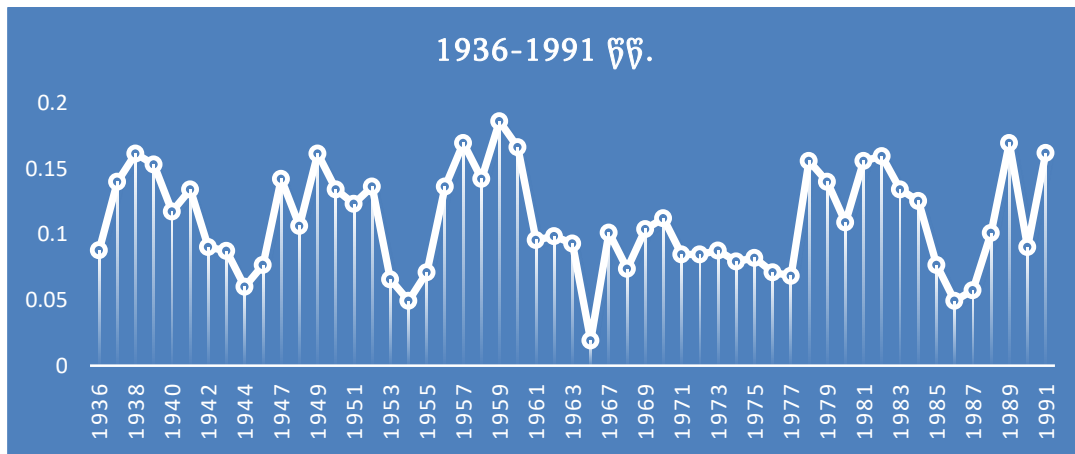


Fig. 2. Quantitative time period analysis of 2σ order anomalies of the magnetic field declination during 1936-1991.

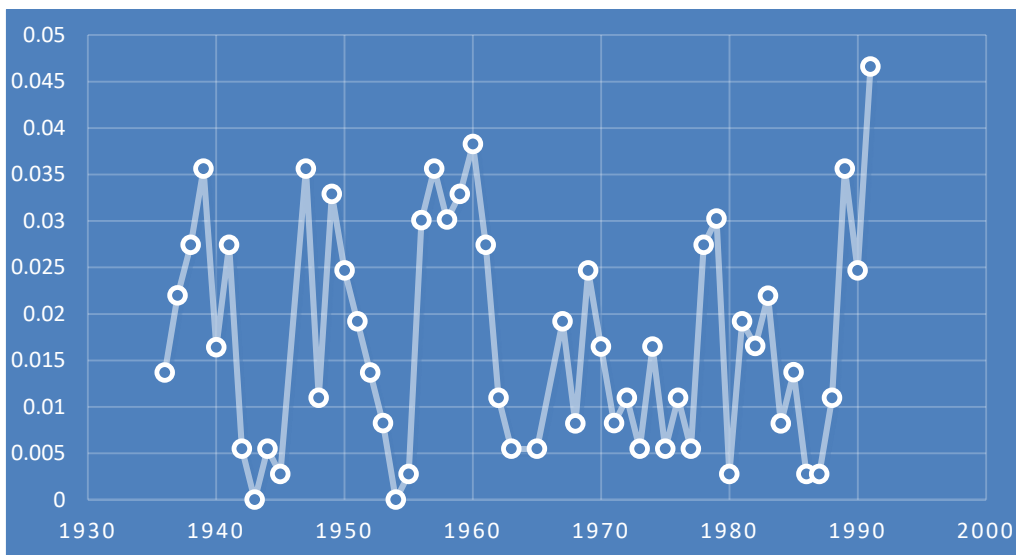


Fig. 3. Quantitative time period analysis of 3σ order anomalies of the magnetic field declination during 1936-1991.

Figure 2 shows shares of more than 2σ anomalies in a concrete year among the whole annual data. In 1936, for example, among the whole data the anomalies were declined from average value by more than 2σ altogether in 8% cases. Like this, in 1959 the number of such anomalies makes only 20% of the annual data. The issue of anomaly measurements according to the distinguished years is quite informative by frequency calculations in this figure.

Figure 3 shows shares of more than 3σ anomalies in a concrete year among the whole annual data. Here, statistical stability is under doubt and verification of statistical hypotheses of strong anomaly zones is prevented as far as their reliability is doubtful. There a question arises: why? It is due their limited involvement

with all the data. However, from 1936 to 1960 their share is more than in 1960-1985, whereas at the end of the 80-s of the past century, processes similar to the ones in 1936-1960 take place.

From *Figures 1, 2 and 3* we may conclude that there is an alternation of rise (more dynamics) and fall in the magnetic declination. We are interested in the proceeding of the data till now. Does the magnetic declination parameter maintain its characteristics? Additionally, it is noteworthy to mention that there an alternation of average declination takes place from year to year, an average of so called *Up* and an average of so called *Down*. For example, so called *Up-s* and *Down-s* (frequent alternations of rise and fall) according to months must be a potential to build a good forecasting model.

If we bring in the following marks for our magnetic declination statistics:

$$X_i, \quad \text{where } (i \text{ is changed } 1:19332)$$

we can build the following matrix:

Let us make components from previous 30 data in time for each i member. Our task is to find such F function, which will find links between any X_i and previous 30 records.

In this way we receive:

$$F(X_i, X_{i+1}, X_{i+2} \dots X_{i+30}) \cong X_{i+31} \quad (1)$$

The whole machine learning theory considers optimization of F , which uses different algorithms. In this paper we will consider Adam Optimization Algorithm for Deep Learning (1).

The normalized learning sampling for Formula (1) in the form of a real table is given as follows:

X1	X2	X3	X29	X30	X(31)
2899	2898	2893	2922	2925	2921
2898	2893	2898	2925	2921	2920
2893	2898	2902	2921	2920	2917
2898	2902	2908	2920	2917	2917
2902	2908	2918	2917	2917	2920
2908	2918	2930	2917	2920	2913
..
2917	2911	2912	2925	2925	2925
2911	2912	2910	2925	2925	2919

Finally, we received a learning sampling with 30 inputs and one output.

A significant detail in the algorithm is the one, which can find links between each reading and its predecessor 30 data. In this case we will construct such a model, which will enable us to make an optimal prognosis by every 30 predecessor reading by means of Adam network.

$$w_{t+1} = (1 - \lambda)w_t - \eta \nabla f_t(w_t) \quad (2)$$

As we know, in Adam algorithm, optimization of (1) is a solution to equation (2) for each necessary iteration (usually, there are 1000 iterations) for weight sampling, where λ is a parameter guiding the weight drop. Of course, here we have L2 regularization for decreasing unjustified weights, which is based on so called penalty principle, finally, on search function modification to have L2 -norm weight vector (4).

Table 1. Matrix of confusion (clarity) of the prognosis by 30-day predecessors of 1936-1991 magnetic field declination.

Target output:	Network output:									
	<2851	2851 .. 2887.6	2887.6 .. 2924.2	2924.2 .. 2960.8	2960.8 .. 2997.4	2997.4 .. 3034	3034 .. 3070.6	3070.6 .. 3107.2	3107.2 .. 3143.8	3143.8 .. 3180.4
2851 .. 2887.6	0	0	2	1	0	0	0	0	0	0
2887.6 .. 2924.2	0	0	65	158	1	0	0	0	0	0
2924.2 .. 2960.8	0	0	0	1810	70	0	0	0	0	0
2960.8 .. 2997.4	0	0	0	197	1813	37	0	0	0	0
2997.4 .. 3034	0	0	0	2	444	1224	73	0	0	0
3034 .. 3070.6	0	0	0	0	1	100	1598	142	0	0
3070.6 .. 3107.2	0	0	0	0	1	3	90	2624	165	2
3107.2 .. 3143.8	0	0	0	0	0	0	0	105	1432	39
3143.8 .. 3180.4	0	0	0	0	0	0	0	2	401	494
3180.4 .. 3217	0	0	0	0	0	0	0	1	3	25

Table 1 shows the qualitative assessment of the obtained prognostic model according to their interval (cluster) distribution. More exactly, here we have real and false coincidences of prognostic and real data, which are expressed quantitatively and by clusters.

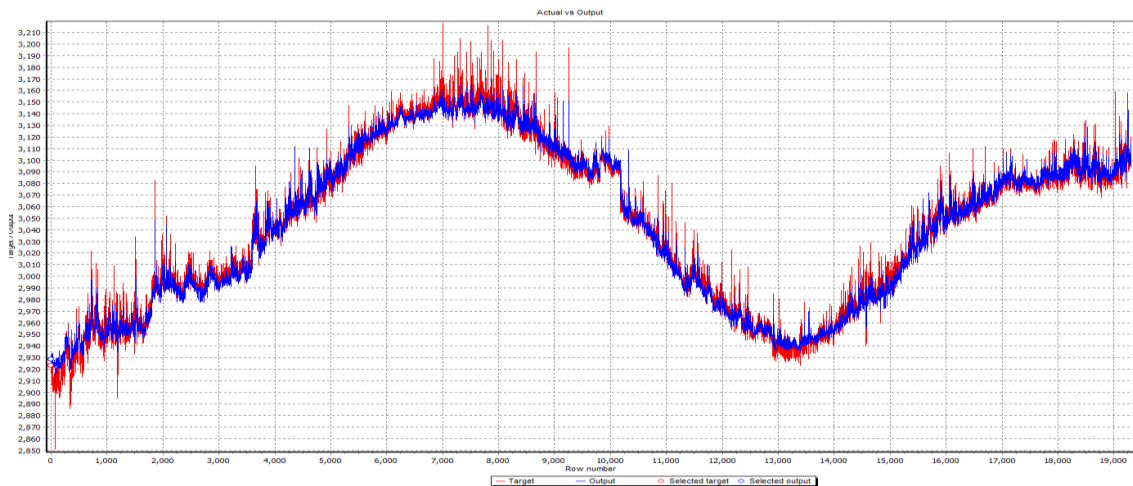


Fig. 4. Scheme of the prognosis by 30-day predecessors of 1936-1991 magnetic field declination.

Figure 4 clearly shows the compatibility between learning sampling and prognostic model. The real data are given in red colour and the prognostic values are shown in blue. It is obvious that the values of the strong anomalies, which took place in 1940-1950, are not reliable. However, taking into account these anomalies, the model revealed similar strong anomalies of the same order more accurately at the end of 1980-s. Finally, we have following characteristics for the constructed model:

	Target	Output	AE	ARE
Mean:	3042.505371	3041.840249	5.740834	0.001887
Std Dev:	68.173232	67.271728	5.886604	0.001927
Min:	2851	2908.161365	0.000849	2.71E-07
Max:	3217	3171.404207	101.941242	0.033066
Correlation: 0.992766				
R-squared: 0.98506				

The obtained result makes it clear that there is a tight correlation (0.992766) between the value of magnetic declination of any day and its predecessor 30-day data. The corrected determination coefficient 0.98506 means that our prognostic model, taking into account the predecessor 30 days, forecasts the probable value of the following day declination by nearly 99% on the eve.

Conclusion

The statistical study of the Dusheti Observatory (overall 19332 readings, y.y.1935-1950) shows that the dynamics of average annual values is characterized with 45-year cycle. Here statistically stable periodicity of increase of magnetic declination variations is distinguished. Namely, increase from minimal meaning to peak value needs approximately 15 years. There was a peak of such variation in 1952 and during the following 15 years a decrease took place. Further on, similar increase was observed during the next 15 years.

Taking into consideration, in our work we constructed an Adam deep learning network and received a high reliability prognostic model. The obtained model can be used for every-day forecasting as well as for more long-term prognoses.

References

- [1] Glatzmaiers G.A., Roberts P.H. Three-dimensional self-consistent computer simulation of a geomagnetic field reversal. Nature. 377(6546), X, 1995, pp. 203–209.
- [2] Loshchilov I., Hutter F. Fixing Weight Decay Regularization in Adam. Archives: 1711.05101 v. 2, 2017.
- [3] Kingma D.P., Ba J. A Method for Stochastic Optimization. 22 Dec 2014 (v.1), last revised 30 Jan 2017 (this version, v. 9).
- [4] Anastasis K. Deep Arbitrage-Free Learning in a Generalized HJM Framework via Arbitrage-Regularization Data". 2020.

მაგნიტური (დუშეთი 1935-1989 წ.წ.) მიხრილობის სტატისტიკა და ღრმა თვითსწავლებადი მოდელი

თ. ქირია, ა. ესაკია, მ. ნიკოლაიშვილი, ე. ლომაძე

რეზიუმე

ნაშრომში წარმოდგენილია დუშეთის ობსერვატორიის მაგნიტური მიხრილობის საკმაოდ ხანგრძლივ (1935 – 1950 წ.წ., სულ 19332 ანათვალი) მონაცემთა ჯერ სტატისტიკური, ხოლო შემდეგ დღეისათვის ფართოდ გავრცელებული ე. წ. მანქანური სწავლების (ML) მეთოდის შესაძლებლობები. მოყვანილია ჰიპოთეზები ზოგიერთი გეომაგნეტური პარამეტრისთვის გარკვეული ფარული კანონზომიერებების და პერიოდულობის დასასაბუთებლად. დადგენილია მაღალი სტატისტიკური მდგრადობის მქონე ე. წ. „მეხსიერებები“. სწორედ ესაა ეტალონური ნიმუშები, რომელსაც ე. წ. ადამის ღრმა სწავლების ქსელის გამოყენებით მიკუთვნების ფუნქციის ასაგებად ვიყენებთ.

Статистика магнитного склонения (Душети, 1935-1989 гг.) и модель глубокого самообучения

Т.В. Кириа, А.А. Эсакиа, М.М. Николайшвили, Э.Д. Ломадзе

Реферат

В статье представлены довольно длительные (1935 - 1950 г.г., всего 19332 отсчета) данные магнитного склонения обсерватории Душети, сначала статистические, а затем, широко используемые, методом машинного обучения (ML). Приводятся гипотезы для обоснования определенных скрытых закономерностей и периодичности некоторых геомагнитных параметров.

С высокой статистической стабильностью установлена т. н. "устойчивость". Это те стандартные образцы, которые мы используем с помощью сети глубокого самообучения Адама для построения функции принадлежности.

The Sharle Statistical Structure

¹Zurab S. Zerakidze, ²Jemal K. Kiria, ²Tengiz V. Kiria

¹Gori State Teaching University

e-mail: zura.zerakidze@mail.ru

²M. Nodia Institute of Geophysics of I. Javakhishvili Tbilisi State University

e-mail: kiria51@yahoo.com

ABSTRACT

In this paper the consistent criteria for testing Hypothesis for the Sharle statistical structure are defined. It is shown that the necessary and sufficient conditions for the existence of these critical are considered. Also the necessary and sufficient conditions for the existence of such criteria for the Sharle strongly statistical structure.

Key words: Consistent criterion, statistical structures.

I. The Sharle statistical structure.

Definition 1.1 we will say that X random value is (see [1]-[6]) the Sharle distribution if this density given by formula

$$f_{sh}(x) = f(x) + \frac{1}{\sigma} \left[\frac{S_k(X)}{6} \cdot Z_u \cdot (U^3 - 3U) + \frac{E_x(X)}{24} \cdot Z_u \cdot (U^4 - 6U^2 + 3) \right]$$

$$f(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-m)^2}{2\sigma^2}}, \quad U = \frac{x-m}{\sigma}, \quad Z_u = \frac{1}{\sqrt{2\pi}} e^{-\frac{u^2}{2}},$$

$s_k(x)$ - asymmetry, $E_x(X)$ -excess

Let (E, S) – be a measurable space. Let

$$\mu(A) = \int_A f_{sh}(x) dx, \quad A \in L(R)$$

Probability Sharle given on $(R, L(R))$, where $f_{sh}(x)$ be spectral density Sharle and $L(R)$ Lebesgue σ -algebra in R . Let $\{\mu_i, i \in I\}$ the corresponding Sharle probability measures.

Definition 1.2. An Object $\{E, S, \mu_h, h \in H\}$ is called a Sharle statistical structure.

Definition 1.3. A Sharle statistical structure $\{E, S, \mu_h, h \in H\}$ is called orthogonal (singular) if a family of Sharle probability measures $\{\mu_i, i \in I\}$ consists of pairwise singular measures (i. e. $\mu_{h'} \perp \mu_{h''} \quad \forall h' \neq h''$).

Definition 1.4. Sharle statistical structure $\{E, S, \mu_h, h \in H\}$ is called weakly separable if then exists a family of S-measurable sets $\{X_h, i \in I\}$ Such that the relations are fulfilled:

$$\mu_h(X_{h'}) = \begin{cases} 1, & \text{if } h = h' \\ 0, & \text{if } h \neq h' \end{cases}$$

Definition 1.5. A Sharle statistical structure $\{E, S, \mu_h, h \in H\}$ Is called separable if there exists a family of S-measurable sets $\{X_h, h \in H\}$ Such that the relations are fulfilled:

1. $\mu_h(X_{h'}) = \begin{cases} 1, & \text{if } h = h' \\ 0, & \text{if } h \neq h' \end{cases}$
2. $\forall h, h' \in H \text{ card}(X_h \cap X_{h'}) < c \text{ if } h \neq h'$

Where c denotes the continuum power

Definition 1.6 A Sharle statistical structure $\{E, S, \mu_h, h \in H\}$ is called strongly separable if there exists a disjoint family of S-measurable sets $\{X_h, h \in H\}$ such that the relations are fulfilled:

$$\mu_h(X_h) = 1, \forall h \in H.$$

Remark 1.1 from strong separability there follows separability there follows orthogonality but not vice versa.

Example 1.1 let $E = R \times R$ (where $R = (-\infty, +\infty)$) and $L(R \times R)$ be a lebesgue σ - algebra of subsets of $R \times R$. As a set of hypotheses consider the set $H = R$. let $X_h = \{(-\infty < x < +\infty), y = h, h \in R\}$. And let

$$\mu(A) = \int_A f_{Sh}(x) dx$$

Be the Sharle linear measures on $X_h, h \in R$

The Sharke statistical structure $\{R \times R, L(R \times R), \mu_h, h \in R\}$ is continuum strongly separable statistical structure.

Let H be the set of hypotheses and let $B(H)$ be σ -algebra of subsets of H which contains all finite subsets of H .

Definition 1.7. we will say that the statistical structure $\{E, S, \mu_h, h \in H\}$ admits a consistent criterium for hypotheses testing if there exist at least one measurable mapping

$$\delta: (E, S) \rightarrow (H, B(H)),$$

Such that

$$\mu_h(\{x: \delta(x) = h\}) = 1, \forall h \in H.$$

The notion and corresponding construction of consistent criteria for hypotheses testing was introduced and studied by Z. Zerakidze (see [5]).

Remark 1.2. if the Sharle statistical structure $\{E, S, \mu_h, h \in H\}$ admits a consistent criterion for hypothesis testing, then the Sharle statistical structure $\{E, S, \mu_h, h \in H\}$ is strongly separable but not vice versa.

(see example 1.1).

Example 1.2. let $E = R \times R \times R = R^3$, let S be a Borel σ -algebra on R^3 . let take S-measurable sets

$$X_h = \begin{cases} -\infty < x < +\infty, & -\infty < y < +\infty, & z = h \text{ if } h \in [0,1]; \\ x = h - 2, & -\infty < y < +\infty, & -\infty < z < +\infty, \text{ if } h \in [2,3]; \\ -\infty < x < +\infty, & y = h - 4, & -\infty < z < +\infty, \text{ if } h \in [4,5]; \end{cases}$$

and assume that μ_h are plane Sharle measures on X_h Then the Sharle statistical structure $\{R^3, S, \mu_h, h \in [0,1] \cup [2,3] \cup [4,5]\}$ is weakly separable, but not strongly separable.

Example 1.3 let $E = [0,1] \times [0,1]$, let $B = [0,1] \times [0,1]$, be a Borel σ -algebra of subsets of E . As a set of hypotheses consider the set $H = [0,1] \cup [2,3]$ let us take the $B = [0,1] \times [0,1]$ – measurable sets

$$X_h = \begin{cases} 0 \leq x \leq 1, & y = h, & \text{if } h \in [0,1]; \\ x = h - 2, & 0 \leq y \leq 1, & \text{if } h \in [2,3]; \end{cases}$$

and denote by $\mu_h, h \in [0,1] \cup [2,3]$ linear, normed, Sharle measures on X_h . Then the Sharle statistical structure $\{[0,1] \times [0,1], B([0,1] \times [0,1]), \mu_h, h \in [0,1] \cup [2,3]\}$, is a separable statistical structure. Suppose that it admits a consistent criterium for hypotheses testing

$$\delta: ([0,1] \times [0,1], B([0,1] \times [0,1])) \rightarrow (H, B(H)),$$

with

$$\mu_h(\{x: \delta(x) = h\}) = 1, \quad \forall h \in [0,1] \cup [2,3].$$

Let's introduce sets $A_1 = \{x: \delta(x) \in [0,1]\}$ and $A_2 = \{x: \delta(x) \in [2,3]\}$ it is clear that A_1 and A_2 $B([0,1] \times [0,1])$ measurable Sets and we have

$$\mu_h(A_1 \cap \{[0,1] \times \{h\}\}) = 1 \quad \forall h \in [0,1] \quad \text{and}$$

$$\mu_h(A_2 \cap \{h - 2\} \times [0,1]) = 1 \quad \forall h \in [2,3]$$

Further, according to the Fubin theorem we conclude that $\mu(A_1) = 1$ and $\mu(A_2) = 1$ (where μ is the Sharle plane measure).

From here, taking into account that $A_1 \cap A_2 = \emptyset$ and $A_1 \cup A_2 = [0,1] \times [0,1]$ we verify that $\mu([0,1] \times [0,1]) = 2$ which contradicts the fact that $\mu([0,1] \times [0,1]) = 1$. Hence, the Sharle separable statistical structure does not admit a consistent criterium for hypotheses testing.

Theorem 1.1 let $\{E, S, \mu_{h_i}, i \in N\}$ ($N = 1, 2, \dots, n, \dots$) Be on orthogonal Sharle statistical structure, then this statistical structure admit a consistent criterium for hypotheses testing.

Proof due to the singulatito of Sharle statistical structure $\{E, S, \mu_h, h \in N\}$ there exists the family of S-measurable sets $\{X_{ik}\}$ such that for any $i \neq k$: $\mu_{h_k}(X_{ik}) = 0$ and $\mu_{h_i}(E - X_{ik}) = 0$, Therefore if consider the sets

$$X_i = \bigcup_{k \neq i} (E - X_{ik}),$$

we get $\mu_{h_i}(X_i) = 0$, Hence, $\mu_{h_i}(E - X_i) = 1$. On the other hand, for $k \neq i$ we have $\mu_k(E - X_i) = 0$, It means that the Sharle statistical structure $\{E, S, \mu_{h_i}, i \in N\}$ is weakly seperable. Therefore, there exists the family of S-measurable sets $\{X_{h_i} \mid i \in N\}$ such, that

$$\mu_{h_i}(X_{h_i}) = \begin{cases} 1, & \text{if } i = j \\ 0, & \text{if } i \neq j \end{cases}$$

Consider now the sets

$$\bar{X}_{h_i} = X_{h_i} - \left(X_{h_i} \cap \left(\bigcup_{k \neq 0} X_{h_k} \right) \right), \quad i \in N$$

It is obvious that these sets are S-measurable disjocht sets and

$$\mu_{h_i}(\bar{X}_{h_i}) = 1, \forall i \in N$$

Let us define the mapping

$$\delta: (E, S) \rightarrow (H, B(H))$$

in the following way: $\delta(\bar{X}_i) = h_i, \forall i \in N$.

Then we have

$$\{x: \delta(x) = h_i\} = \bar{x}_i$$

and

$$\mu_{h_i}(\bar{X}_{h_i}) = \mu_{h_i}(\{x: \delta(x) = h_i\}) = 1, \forall i \in N$$

Hence δ is a consistent criterion for hypothesis testing.

2. The consistent criterion for hypothesis testing of Sharle strongly separable statistical structures.

Let $\{\mu_h, h \in H\}$ be Sharle probability measures defined on the measurable space (E, S) , For each $h \in H$ denote by $\bar{\mu}_h$ the completion of the measure μ_h and denote by $\text{dom}(\bar{\mu}_h)$ the σ -algebra of all $\bar{\mu}_h$ -measurable subsets of E . Let

$$S_1 = \bigcap_{h \in H} \text{dom}(\bar{\mu}_h) .$$

Definition 2.1. A Sharle statistical structure $\{E, S_1, \bar{\mu}_h, h \in H\}$ is called strongly separable if there exists the family of S_1 -measurable sets $\{z_h, h \in H\}$ such that the relations are fulfilled:

- 1 $\bar{\mu}_h(z_h) = 1, \forall h \in H$;
- 2 $Z_{h_1} \cap z_{h_2} = \emptyset \forall h_1 \neq h_2$;
- 3 $\bigcup_{h \in H} z_h = E$.

Definition 2.2. We will say that the orthogonal Sharle statistical structure admits a consistent criterion for testing hypothesis if there exists at least one measurable mapping

$$\delta: (E, S_1) \rightarrow (H, B(H)) , \text{ such that}$$

$$\bar{\mu}_h(\{x: \delta(x) = h\}) = 1, \forall h \in H.$$

Theorem 2.1 in order that the Sharle statistical structure $\{E, S_1, \bar{\mu}_h, h \in H\}$, $\text{card}H = c$ admitted a consistent criterion for hypothesis testing it is necessary and sufficient that this statistical structure was strongly separable (see definition 2.1).

Proof. Necessity. The existence of a consistent criterion for hypothesis testing (see definition 2.2) means that there exist at least one measurable mapping

$$\delta: (E, S_1) \rightarrow (H, B(H))$$

Such that

$$\bar{\mu}_h(\{x: \delta(x) = h\}) = 1, \forall h \in H.$$

Denoting $z_h = \{x: \delta(x) = h\}$ for $h \in H$ we get:

- 1) $\bar{\mu}_h(z_h) = \bar{\mu}_h(\{x: \delta(x) = h\}) = 1, \forall h \in H;$
- 2) $z_{h_1} \cap z_{h_2} = \emptyset \quad \{x: \delta(x) = h_1\} \cap \{x: \delta(x) = h_2\} \quad \forall h_1 \neq h_2, \quad h_1, h_2 \in H;$
- 3) $\bigcup_{h \in H} z_h = \{x: \delta(x) \in H\} = E$

Sufficiency, since the Sharle statistical structure $\{E, S_1, \bar{\mu}_h, h \in H\}$ Is strongly separable (see definition 2.1.) there exist a family $\{z_h, h \in H\}$ of elements of the σ -algebra

$$S_1 = \bigcap_{h \in H} \text{dom}(\bar{\mu}_h).$$

Such that

- 1) $\bar{\mu}_h(z_h) = 1, \forall h \in H;$
- 2) $z_{h_1} \cap z_{h_2} = \emptyset \quad \forall h_1 \neq h_2;$
- 3) $\bigcup_{h \in H} z_h = E$

For $x \in E$ we put

$$\delta(x) = h$$

Where h is a unique hypothesis from the H for which $x \in z_h$ the existence and uniqueness of such hypothesis h can be proved using conditions 2) and 3)

Take now $y \in B(H)$. Then

$$\{x: \delta(x) \in y\} \in \bigcup_{h \in y} z_h$$

We have to show that

$$\{x: \delta(x) \in y\} \in \text{dom}(\bar{\mu}_{h_0})$$

For each $h_0 \in H$

If $h_0 \in y$ then

$$\{x: \delta(x) \in y\} \in \bigcup_{h \in y} z_h = z_{h_0} \cup \left(\bigcup_{h \in y - \{h_0\}} z_h \right)$$

On the one hand the conditions 1), 2) and 3) follows that

$$z_{h_0} \in S_1$$

On the other hand the inclusion

$$\bigcup_{h \in y - \{h_0\}} z_h \subseteq (E - z_{h_0})$$

Implies that

$$\bar{\mu}_{h_0} \left(\bigcup_{h \in y - \{h_0\}} Z_h \right) \leq \bar{\mu}_{h_0}(E - Z_0) = 0$$

And hence

$$\bar{\mu}_{h_0} \left(\bigcup_{h \in y - \{h_0\}} Z_h \right) = 0.$$

And hence

$$\bigcup_{h \in y - \{h_0\}} Z_h \in \text{dom}(\bar{\mu}_{h_0}).$$

Since $\text{dom}(\bar{\mu}_{h_0})$ Is σ -algebra, we conclude that

$$\{x: \delta(x) \in y\} = Z_{h_0} \cup \left(\bigcup_{h \in y - \{h_0\}} Z_h \right) \in \text{dom}(\bar{\mu}_{h_0}).$$

If $h_0 \notin y$ Then

$$\{x: \delta(x) \in y\} \in \bigcup_{h \in y} Z_h \subseteq (E - Z_{h_0})$$

And we conclude that

$$\bar{\mu}_h(\{x: \delta(x) \in y\}) = 0.$$

The last relation implies that

$$\{x: \delta(x) \in y\} \in \text{dom}(\bar{\mu}_{h_0}), \forall y \in B(H).$$

Thus we have shown the validity of the relation

$$\{x: \delta(x) \in y\} \in \text{dom}(\bar{\mu}_{h_0})$$

For an arbitrary $h_0 \in H$ Hence,

$$\{x: \delta(x) \in y\} \in \bigcap_{h \in H} \text{dom}(\bar{\mu}_h) = S_1$$

We have shown that the map

$$\delta: (E, S_L) \rightarrow (H, B(H))$$

Is measurable map

Since $B(H)$ contains all singletons of H , we ascertain that $\bar{\mu}_h(\{x: \delta(x) = h\}) = \bar{\mu}_h(Z_h) = 1, \forall h \in H$.

References

- [1] Borovkov A.A. Mathematics statistics. Estimation of parameters. Testing of hypotheses. "Nauka", Moscow, 1984, English transl. mathematical statistics. Gordon. Breach. Amsterdam 1998.
- [2] Kharazishvili A. Topological aspects of measure theory (in Russian) "Naukova Dumka", Kiev, 1984.

- [3] Kharazishvili A. Nonmeasureable Sets and Functions. North-Holland Mathematics Studies. Elsevier, Amsterdam, 2004.
- [4] Zerakidze Z. On weakly divisible and divisible families of probability measures (in Russian). Bull. Acad. Sci. Georg. SSR 113(2), 1984, pp. 273-275.
- [5] Zerakidze Z. Generalization criteria of Neiman-Pearson. Georgian Technical University. Proceedings of the International Scientific Conference on "Information Technologies", Tbilisi, 2008, pp. 22-24.
- [6] Zerakidze Z., Purtokhia O. The weakly consistent, strongly consistent and consistent estimates of the parameters. Reports of Enlarged Sessions of the I. Vekua Institute of Applied Mathematics, 31, 2017, pp. 151-154.
- [7] Danko P. Ye., Popov A.G., Kozhevnikova T.Ya., Danko S. P. Vysshaya matematika v uprazhneniyakh i zadachakh, chast' 2. «Mir i obrazovaniye», Moskva, 2009.

შარლეს სტატისტიკური სტრუქტურები

ზ. ზერაკიძე, ჯ. ქირია, თ. ქირია

რეზიუმე

სტატიაში აგებულია შარლეს ძლიერად და სუსტად განცალკეადი სტრუქტურები, რომელთათვისაც არ არსებობს ჰიპოთეზათა შემოწმების ძალდებული კრიტერიუმი. ამავე სტატიაში განხილულია უფრო ფართე σ -ალგებრა და ახალი განმარტება ძლიერად განცალკეადი სტატისტიკური სტრუქტურები, რომელთათვისაც ყოველთვის არსებობს ჰიპოთეზა შემოწმების ძალდებული კრიტერიუმი.

Статистические структуры Шарлье

З. С. Зеракидзе, Дж. К. Кириа, Т. В. Кириа

Резюме

В статье построены сильно и слабо разделимые статистические структуры Шарлье, для которых не существует состоятельных критериев для проверки гипотез. Далее строится более широкая σ – алгебра и по новому определены сильно разделимые статистические структуры Шарлье, для которых всегда существуют состоятельные критерии для проверки гипотез.

The Consistent Criteria for Hypotheses Testing for Sharle Statistical Structure

¹Zurab S. Zerakidze, ²Jemal K. Kiria, ²Tengiz V. Kiria

¹Gori State Teaching University

e-mail: zura.zerakidze@mail.ru

²M. Nodia Institute of Geophysics of I. Javakhishvili Tbilisi State University

e-mail: kiria51@yahoo.com

ABSTRACT

In this paper, we define Sharle statistical Structure such that the probability of any kind of error is zero for given criterion. The necessary and sufficient conditions for the existence of such criteria are given.

Key words: Consistent criterion, statistical structures.

I. Introduction

Statistics of random processes is used in various fields of science and technology (for example, in theoretical physics, genetic, economics, radio physics, geophysics).

When using random processes as models of real phenomena, the question of determining the probabilistic characteristics of the processes arises. To determine these characteristics should be use statistical methods. Among the problems of statistics, a class of problems is distinguished in which the number of observations is unique.

Despite the uniqueness of observation, in many cases, one can authentically determine the values of unknown distribution parameters or reliably choose one of an infinite number of competing hypotheses about the exact form of the distribution. In the case when a parameter or hypothesis is determined by one observation reliably, it is said that for it there exists a consistent estimate of parameter or a consistent criterion for hypothesis testing. This article is devoted to the question a consistent criterion for hypothesis testing for Sharle statistical structure.

It is said that an error of the h-th kind of the δ criterion accors, if the criterion rejects the main hypothesis of H_0 . The following probability $\{\alpha_h(\delta) = \mu_h(\{x: \delta(r) \neq h\})$ Is called the probability of an erroe of the h-th kind for a given criterion δ . Example 1 nd 2 show a general trend-when we decrease one of the probabilities of an error, the other, as a rule, increases.

Example 1.1 Consider the case when there are two simple hypseses h_1 and h_2 about the distribution of population and criterion $\delta: R^n \rightarrow \{h_1, h_2\}$ such that $\delta(x) \equiv h_1$ then the probability of an error of the first kind is zero $\alpha_1 = \mu_{h_1}(\{x: \delta(r) \neq h_1\}) = 0$ and the probability of an error of the second kind is equal to one $\alpha_2 = \mu_{h_2}(\{x: \delta(r) \neq h_2\}) = 1$.

Example 1.2. There are observations from the normal distribution in R with variation one and different means $a \in R$ And two simple hypotheses $H_1 = \{a = 0\}$ And $H_2 = \{a = 1\}$ consider the following criteria

$$\delta(x) = \begin{cases} H_1 & \text{if } x \leq c; \\ H_2 & \text{if } x > c, \end{cases}$$

From $c \in R$. It obvious that with the increase of the number c the probability of an error of the first type decreases, and the probability of an error of the second kind increases.

2. The consistent criteria for hypotheses testing of Sharle statistical structure let (E, S) – be a measurable space. The following definitions are taken from ([1] – [7])

Definition 2.1. we will say that X random value is the Sharle (see [7]) distribution if this density given by formul

$$f_{sh} = f(x) + \frac{1}{\sigma} \left[\frac{S_k(x)}{6} \cdot Z_u \cdot (u^3 - 3u) + \frac{E_x(x)}{24} Z_u \cdot (u^4 - 6u^2 + 3) \right] \quad (1)$$

$$f_x = \frac{1}{\sqrt{2\pi}\sigma} \cdot e^{-\frac{(x-m)^2}{2\sigma^2}} \quad u = \frac{x-m}{\sigma} \quad Z_u = \frac{1}{\sqrt{2\pi}} \cdot e^{-\frac{u^2}{2}}$$

$S_k(x)$ - asymmetry, $E_x(x)$ - excess

Let

$$\mu(A) = \int_A f_{sh}(x) dx, \quad A \in L(R)$$

Probability Sharle given on $(R, L(R))$ where $f_{sh}(x)$ Be spectral density Sharle and $L(R)$ lebesgue σ –algebra in R .

Let $\{\mu_i, i \in I\}$ the corresponding Sharle probability measures .

Definition 2.2. an object $\{E, S, \mu_i, i \in I\}$ is called a Sharle statistical structure.

Definition 2.3. a Sharle statistical structure $\{E, S, \mu_i, i \in I\}$ Is called orthogonal (singular) if a family of Sharle probability measures $\{\mu_i, i \in I\}$ Consists of pairwise singular measures (i.e. $\mu_i \perp \mu_j, \forall i \neq j$).

Example 2.1 let $E = [0, 1]$ S be a Borel σ –algebra of subsets of $[0, 1]$.

Let $\mu_1(B) = 2l\left(B \cap \left[0, \frac{1}{2}\right]\right)$, $\mu_2(B) = 2l\left(B \cap \left[\frac{1}{2}, 1\right]\right)$ and $\mu_3(B) = 3l\left(B \cap \left[0, \frac{1}{3}\right]\right)$ BCS., where l lebesgue measure on S . then $\mu_1 \perp \mu_2$ And $\mu_1 \perp \mu_3$ But $\mu_2 \perp \mu_3$ is not orthogonal to μ_3 .

Definition 2.4. A Sharle statistical structure $\{E, S, \mu_i, i \in I\}$ is called weakly separable if there exists a family of S -measurable sets $\{X_i, i \in I\}$ such that the relations are fulfilled:

$$\mu_i(X_j) = \begin{cases} 1, & \text{if } i = j \\ 0, & \text{if } i \neq j \end{cases}$$

Definition 2.5 A Sharle statistical structure $\{E, S, \mu_i, i \in I\}$ Is called separable if there exists a family of S -measurable sets $\{X_i, i \in I\}$ Such that the relations are fulfilled:

$$1) \mu_i(X_j) = \begin{cases} 1, & \text{if } i = j \\ 0, & \text{if } i \neq j \end{cases}$$

$$2) \forall i, j \in I \quad \text{card}(X_i \cap X_j) < c \text{ if } i \neq j,$$

Where c denotes the continuum power.

Definition 2.6. A Sharle statistical structure $\{E, S, \mu_i, i \in I\}$ is called strongly separable if there exists a disjoint family of S -measurable sets $\{X_i, i \in I\}$ such that the relations are fulfilled: $\mu_i(X_i) = 1, \forall i \in I$

Remark 2.1. From strong separability there follows separability there follows weak separability and from weak separability there follows orthogonality but not vice versa.

Remark 2.2. On an arbitrary set E of continuum power one can define a Sharle orthogonal statistical structure having the maximal possible power equal to 2^{2^c} and Sharle weakly separable statistical structure having the maximal possible power equal to 2^c , and a Sharle strongly statistical structure with the maximal possible power equal c , where c is continuum power.

Lemma 2.1. If a Sharle statistical structure $\{E, S, \mu_i, i \in I\}$ Is separable then it is orthogonal.

Proof. If a Sharle statistical structure $\{E, S, \mu_i, i \in I\}$ is separable, then there exists S -measurable sets $\{X_i, i \in I\}$ Such that

$$\mu_i(X_j) = \begin{cases} 1, & \text{if } i = j \\ 0, & \text{if } i \neq j \end{cases} \quad i, j \in I$$

Since $\mu_i(X_i) = 1, \forall i \in I$ and $\mu_j(X_i) = 0, \forall i \in I$ We have $\mu_i(E - X_i) = 0$. Hence, the measurable μ_i and μ_j Are orthogonal.

The notion and corresponding construction of consistent criteria for hypotheses testing was introduced and studied by Z.Zerakidze (see [6]).

Definition 2.7 We consider the concept of the hypothesis as any assumption that determines the form of the distribution on population.

Definition 2.8. A statical criterion is any measurable mapping $\delta(E, S) \rightarrow (H, B(H))$ where H be the set of hypotheses and $B(H)$ be σ -algebra of subsets of H which contains all finite subsets of H .

Definition 2.9. We will say that the Sharle statistical structure $\{E, S, \mu_h, h \in I\}$ admits a consistent criterion for hypothesis testing if there exists at least one measurable mapping

$$\delta(E, S) \rightarrow (H, B(H))$$

Such that

$$\mu_h: (x: \delta(x) = h) = 1 \quad \forall h \in H$$

Definition 3.0. the probability

$$\alpha_h(\delta) = \mu_h: (x: \delta(x) \neq h)$$

Is called the probability of error of the h -th type for a given criterion....

Definition 3.1 we will say that the Sharle statistical structure $\{E, S, \mu_i, i \in I\}$ admits a consistent criterion for hypothesis testing of any parametric function is for any real bounded measurable function

$$g: (H, B(H)) \rightarrow (R, B(R))$$

There exists at least one measurable function

$$f : (E, S) \rightarrow (R, B(R))$$

Such that

$$\mu_h : (x: f(x) = g(h)) = 1 \quad \forall h \in H$$

Definition 3.1. We will say that the Sharle statistical structure $\{E, S, \mu_i, i \in I\}$ admits an unbiased criterion for hypothesis testing if for any real bounded function

$$g : (H, B(H)) \rightarrow (R, B(R))$$

There exists at least one measurable function

$$f : (E, S) \rightarrow (R, B(R))$$

Such that

$$\int_E f \mu_h(dx) = g(x) \quad , \quad \forall h \in H$$

Theorem 2.1. Let the Sharle statistical structure $\{E, S, \mu_h, h \in I\}$ Admit a consistent criterion for hypothesis testing of any parametric function, then the Sharle statistical structure $\{E, S, \mu_h, h \in I\}$ Is weakly separable.

Theorem 2.2. Let the Sharle statistical structure $\{E, S, \mu_h, h \in I\}$ Admit a consistent criterion for hypothesis testing, then the statistical structure $\{E, S, \mu_h, h \in I\}$ Admits a consistent criterion for hypothesis testing of any parametric function, which in turn, implies the existence of an unbiased criterion for hypothesis testing.

Theorem 2.1. Since the Sharle statistical structure $\{E, S, \mu_h, h \in I\}$ Admits a consistent criteria of hypothesis any parametric function, denote by $f(x)$ one of the corresponding consistent criterion for indicator $I_{h'}(h)$. Hence, for the sets

$$\{x: f_h(x) = I_{h'}(h)\} = X_h$$

We have

$$\mu_{h'}(h) = \begin{cases} 1, & \text{if } h' = h \\ 0, & \text{if } h' \neq h \end{cases}$$

The Sharle statistical stricture $\{E, S, \mu_h, h \in I\}$ is weakly separable.

Proof Theorem 2.2. Since Sharle statistical structure $\{E, S, \mu_h, h \in I\}$ admits a consistens criteria of hypothesis testing there existing measurable mapping

$$\delta : (E, S) \rightarrow (H, B(H))$$

Such that

$$\mu_h(x: \delta(x) = h) = 1 \quad \forall h \in H$$

Let

$$g : (H, B(H)) \rightarrow (R, B(R))$$

Be any real, bounded, measurable function and define the function δ_g as follows:

$$\delta_g = g(\delta(x))$$

Then we obtain $\{x: \delta_g = g(h)\} = \{x: \delta(x) = h\}$

$$\mu_h(x: \delta(x) = g(h)) = \mu_h(x: \delta(x) = h) = 1 \quad \forall h \in H \quad \text{and}$$

$$\int_E f \mu_h(dx) = g(x) \quad , \quad \forall h \in H$$

Theorem 2.3. If the Sharle statistical structure $\{E, S, \mu_h, h \in I\}$ Admits a consistent δ criteria of hypothesis testing then this statistical structure $\{E, S, \mu_h, h \in I\}$ is strongly separable, but not versa.

Proof. Since the Sharle statistical stricture admits a consisten criterion of hypothesis testing there existing δ measersble maping

$$\delta : (E, S) \rightarrow (H, B(H))$$

Such that

$$\mu_h(x: \delta(x) = h) = 1 \quad \forall h \in H$$

Let $X_h = \{x: \delta(x) = h\} = 1$. $X_h \cap X_{h'} = \emptyset, \quad \forall h \neq h'$ and

$$\mu_h(X_h) = \mu_h(x: \delta(x) = h) = 1 \quad \forall h \in H$$

Theorem 2.4. The Sharle statistical structure $\{E, S, \mu_h, h \in I\}$ admit a consistent criteria for hypothesis testing if and only if the probability of error of any kind is equal to zero for the criterion δ .

Proof. Necessity since the sharle statistical structure $\{E, S, \mu_h, h \in I\}$ admit a consistent criterion for hypothesis testing, there exists a measurable maping

$$\delta : (E, S) \rightarrow (H, B(H))$$

Such that

$$\mu_h\{x: \delta(x) = h\} = 1 \quad \forall h \in H$$

There for

$$\alpha_h(\delta) = \mu_h(x: \delta(x) \neq h) = 0, \quad \forall h \in H$$

Sufficiency since thes probability of error of any kind is equal to zero, we have

$$\alpha_k(\delta) = \mu_h(x: \delta(x) \neq h) = 0, \quad \forall h \in H$$

On the ather hand

$$\{x: (\delta(x) = h) \cup (x: \delta(x) \neq h)\} = E$$

and

$$\mu_h[(x: (\delta(x) = h) \cup (x: \delta(x) \neq h))] = \mu_h(E) = 1$$

and

$$\begin{aligned} & \mu_h[(x: \delta(x) = h) \cup (x: \delta(x) \neq h)] = \\ & = \mu_h(x: \delta(x) = h) + \mu_h(x: \delta(x) \neq h) = \mu_h(x: (\delta(x) = h)) = 1, \quad \forall h \in H \end{aligned}$$

Example 2.1. let $E = R \times R$ (where $R = (-\infty, +\infty)$ and $B(R \times R)$ be a Borel σ -algebra of subsets of $R \times R$. As a set of hypotheses consider the set $H = Q^+$, where Q^+ is the set of positive rational numbers.

Let $X_h = \{-\infty < x < +\infty, y = h, h \in Q^+\}$

And let

$$\mu_h(A) = \int_A f_{sh}(x) dx$$

be the Sharle linear measures on $X_h, h \in Q^+$.

The Sharle statistical structure $X_h = \{R \times R, B(R \times R), \mu_h, h \in Q^+\}$ is countable Sharle strongly separable statistical structure.

If we now define the mapping

$$\delta: (R \times R, B(R \times R)) \rightarrow (Q^+, B(Q^+))$$

By the formula

$$\delta(X_h) = h, \quad h \in Q^+$$

We get

$$\mu_h(\{x: \delta(x) = h\}) = 1, \quad \forall h \in Q^+$$

Thus, this Sharle statistical structure $\{R \times R, B(R \times R), \forall h \in Q^+\}$ admits a consistent criterion for hypothesis testing.

References

- [1] Ibramkhalilov I., Skorokhod A. Consistent estimates of parameters of random processes (in Russian) "Naukova Dumka", Kiev, 1980.
- [2] Kharazishvili A. Topological aspects of measure theory (in Russian) "Naukova Dumka", Kiev, 1984.
- [3] Kharazishvili A. Nonmeasureable Sets and Functions. North-Holland Mathematics Studies. Elsevier, Amsterdam, 2004.
- [4] Zerakidze Z. On weakly divisible and divisible families of probability measures (in Russian). Bull. Acad. Sci. Georg. SSR 113(2), 1984, pp. 273-275.
- [5] Zerakidze Z., Purtokhia O. The weakly consistent, strongly consistent and consistent estimates of the parameters. Reports of Enlarged Sessions of the I. Vekua Institute of Applied Mathematics, 31, 2017, pp. 151-154.
- [6] Zerakidze Z. Generalization criteria of Neiman-Pearson. Georgian Technical University. Proceedings of the International Scientifics Conference on "Information Technologies", Tbilisi, 2008, pp. 22-24.
- [7] Danko P. Ye., Popov A.G., Kozhevnikova T.YA., Danko S. P. Vysshaya matematika v uprazhneniyakh i zadachakh, chast' 2. «Mir i obrazovaniye», Moskva, 2009.

ჰიპოთეზათა შემოწმების ძალდებული კრიტერიუმები შარლეს სტატისტიკური სტრუქტურებისათვის

ზ. ზერაკიძე, ჯ. ქირია, თ. ქირია

რეზიუმე

სტატიაში განმარტებულია შარლეს სტატისტიკური სტრუქტურები. აგებულია ისეთი შარლეს სტატისტიკური სტრუქტურა, რომლისთვისაც ჰიპოთეზათა შემოწმების ძალდებული კრიტერიუმისათვის ყველა რიგის შეცდომის ალბათობა ნულის ტოლია.

Состоятельные критерии для проверки гипотез статистических структур Шарлье

З. С. Зеракидзе, Дж. К. Кириа, Т. В. Кириа

Резюме

В этой статье мы определяем статистическую структуру Шарлье так, что вероятность любого вида ошибки равна нулю для данного критерия. Приведены необходимые и достаточные условия существования таких критериев.

Modeling the Distribution of Hailstones by Mean Max Sizes on the Territory of Kakheti (Georgia) using Data of the Freezing Level in the Atmosphere and Radar Measurements

**¹Avtandil G. Amiranashvili, ²Nana R. Bolashvili, ²Zaza M. Gulashvili,
¹Nino K. Jamrishvili, ²Nikoloz E. Suknidze, ¹Khatia Z. Tavidashvili**

*M. Nodia Institute of Geophysics of Iv. Javakhsishvili Tbilisi State University, 1 M Alexsidze Str 0160, Tbilisi, Georgia,
avtandilamiranashvili@gmail.com*

Vakhushti Bagrationi Institute of Geography of Iv. Javakhsishvili Tbilisi State University, Tbilisi, Georgia

ABSTRACT

Results of modeling of the distribution of hailstones by mean max diameter (D) on the territory of Kakheti (Georgia) using data of the freezing level in the atmosphere and radar measurements of hail max sizes in clouds are presented.

Maps of the distribution of hail by the average maximum diameter in the territory of Kakheti for individual months, from April to September, have been built. The vertical distribution of D on the indicated territory in the range of heights from 0.11 to 3.84 km was studied.

Key words: *Hail, map of hail distribution by size.*

Introduction

Hail phenomena occur in many regions of the world [1-3], including Georgia [4-7]. Annual global losses of agricultural products from hail damage range from 4 to 18% of the harvest, and in monetary terms, they exceed 11 billion US dollars [<https://www.meteorf.ru/activity/activ/antigrad/obs-info/>]. At the same time, with regard to damage from hail, Georgia is one of the most hail-hazardous countries in the world. Therefore, the problem of hail in this country is devoted to numerous works covering a wide range of studies, such as climatology of hail [5,8-15], radar observation on hail processes [16-19], theoretical and experimental studies of the mechanisms of hail formation [20-22], methods of impact on hail processes [18,23,24], analysis of impact results [25-28], etc.

To solve various problems of scientific or applied significance (the impact of climate change on hail processes, comparison of experimental data on hailstorms with theoretical models of hail processes, assessment of the expected damage from hailstorms, planning of work on active impacts on hail processes, etc.), detailed information on the spatial-temporary characteristics of hail distributions and its sizes on different locations is necessary. Corresponding ground-based network studies (the number of days with hail per year, determination of the size of hail, their structure, kinetic energy, etc.) have been widely carried out and are being carried out both in different countries of the world [2,3,29-34] and in Georgia [1,4,5-15].

To construct of spatial-temporary maps of the distribution of hail processes, data from radar observations of convective clouds are also used [3,16,19,22,25,35-37]. In particular, in the paper [19] presents the results of a statistical analysis of such parameters of hail processes for separate municipalities of

Kakheti in the period from 2016 to 2019, as: the maximum height of hail clouds, the maximum diameter of hailstone in clouds, the number of hail clouds of various categories, repetition of hail clouds of various categories, the mean hail hazard relative ratio G. It was found that during the study period, the greatest hail hazard was observed in the Gurjaani municipality (G = 1.74), and the smallest in the Dedoplistskaro municipality (G = 0.39).

This work is a continuation of the study [19]. Results of modeling of the distribution of hailstones by mean max sizes on the territory of Kakheti (Georgia) using data of the freezing level in the atmosphere and radar measurements of hail sizes in clouds are presented below.

Study area, material and methods

Study area – Kakheti region of Georgia. Data of meteorological radar “METEOR 735 CDP 10 - Doppler Weather Radar” of Anti-hail service of Georgia about the max diameter of hailstones in the clouds (cm) - radar products HAILSZ (Size) [38] - are used.

Period of observation: April-September, 2016-2019. The area of shielded from the hail territory - 11 309.5 km².

The expected diameter of hailstones falling out to the earth's surface according to the Zimenkov-Ivanov model of hail melting in the atmosphere [1,39-41] by taking into account the radar data about their max diameter in the clouds and freezing level in atmosphere was calculated [42, 43].

To calculate the mean max diameter of hailstones (D) on the surface of the earth, the territory of Kakheti was divided into 465 squares, the range of heights was 0.11 ÷ 3.84 km. The monthly average values of the max sizes of hailstones and their 99% values of the lower and upper levels of the average were calculated.

The initial dimensions of hailstones in clouds from April to September in Table 1 are presented.

Table 1. The statistical characteristics of hailstones in the clouds above Kakheti territory by mean max diameter from April to September 2016-2019 (cm).

Parameter	April	May	June	July	August	September
Mean	1.17	1.69	1.97	1.97	1.27	1.79
Min	0.16	0.09	0.09	0.29	0.09	0.20
Max	2.55	4.30	4.83	3.58	4.05	3.58
Range	2.40	4.21	4.74	3.29	3.96	3.38
St Dev	0.65	0.90	0.96	0.82	0.94	0.87
σ_m	0.10	0.06	0.07	0.09	0.22	0.11
99%_Low	0.91	1.53	1.79	1.73	0.70	1.49
99%_Upp	1.42	1.84	2.16	2.21	1.84	2.08

GIS technologies to construct maps of the distribution of hailstones by size near the surface of the earth in the territory of Kakheti were used.

For the data analysis the standard statistical methods are used. The following designations of statistical information are used below: Mean – average values; Min – minimal values; Max - maximal values; Range - Max – Min; St Dev - standard deviation; σ_m - standard error; %; 99% _Low and 99%_Upp – 99% of lower and upper levels of the mean accordingly.

Results

Results in Fig. 1-12 and Table 2,3 are presented.

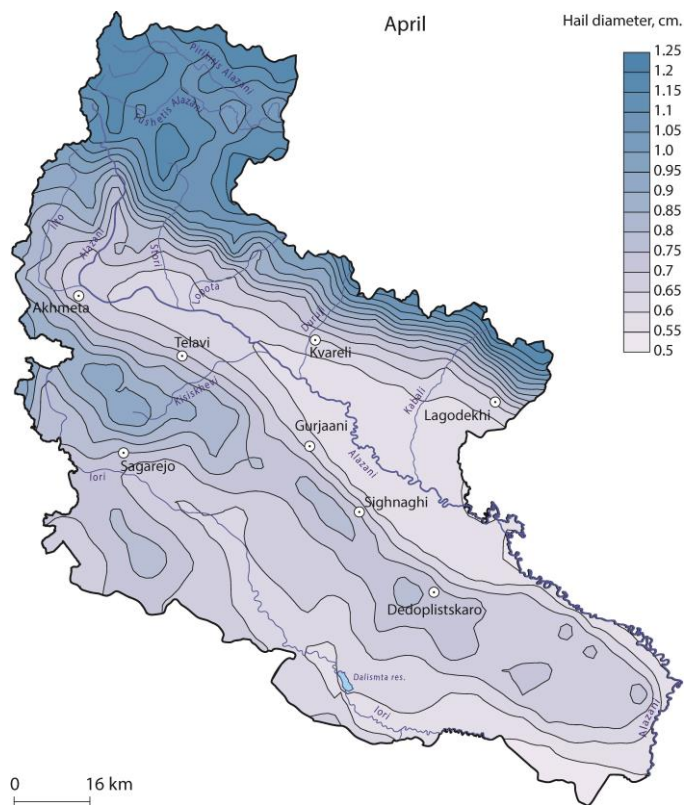


Fig. 1. Distribution of hailstones by mean max diameter on the territory of Kakheti in April.

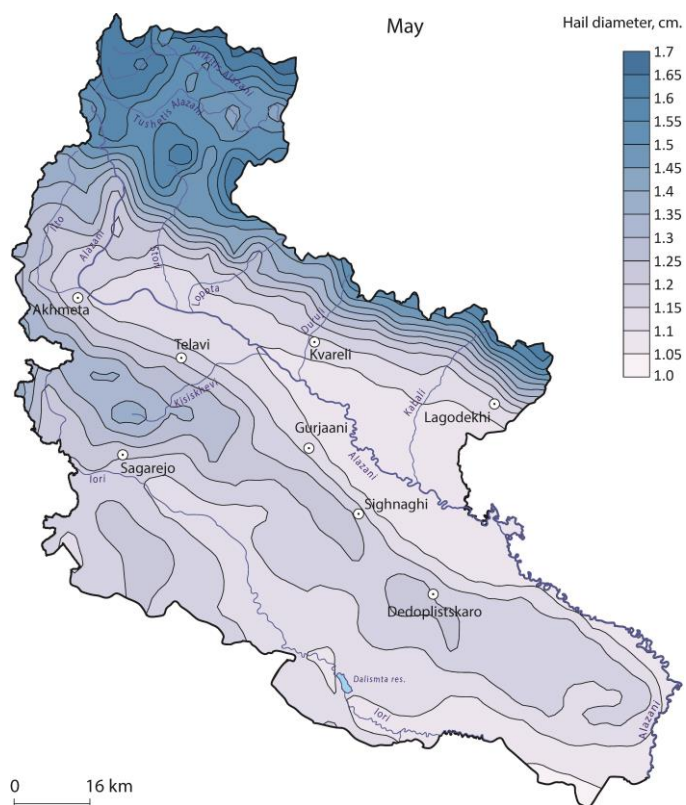


Fig. 2. Distribution of hailstones by mean max diameter on the territory of Kakheti in May.

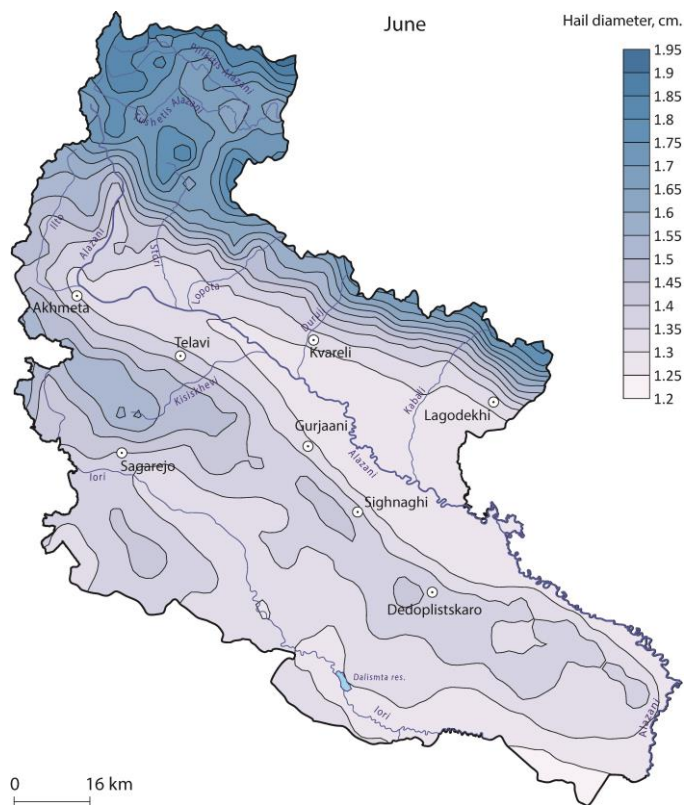


Fig. 3. Distribution of hailstones by mean max diameter on the territory of Kakheti in June.

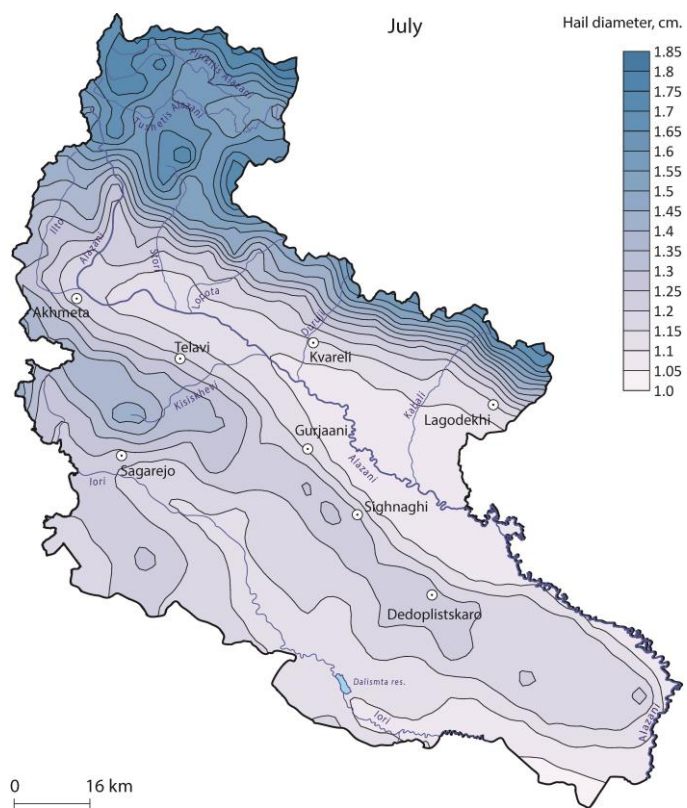


Fig. 4. Distribution of hailstones by mean max diameter on the territory of Kakheti in July.

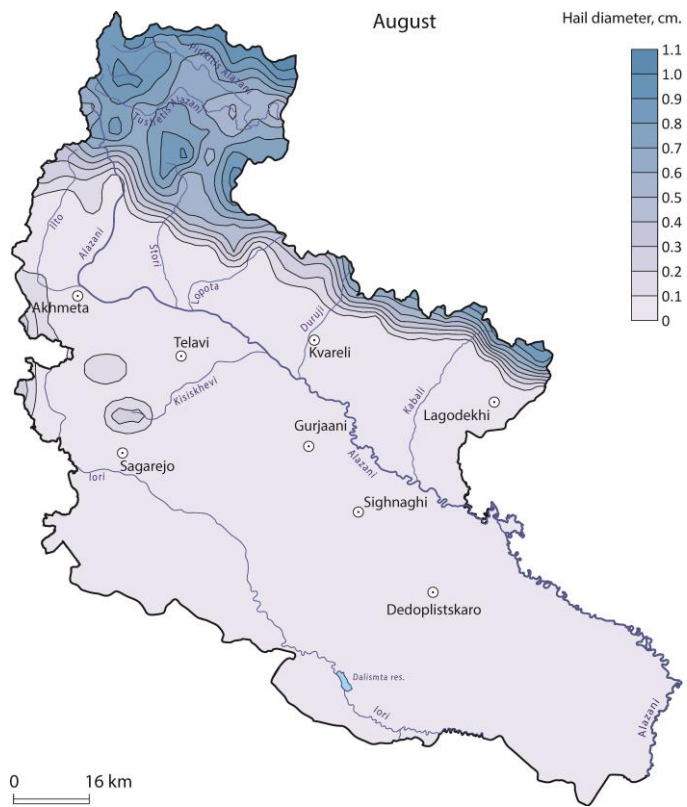


Fig. 5. Distribution of hailstones by mean max diameter on the territory of Kakheti in August.

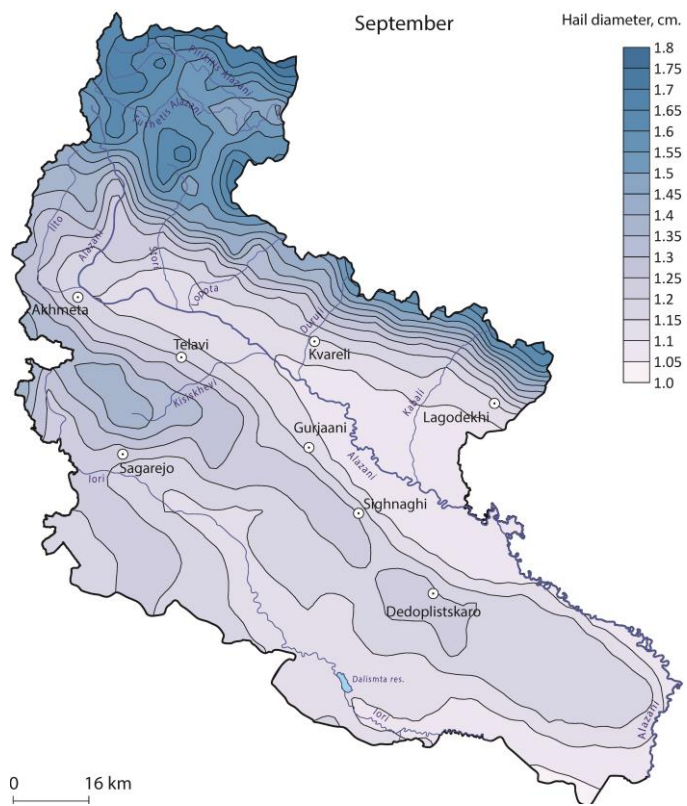


Fig. 6. Distribution of hailstones by mean max diameter on the territory of Kakheti in September.

Fig. 1-6 clearly demonstrated distribution of hailstones by mean max diameter on the territory of Kakheti from April to September.

The statistical characteristics of hailstones by mean max diameter on the earth surface in Kakheti from April to September in Table 2 and 3 are presented. In particular, as follows from Table 2,3 mean values of D to full Kakheti territory change from 0.12 cm (August) to 1.41 cm (June), mean values of D_99%_Low change from 0 cm (August) to 1.17 cm (June) and mean values of D_99%_Upp – from 1.06 cm (August) to 1.65 cm (June).

Table 2. The statistical characteristics of hailstones by mean max diameter on the earth surface in Kakheti from April to June.

Month	April			May			June		
Parameter	99%_Low	Mean	99%_Upp	99%_Low	Mean	99%_Upp	99%_Low	Mean	99%_Upp
Mean	0.29	0.75	1.09	1.02	1.23	1.42	1.17	1.41	1.65
Min	0.00	0.52	0.92	0.82	1.05	1.26	0.96	1.24	1.49
Max	0.91	1.17	1.42	1.53	1.69	1.84	1.75	1.94	2.12
Range	0.91	0.65	0.50	0.72	0.65	0.58	0.79	0.70	0.64
St Dev	0.29	0.18	0.14	0.17	0.15	0.14	0.17	0.15	0.14
σ_m	0.01	0.01	0.01	0.01	0.01	0.01	0.01	0.01	0.01
99%_Low	0.25	0.73	1.07	1.00	1.21	1.40	1.15	1.39	1.63
99%_Upp	0.32	0.78	1.11	1.04	1.24	1.43	1.19	1.43	1.66

Table 3. The statistical characteristics of hailstones by mean max diameter on the earth surface in Kakheti from July to September.

Month	July			August			September		
Parameter	99%_Low	Mean	99%_Upp	99%_Low	Mean	99%_Upp	99%_Low	Mean	99%_Upp
Mean	0.88	1.25	1.58	0.00	0.12	1.06	0.81	1.24	1.61
Min	0.59	1.04	1.40	0.00	0.00	0.82	0.54	1.05	1.46
Max	1.60	1.86	2.11	0.42	1.11	1.72	1.49	1.79	2.09
Range	1.01	0.82	0.71	0.42	1.11	0.90	0.95	0.75	0.63
St Dev	0.22	0.18	0.15	0.02	0.26	0.20	0.21	0.16	0.14
σ_m	0.01	0.01	0.01	0.00	0.01	0.01	0.01	0.01	0.01
99%_Low	0.85	1.23	1.56	0.00	0.09	1.04	0.79	1.22	1.60
99%_Upp	0.90	1.27	1.60	0.00	0.15	1.08	0.84	1.26	1.63

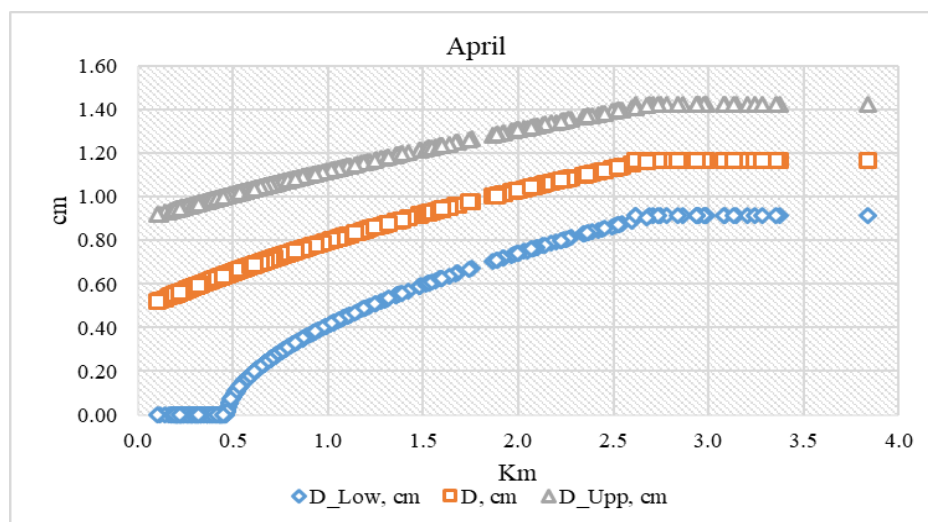


Fig. 7. Vertical distribution of mean max diameter of hailstones and their 99% lower and upper levels in Kakheti in April.

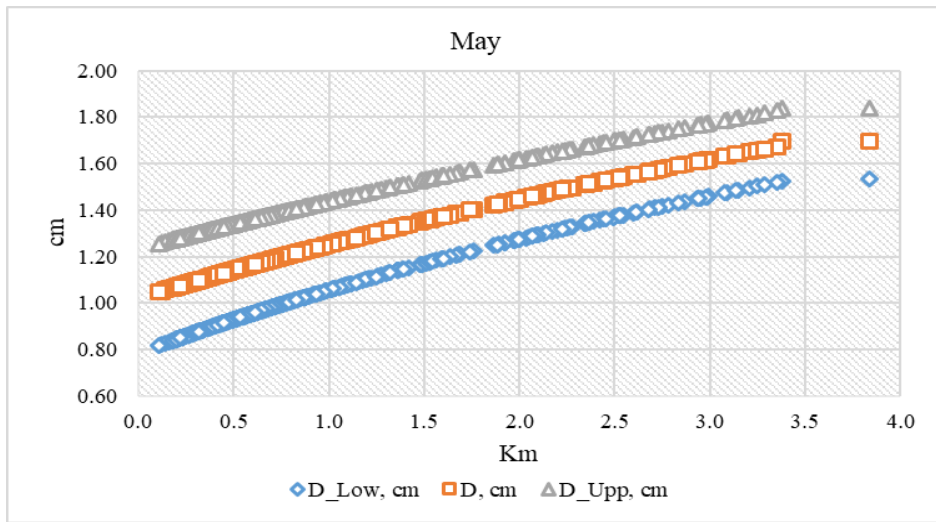


Fig. 8. Vertical distribution of mean max diameter of hailstones and their 99% lower and upper levels in Kakheti in May.

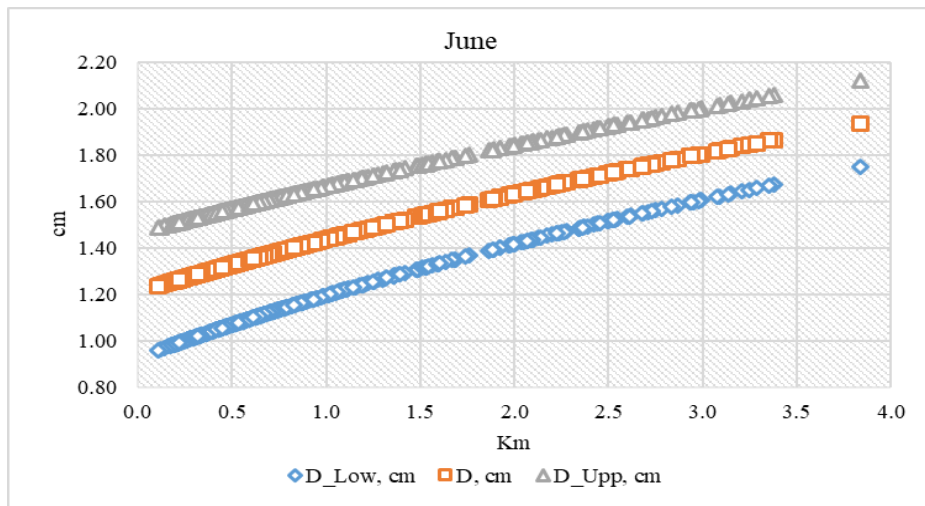


Fig. 9. Vertical distribution of mean max diameter of hailstones and their 99% lower and upper levels in Kakheti in June.

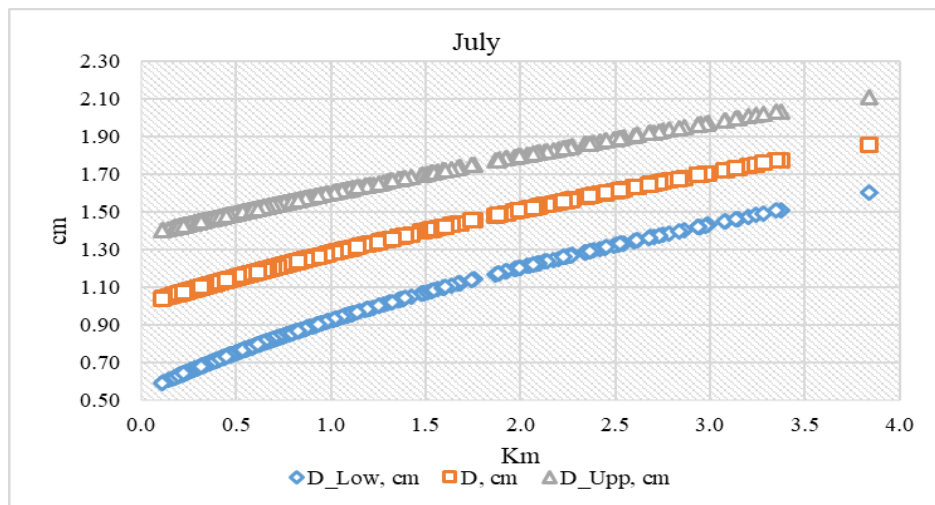


Fig. 10. Vertical distribution of mean max diameter of hailstones and their 99% lower and upper levels in Kakheti in July.

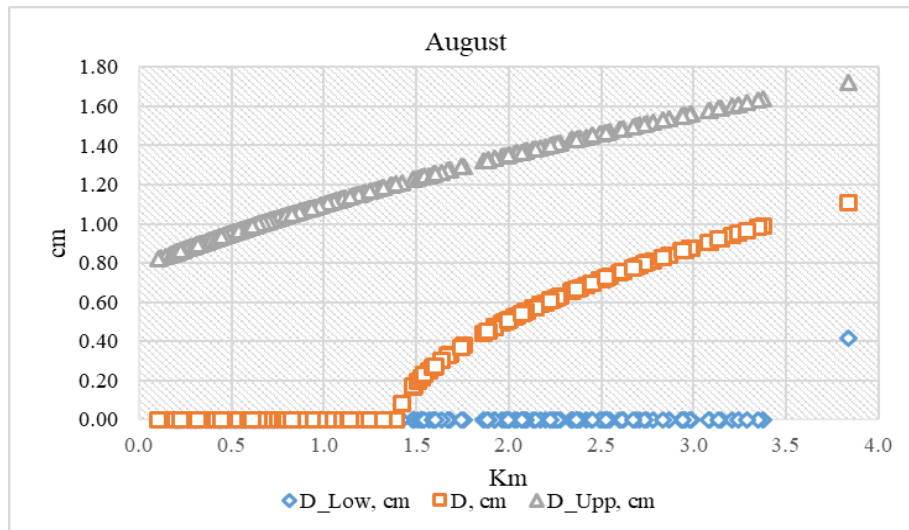


Fig. 11. Vertical distribution of mean max diameter of hailstones and their 99% lower and upper levels in Kakheti in August.

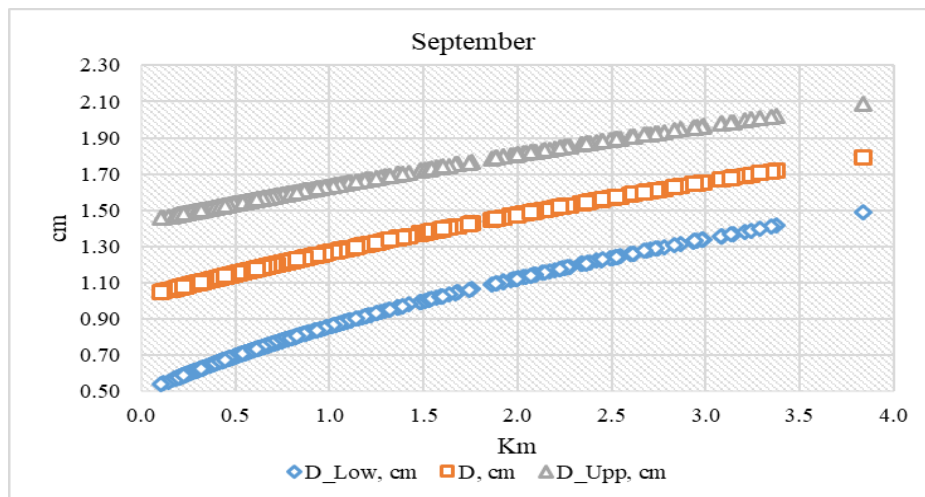


Fig. 12. Vertical distribution of mean max diameter of hailstones and their 99% lower and upper levels in Kakheti in September.

In Fig. 7-12 vertical distribution of mean max diameter of hailstones and their 99% lower and upper levels in Kakheti from April to September are presented. In particular, Fig. 5-12 and Table 2, 3 shows, that the variability of the mean maximum hail diameter on the territory of Kakheti in the range of heights from 0.11 to 3.84 km is as follows:

- **D_99%_Low.** April: 0÷0.91 cm; May: 0.82÷1.53 cm; June: 0.96÷1.75 cm; July: 0.59÷1.60 cm; August: 0÷0.42 cm; September: 0.54÷1.49 cm.
- **D.** April: 0.52÷1.17 cm; May: 1.05÷1.69 cm; June: 1.24÷1.94 cm; July: 1.04÷1.86 cm; August: 0÷1.11 cm; September: 1.05÷1.79 cm.
- **D_99%_Upp.** April: 0.92÷1.42 cm; May: 1.26÷1.84 cm; June: 1.49÷2.12 cm; July: 1.40÷2.11 cm; August: 0.82÷1.72 cm; September: 1.46÷2.09 cm.

Conclusion

In the near future, we plan to study the statistical characteristics of the mean max size of hailstones for the municipalities of Kakheti as well as modeling the damage from hail to vineyards, wheat and corn in the agricultural regions of Kakheti.

References

- [1] Sulakvelidze G.K. Livnevyye osadki i grad. L., Gidrometeoizdat, 1967, 412 s., (in Russian).
- [2] Dessens J., Frale R. Haistone Size Distributions in Southwestern France. *Atmospheric Research*, 33, 1994, pp. 57-73.
- [3] Abshaev A.M., Abshaev M.T., Barekova M.V., Malkarova A.M. Rukovodstvo po organizacii i provedeniu protivogradovih rabot. ISBN 978-5-905770-54-8, Nalchik, Pechatni dvor, 2014, 508 s, (in Russian).
- [4] Elizbarashvili E.Sh., Elizbarashvili M. E. Extreme Weather Events over the Territory of Georgia, 2012, Tbilisi (in Russian).
- [5] Varazanashvili O., Tsereteli N., Amiranashvili A., Tsereteli E., Elizbarashvili E., Dolidze J., Qaldani L., Saluqvadze M., Adamia Sh., Arevadze N., Gventcadze A. Vulnerability, Hazards and Multiple Risk Assessment for Georgia, *Natural Hazards*, Vol. 64, Number 3 (2012), 2021-2056, DOI: 10.1007/s11069-012-0374-3, <http://www.springerlink.com/content/9311p18582143662/fulltext.pdf>.
- [6] Amiranashvili A., Varazanashvili O., Pipia M., Tsereteli N., Elizbarashvili M., Elizbarashvili E. Some Data About Hail Damages in Eastern Georgia and Economic Losses from Them. *Int. Conf. "Advanced Problems in Geophysics"*. Reports, presented on the Scientific Conference "80 years of the M. Nodia Institute of Geophysics". Tb., 2014, pp. 145-150, (in Russian).
- [7] Elizbarashvili E. Sh., Amiranashvili A. G., Varazanashvili O. Sh., Tsereteli N. S., Elizbarashvili M. E., Elizbarashvili Sh. E., Pipia M. G. Hailstorms in the Territory of Georgia. *European Geographical Studies*, ISSN: 2312-0029, vol.2, № 2, 2014, pp. 55-69, DOI: 10.13187/egs.2014.2.55, www.ejournal9.com, (in Russian).
- [8] Gigineishvili V.M. Gradobitia v vostochnoi Gruzii. Leningrad, Gidrometeoizdat, 1960, 123 s.
- [9] Amiranashvili A.G., Nodia A.G., Toronjadze A.F., Khurodze T.V. Some Statistical Characteristics of the Number of Days with Hail into the Warm Half-Year in Georgia in 1941-1990. *Trans. of Institute of Geophysics of Acad. of Sc. of Georgia*, ISSN 1512-1135, v. 58, 2004, pp. 133-141, (in Russian).
- [10] Amiranashvili A., Varazanashvili O., Nodia A., Tsereteli N., Khurodze T. Statistical Characteristics of the Number of Days with Hail Per Annum in Georgia. *Trans. of the Institute of Hydrometeorology*, ISSN 1512-0902, vol. 115, Tb., 2008, pp. 427 – 433, (in Russian).
- [11] Amiranashvili A.G., Nodia A.G., Toronjadze A.F., Khurodze T.V. The Changeability of the Number of Days with the Hail in Georgia in 1941-1990. *Trans. of Institute of Geophysics of Acad. of Sc. of Georgia*, ISSN 1512-1135, v. 58, 2004, pp. 127-132, (in Russian).
- [12] Amiranashvili A.G., Amiranashvili V.A., Nodia A.G., Khurodze T.V., Toronjadze A.F., Bibilashvili T.N. Spatial-Temporary Characteristics of Number of Days with a Hails in the Warm Period of Year in Georgia, *Proc. 14th International Conference on Clouds and Precipitation*, Bologna, Italy, 18-23 July 2004, 2_2_215.1-2_2_215.2.
- [13] Amiranashvili A.G., Bliadze T.G., Jamrishvili N.K., Khurodze T.V., Pipia M.G., Tavidashvili Kh.Z. Comparative Analysis of the Distribution of Number of Days with Hail Per Annum on the Territory of Kakheti According to the Data of the Meteorological Stations and State Insurance Service of Georgia. *Journal of the Georgian Geophysical Society, Issue A. Physics of Solid Earth*, v.20A, 2017, pp. 44 - 56.

- [14] Janelidze I., Pipia M. Hail storms in Georgia in 2016-2018. Int. Sc. Conf. "Natural Disasters in Georgia: Monitoring, Prevention, Mitigation". Proc., ISBN 978-9941-13-899-7, Publish House of Iv. Javakhishvili Tbilisi State University, December 12-14, Tbilisi, 2019, pp. 144 -146.
- [15] Beglarashvili N., Janelidze I., Pipia M., Varamashvili N. Hail Storms in Kakheti (Georgia) in 2014-2018. Int. Sc. Conf. „Modern Problems of Ecology“, Proc., ISSN 1512-1976, v. 7, Tbilisi-Telavi, Georgia, 26-28 September, 2020, pp. 176-179.
- [16] Amiranashvili A., Amiranashvili V., Doreuli R., Khurodze T., Kolesnikov Yu. Some Characteristics of Hail Processes in the Kakheti Region of Georgia, Proc.13th Int. Conf. on Clouds and Precipitation, Reno, Nevada, USA, August 14-18, vol.2, 2000, 1085-1087.
- [17] Abaiadze O., Avlokhashvili Kh., Amiranashvili A., Dzodzuashvili U., Kiria J., Lomtadze J., Osepashvili A., Sauri I., Telia Sh., Khetashvili A., Tskhvedishvili G., Chikhladze V. Radar Providing of Anti-Hail Service in Kakheti. Trans. of Mikheil Nodia Institute of Geophysics, ISSN 1512-1135, vol. 66, Tb., 2016, pp. 28-38, (in Russian).
- [18] Amiranashvili A., Chikhladze V., Dzodzuashvili U., Ghlonti N., Sauri I., Telia Sh., Tsintsadze T. Weather Modification in Georgia: Past, Present, Prospects for Development. International Scientific Conference "Natural Disasters in Georgia: Monitoring, Prevention, Mitigation". Proceedings, ISBN 978-9941-13-899-7, Publish House of Iv. Javakhishvili Tbilisi State University, December 12-14, Tbilisi, 2019, pp. 216-222.
- [19] Amiranashvili A., Chikhladze V., Kveselava N., Kvilitaia N., Sauri I., Shavlakadze Sh. Some Characteristics of Hail Processes in Kakheti (Georgia) According to Radar Observations into 2016-2019. Journal of the Georgian Geophysical Society, ISSN: 1512-1127, Physics of Solid Earth, Atmosphere, Ocean and Space Plasma, v. 23(2), 2020, pp. 50 – 56.
- [20] Bartishvili I. T., Nadibaidze G.A., Begalishvili N.A., Gudushauri Sh. L. K fizicheskim osnovam metoda ZakNIGMI. Tr. ZakNIGMI, vip. 67(73), 1978, s. 73-82, (in Russian).
- [21] Amiranashvili A., Gzirishvili T. Aerosols and Ice Crystals in the Atmosphere. Tbilisi, Metsniereba, 1991, 113 p., (in Russian).
- [22] Amiranashvili A., Bliadze T., Jamrishvili N., Kekenadze E., Tavidashvili Kh., Mitin M. Some Characteristics of Hail Process in Georgia and Azerbaijan on May 28, 2019. Journal of the Georgian Geophysical Society, ISSN: 1512-1127, Physics of Solid Earth, Atmosphere, Ocean and Space Plasma, v. 22(2), 2019, pp. 40–54.
- [23] Amiranashvili A., Bakgsoliani B., Begalishvili N., Beritashvili B., Rekhviashvili R., Tsintsadze T., Chitanava R. On the Necessity of Resumption of Atmospheric Processes Modification Activities in Georgia. Trans. of the Institute of Hydrometeorology, Georgian Technical University, 2013, v. 119, pp.144-152, (in Russian).
- [24] Amiranashvili A., Burnadze A., Dvalishvili K., Gelovani G., Ghlonti N., Dzodzuashvili U., Kaishauri M., Kveselava N., Lomtadze J., Osepashvili A., Sauri I., Telia Sh., Chargazia Kh., Chikhladze V. Renewal Works of Anti-Hail Service in Kakheti. Trans. of Mikheil Nodia institute of Geophysics, ISSN 1512-1135, vol. 66, Tb., 2016, pp. 14 – 27, (in Russian).
- [25] Amiranashvili A., Dzodzuashvili U., Lomtadze J., Sauri I., Chikhladze V. Some Characteristics of Hail Processes in Kakheti. Trans. of Mikheil Nodia Institute of Geophysics, ISSN 1512-1135, vol. 65, Tb., 2015, pp. 77 – 100, (in Russian).
- [26] Amiranashvili A., Nodia A., Khurodze T., Kartvelishvili L., Chumburidze Z., Mkurnalidze I., Chikhradze N. Variability of Number of Hail and Thunderstorm Days in the Regions of Georgia with Active Influence on Atmospheric Processes, Bull. of the Georgian Acad. of Sciences, 172, N3, 2005, 484-486.
- [27] Amiranashvili A., Chikhladze V., Kveselava N., Sauri I. Some Results of Anti-Hail Works in Kakheti into 2016-2019. Int. Sc. Conf. „Modern Problems of Ecology“, Proc., ISSN 1512-1976, v. 7, Tbilisi-Telavi, Georgia, 26-28 September, 2020, pp. 153-156.

- [28] Amiranashvili A., Kveselava N., Kvilitaia N., Sauri I., Shavlakadze Sh., Chikhladze V. Some Results of Anti-Hail Works in Kakheti into 2016-2020. Trans. of M. Nodia Institute of Geophysics, ISSN 1512-1135, vol. LXXII, Tbilisi, 2020, pp. 123-128. (in Georgian).
- [29] Tlisov M.I., Khuchunayev B.M. Issledovaniye prostranstvennogo raspredeleniya zarodyshey grada. Tr. Vses. semin. "Aktivnyye vozdeystviya na gradovyye protsessy i perspektivy usovershenstvovaniya l'dobrazuyushchikh reagentov dlya praktiki aktivnykh vozdeystviy", Nal'chik, 16-21 oktyabrya 1989. M., MO Gidrometeoizdata. 1991. C. 61-74.
- [30] Tlisov M.I., Zagidulin A.A., Khuchunayev B.M., Fedchenko L.M. Apparatura, metodika i rezul'taty nazemnykh issledovaniy fizicheskikh kharakteristik grada. Tr. Vses.konf. "Aktivn. vozd. na gidromet. protsessy". Nal'chik, 22-25 oktyabrya 1991. Sankt-Peterburg. Gidrometeoizdat, kniga 2. 1995. c. 24-30. 29.
- [31] Fraile R., Sa´nchez J.L., de la Madrid J.L., Castro A., Marcos J.L. Some Results from the Hailpad Network in Leo´n (Spain): Noteworthy Correlations among Hailfall Parameters. Theor. Appl. Climatol., 64, 1999, pp. 105–117.
- [32] Sa´nchez J.L., Gil-Robles B., Dessens J., Martin E., Lopez L., Marcos J.L., Berthet C., Ferna´ndez J.T., Garcıa-Ortega E. Characterization of Hailstone Size Spectra in Hailpad Networks in France, Spain, and Argentina. Atmospheric Research Vol. 93, Iss. 1-3, 2009, DOI: 10.1016/j.atmosres.2008.09.033
- [33] Hermida L., Merino A., Sa´nchez J.L., Berthet G., Dessens J., L´opez L., Ferna´ndez-Gonz´alez S., Gasc´on E., Garcıa-Ortega E. Spatial Variability of Hailfalls in France: An Analysis of Air Mass Retro-Trajectories. Conference: European Geosciences Union (EGU) General Assembly, 2014, DOI:10.13140/RG.2.1.2647.3843
- [34] Tsitouridis K. The Greek Hailpad Network 2008 to 2019: Analysis of the Hailpad Data - The Equivalent Hail Diameter. International Journal of Engineering Science Invention (IJESI) ISSN (Online): 2319-6734, ISSN (Print): 2319-6726, Vol. 10, Iss. 9, Series II, 2021, pp. 18-32, DOI: 10.35629/6734-1009021832
- [35] Montopoli M, Picciotti E, Baldini L., Fabio S.Di., Marzano F.S., Vulpiani G. Gazing Inside a Giant-Hail-Bearing Mediterranean Supercell by Dual-Polarization Doppler Weather Radar. Atmospheric Research, Vol. 264, 2021, <https://doi.org/10.1016/j.atmosres.2021.105852>
- [36] Barekova M.V., Gazayeva Z.A., Makitov V.S. Kharakternyye trayektorii superyacheykovykh gradovykh protsessov na Severnom Kavkaze. Tr. VGI, Vyp. 90, 1999, S.80-95, (in Russian).
- [37] Gazayeva Z.A., Makitov V.S. Raspredeleniye trayektoriy peremeshcheniya mnogoyacheykovykh gradovykh protsessov na Severnom Kavkaze. Tr. VGI, Vyp. 80, 1991, S.93-99, (in Russian).
- [38] Selex ES GmbH · Gematronik Weather Radar Systems. Rainbow®5 User Guide, 2015, 464 p., www.gematronik.com
- [39] Zimenkov V.A., Ivanov V.V. Raschet tayaniya gradin v estestvennykh protsessakh. Tr. VGI, 1966, vyp. 3(5).
- [40] Jamrshvili N. K., Tavidashvili Kh. Z. Estimation of the Diameter of Fallen to the Earth's Surface Hail Stones Taking Into Account Their Size in the Cloud and the Heights of Zero Isotherm Under the Conditions of Kakheti Region of Georgia. Int. Sc. Conf. „Modern Problems of Ecology“ Proc., ISSN 1512-1976, v. 6, Kutaisi, Georgia, 21-22 September, 2018, pp. 130-133.
- [41] Jamrshvili N., Tavidashvili Kh. Estimation of the Critical Size of Hailstones in Clouds non Prejudiced to Agriculture in Kakheti. International Scientific Conference “Natural Disasters in Georgia: Monitoring, Prevention, Mitigation”. Proceedings, ISBN 978-9941-13-899-7, Publish House of Iv. Javakhishvili Tbilisi State University, December 12-14, Tbilisi, 2019, pp. 126-129.
- [42] Amiranashvili A.G., Berianidze N.T., Jamrshvili N.K., Tavidashvili Kh.Z. Statistical Characteristics of the Monthly Average Values of the Air Temperature in the Layer of Atmosphere 0.54-27 km above the Kakheti Territory (Georgia) in 2012-2016. Journal of the Georgian Geophysical Society, Issue B. Physics of Atmosphere, Ocean and Space Plasma, v. 20B, 2017, pp.24 –42.

[43] Jamrlishvili N. Monthly and Ten-Day Average Values of Freezing Level in the Atmosphere above Kakheti Territory (Georgia) from April to October. Journal of the Georgian Geophysical Society, ISSN: 1512-1127, Iss. A, Physics of Solid Earth, Tb., 2017, vol. 20A, pp. 57-64.

კახეთის ტერიტორიაზე საშუალო მაქსიმალური ზომების მიხედვით სეტყვის განაწილების მოდელირება ატმოსფეროში გაყინვის დონისა და რადარის გაზომვების მონაცემების გამოყენებით

**ა. ამირანაშვილი, ნ. ბოლაშვილი, ზ. გულაშვილი,
ნ. ჯამრიშვილი, ნ. სუქნიძე, ხ. თავიდაშვილი**

რეზიუმე

წარმოდგენილია სეტყვის განაწილების მოდელირების შედეგები საშუალო მაქსიმალური ზომების მიხედვით კახეთის ტერიტორიისათვის ატმოსფეროში გაყინვის დონისა და ღრუბლებში სეტყვის მაქსიმალური ზომების შესახებ რადიოლოკაციური გაზომვების მონაცემების გამოყენებით.

შედგენილია სეტყვის განაწილების რუკები საშუალო მაქსიმალური ზომების მიხედვით კახეთის ტერიტორიისთვის ცალკეულ თვეებში აპრილიდან სექტემბრამდე. შესწავლილია D - ს ვერტიკალური განაწილება აღნიშნული ტერიტორიისთვის 0.11 - დან 3.84 - კმ- მდე სიმაღლეების დიაპაზონში.

Моделирование распределения по средним максимальных размерам града на территории Кахетии (Грузия) с использованием данных об уровне замерзания в атмосфере и радиолокационных измерений

**А.Г. Амиранашвили, Н. Р. Болашвили, З.М. Гулашвили,
Н. К. Джамришвили, Н. Э. Сукнидзе, Х. З. Тавидашвили**

Резюме

Представлены результаты моделирования распределения градин по средним максимальным размерам на территории Кахетии (Грузия) с использованием данных об уровне промерзания в атмосфере и радиолокационных измерений максимальных размеров града в облаках.

Построены карты распределения града по средним максимальным размерам на территории Кахетии для отдельных месяцев, с апреля по сентябрь. Изучено вертикальное распределение D на указанной территории в диапазоне высот от 0.11 до 3.84 км

Numerical Simulation of Dust Distribution in City Tbilisi Territory in the Winter Period

^{1,2}Aleksandre A. Surmava, ³Leila V. Gverdsiteli, ²Liana N. Intskirveli,
²Natia G. Gigauri

¹M. Nodia Institute of Geophysics at the Iv. Javakhishvili Tbilisi State University,
1, M. Aleksidze Str., 0160, Tbilisi, Georgia, e-mail: aasurmava@yahoo.com

²Institute of Hydrometeorology at the Georgian Technical University, 150-a D.Agmashenebeli Ave, 0112
Tbilisi, Georgia, e-mail: intskirvelebi2@yahoo.com

³Georgian Technical University, 7, Kostava str., Tbilisi, Georgia, e-mail: l.gverdsiteli@gtu.ge

ABSTRACT

Dust propagation at Tbilisi city territory in the winter period during western background light wind is modeled and analyzed using 3D regional model of atmospheric process evolution and via combined integration of admixtures transfer and diffusion equations. The motor transport moving at city streets and highways is the main source of atmosphere pollution. Basic peculiarities featuring dust spatial propagation processes under complex terrain conditions are explored. The role of complex terrain in the passive admixture's diffusion process is studied, urban zones with high dust pollution level are established, and the differences between air pollution spatial distribution in winter and summer seasons are determined. Time intervals, when high dust pollution level of the air is formed or air self-purification process occurs, are defined. Time and spatial change of dust concentration in the lower part of the atmospheric boundary layer is studied. It is obtained that 0.8-1.5 maximum allowable concentration (MAC) is registered at 9 AM and 6 PM at Gldani and Temqa district territories situated in the central and northeastern parts of the city.

Key words: Numerical modeling, pollution source, dust distribution, western background wind, winter season.

Introduction.

Tbilisi is one of the largest cities of the South Caucasus, an administrative and touristic center of Georgia. Tbilisi city relief is highly complex. City is confined from the west and the east by high-mountain massifs; from the north the narrow Mtkvari River gorge, and from the south the lowland territory connects it with the external space. Based on theoretical assumptions we suppose that the local air circulations formed under the influence of the city atmosphere don't promote city atmosphere self-purification process. Indeed, according to atmospheric air monitoring network data and other studies the dust concentration in Tbilisi atmosphere frequently reaches and, in some cases, exceeds maximum allowable concentrations (MAC) [1-7].

In order to elaborate air quality substantiated recommendations and carry out practical measures it is necessary to study theoretically the features of polluting agents time and spatial distribution in city atmosphere, especially during unfavorable meteorological conditions.

In the presented work, with the purpose of further extension of carried out researches [8-11], city atmosphere pollution level in winter period is studied via numerical modeling under conditions of western background light wind. A numerical model is used to describe the development of atmospheric processes and the spread of pollutants in the Caucasus [8].

Brief description of problem statement.

The 30,6x24 sq. km area of Tbilisi and surrounding territories is considered. In order to mathematically correctly describe the dynamic fields of atmosphere and meteorological parameters under conditions of complex terrain of the city a relief-following coordinate system ($t, x, y, \zeta = (z - \delta) / h$) is used. Here t is time,

x and y are coordinates directed along parallel and meridian, ζ is a vertical non dimensional coordinate, $\delta(x, y)$ is a relief height above sea level, $h = H - \delta$ – troposphere thickness, $H(t, x, y)$ – tropopause height.

The equation for dust concentration change in the selected coordinate system will be written in the following form

$$\frac{\partial C}{\partial t} + u \frac{\partial C}{\partial x} + v \frac{\partial C}{\partial y} + (\tilde{w} - \frac{w_0}{h}) \frac{\partial C}{\partial \zeta} = \frac{\partial}{\partial x} \mu \frac{\partial C}{\partial x} + \frac{\partial}{\partial y} \mu \frac{\partial C}{\partial y} + \frac{1}{h^2} \frac{\partial}{\partial \zeta} v \frac{\partial C}{\partial \zeta} + F \quad , \quad (1)$$

where, C – ingredients concentration; u, v, w and \tilde{w} wind velocity components along the x, y, z and ζ axes, w_0 – dust deposition rate, $F(t, x, y, \zeta)$ dust dissipation rate in the atmosphere by the source, μ and v – coefficients of horizontal and vertical turbulence. Wind velocity components and coefficients of turbulence are calculated through numerical integration of equations given in [8] and formulas determining the coefficient of turbulence.

Dust dissipation in the free atmosphere and surface layer of the atmosphere is modeled through numerical integration of equation (1), using respective initial and boundary conditions. Numerical grid steps along the x and y axes equal to 300 and 400 m, and vertical step in the free atmosphere is 1/31 that roughly corresponds to 300 m. In the 100 m thick surface layers of the atmosphere a vertical step varies from 0.5 to 15 m, while time step is 1 sec. Calculations are made for 3-day period. A case of western light background wind under dry weather conditions of January is considered. Wind velocity varies from 1 m/sec (at 100 m height above the ground) to 20 m/sec (in the tropopause at 9 km altitude). Relative atmosphere humidity is 50%.

It is assumed that the atmosphere is polluted by a dust originated at city mains and streets due to motor transport traffic. Its quantity changes in time and is determined according to assessment of continuous surveillance materials and transport traffic intensity.

Numerical modeling results

Spatial distribution of dust concentration and wind velocity, obtained through calculation at 2, 100 and 600 m height from the earth surface, when $t = 3$ and 6 h, is shown in Fig. 1. Concentration is given in the units of one-off maximum allowable concentration (MAC = 0.5 mg/m³).

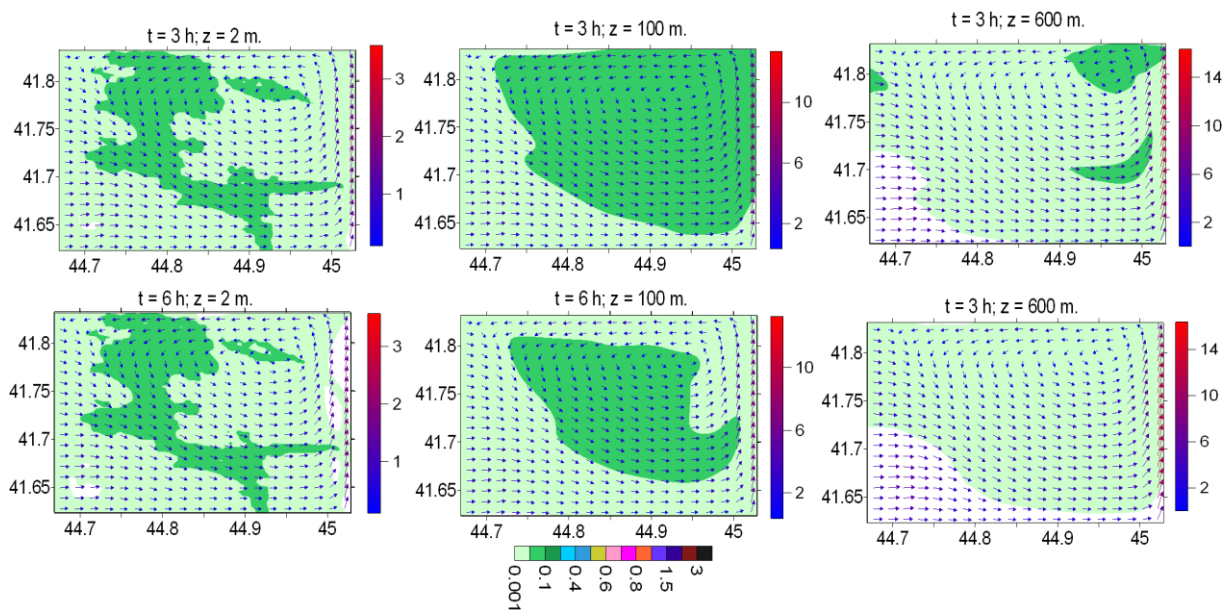


Fig. 1. Wind velocity (m/sec) and dust concentration (MAC) distribution, when $t = 3$ and 6 h, at 2, 100 and 600 m height from the earth surface.

It is seen from Fig. 2 that in time interval from 3h to 6h the dust concentration at 2 m height from the earth surface is practically constant and slightly varies within 0.001-0.1 MAC. 0.01 MAC concentration is obtained at less urbanized territories adjacent to the city. At the central and densely populated (urbanized) territories of the city, dust concentration varies from 0.01 to 0.1 MAC. In the surface layer of the atmosphere dust concentration grows with height increase and at 100 m height it changes within 0.01-0.1MAC in the major part of the atmosphere. Afterwards dust concentration reduces with the further height increase and at 600 m height its value doesn't exceed 0.01 MAC. If we compare spatial distribution and quantitative values of concentration, we come to conclusion that the pollution level in nighttime hours of summer barely exceeds that obtained in winter period.

After $t = 6$ h, the quantity of dust hitting the atmosphere increases along with quick growth of motor transport traffic intensity and rapid pollution of city atmosphere begins. When $t = 9$ h, during first “rush-hour” situation, at 2 m height above the ground dust concentration is high in the proximity of city mains located in the northern part (Gldani, Temka districts) and the central (Vake and Saburtalo districts) parts of the city. Surface distribution of concentration obtained through calculations considerably differs from that obtained in summer period [12]. Maximum concentration value 1.2 MAC is obtained in the surroundings of the crossroad of Sarajishvili, Guramishvili avenues and Kerch street. Maximum value at Vazha-Pshavela and I. Chavchavadze avenues of Vake and Saburtalo districts equals to 0.7 MAC, while highest concentration at A. Tsereteli Avenue of Didube district reaches 0.9 MAC.

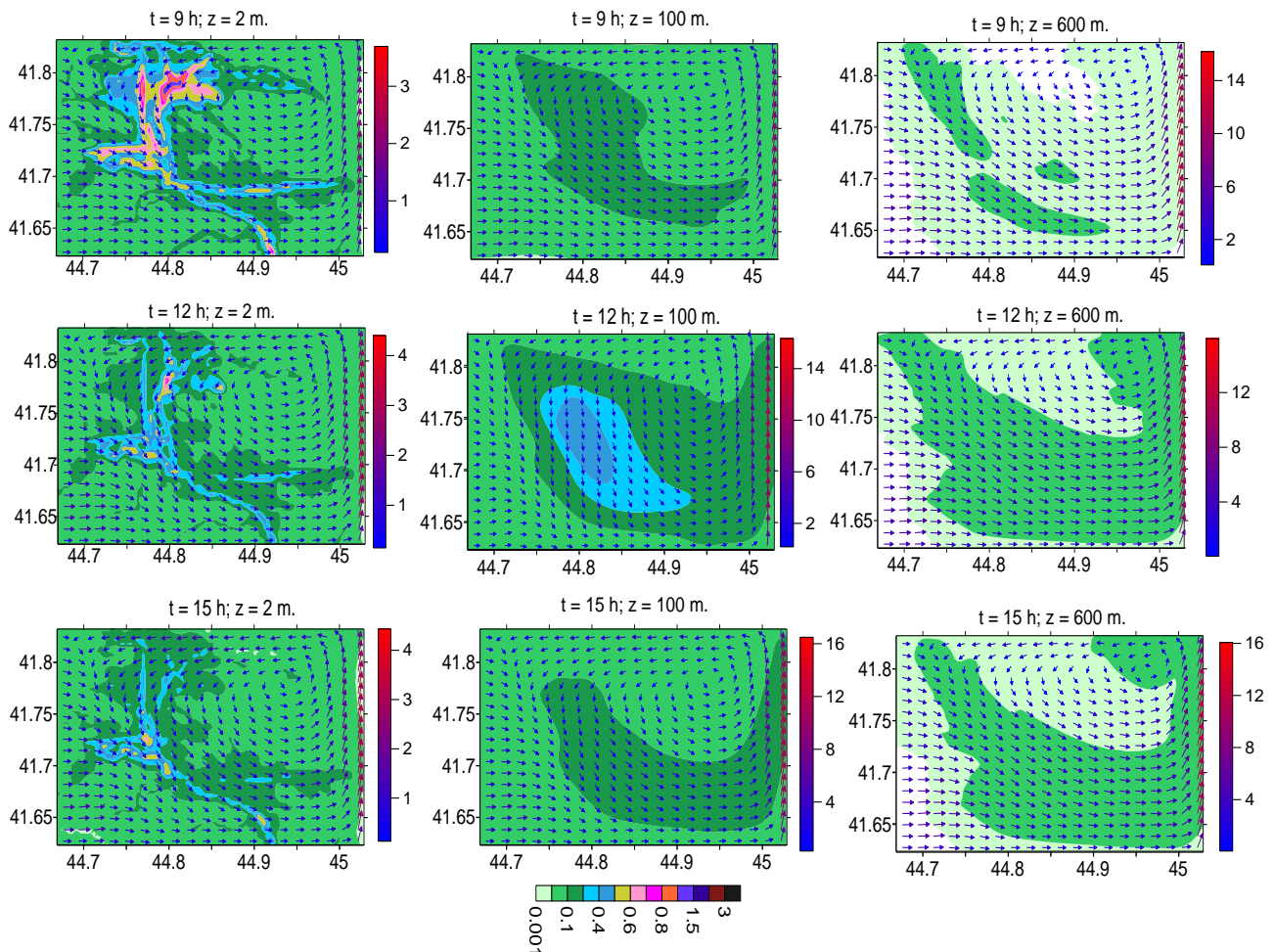


Fig. 2. Wind velocity (m/sec) and dust concentration (MAC) distribution, when $t = 9, 12$ and 15 h, at 2, 100 and 600 m height from the earth surface.

In the time range from 10 AM to midday, under conditions of motor transport traffic constant intensity and dust dissipation constant speed, in the high pollution level area a considerable reduction of concentration

is obtained at 2 m height (Fig. 2). Concentration decrease lasts until 3 PM. When $t = 12$ and 15 h, maximum concentration values vary within 0.6-0.8 MAC interval. The mentioned maximum concentrations are obtained at different urban territories distanced from each other, namely in Temka and Saburtalo districts.

At 100 m from the earth surface a dust is distributed above urbanized territories of the city and its concentration values are equal to 0.3 MAC. The only exception is the concentration values (0.5 MAC) obtained at quite large territories above the central part of the city, when $t = 12$ h. At 600 m height above the ground the concentration doesn't exceed 0.1 MAC.

After 3 PM, at $z = 2$ m height, the second stage of dust pollution level increase begins, which lasts up to 8 PM (Fig. 3). Dust pollution level increases in 1 km thick area of the atmospheric boundary layer. This growth is especially intensive in the lower part of the surface layer in the surroundings of city mains located in the center of the city and its peripheries. In the northern part of Tbilisi (surroundings of Temka and Gldani districts), at 2 m height, dust concentration maximum value reaches and even exceeds 1 MAC, while in the central and southern parts of the city the maximum value equals to 0.8 MAC. At suburban areas, dust concentrations are within 0.2 MAC.

Dust vertical turbulent and convective transfer processes occur along with the increase of its surface concentration. Dust concentration growth in the upper part of the surface layer runs with approximately 3-hour phase lag. As a result, maximum concentration value, which is equal to 0.6 MAC is obtained at 100 m height, when $t = 21$ h. At 600 m height from the earth surface dust pollution level is virtually unchanged and is less than 0.1 MAC.

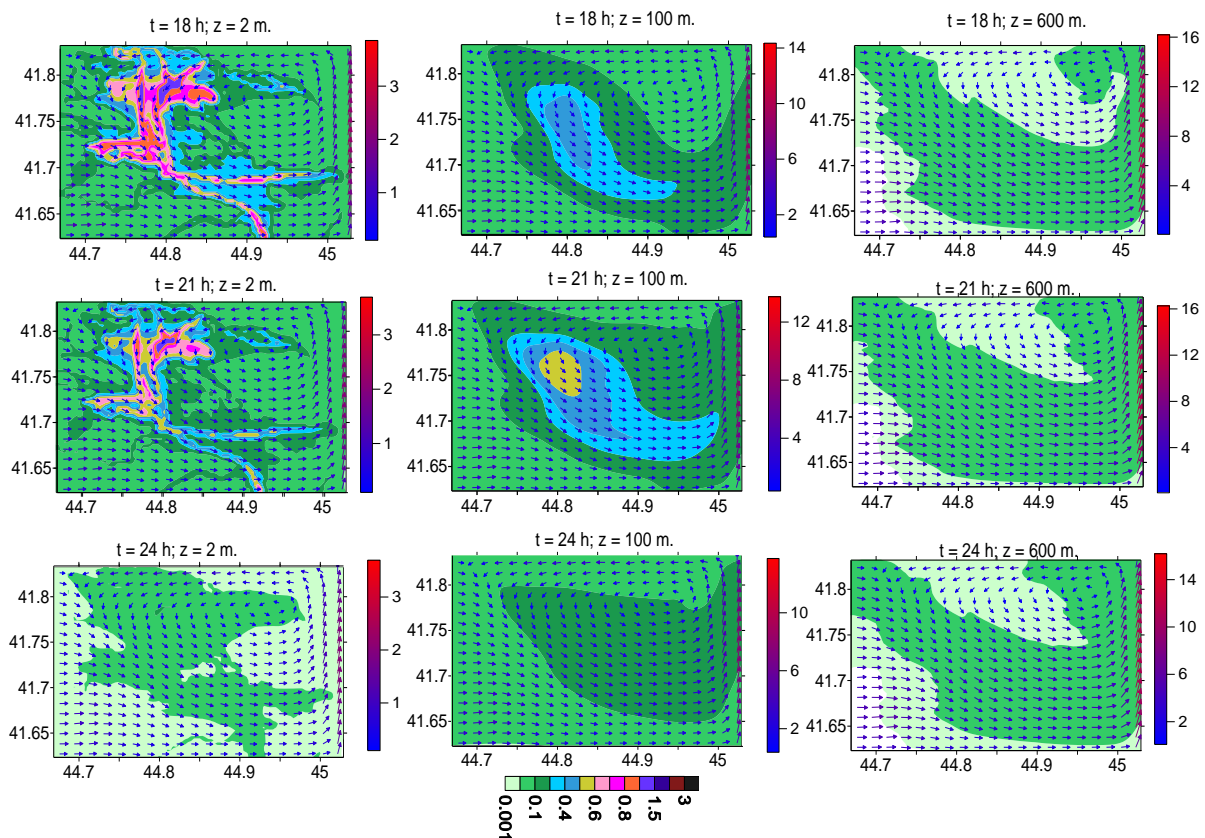


Fig. 3. Wind velocity (m/sec) and dust concentration (MAC) distribution, when $t = 18, 21$ and 24 h, at 2, 100 and 600 m height from the earth surface

After $t = 20$ h the stage of surface concentration reduction begins, which lasts up to $t = 24$ h. Concentration value by this time is within 0.1 MAC at urbanized territory. At 100 m height from the earth surface, dust

concentration roughly equals to 0.2 MAC. Calculations show that in case of constant wind, dust transfer and diffusion process has quasiperiodic character with 24h period.

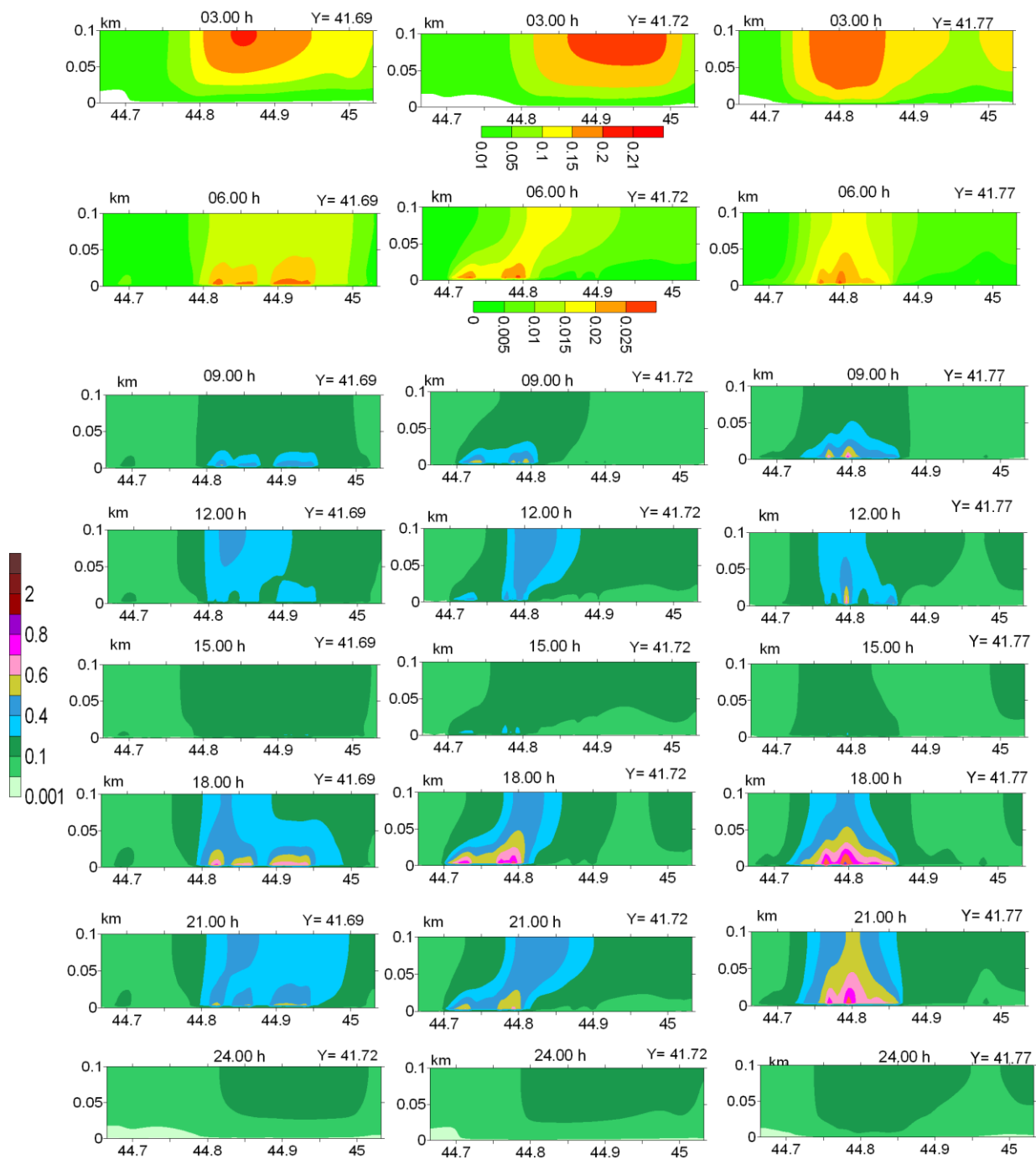


Fig. 4. Dust concentration (MAC) distribution within a day in the surface layer of atmosphere in three vertical planes located along the parallel ($Y= 41.69^\circ$, $Y=41. 72^\circ$ and $Y=41. 77^\circ$).

In Fig. 4 there is shown dust concentration vertical distribution in the surface layer of atmosphere, in three vertical cross-sections drawn along the parallel, with three hour time interval. It is seen from Fig. 4 that from the midnight to 6 AM an atmospheric dust is transferred to the upper part of the surface layer of atmosphere. From $t = 6h$ dust pollution areas start to form near the underlying surface. At following points of time, a surface dust is propagated in vertical and horizontal direction. Intensive vertical dust transfer starts at $t = 18h$ and lasts until $t = 21h$. Dust cloud form points at the convective mechanism of its origination. Prevailing directions of dust propagation are seen in Fig. 4 and from there we can determine respective dynamic processes

causing dust transfer. We can conclude that local circulation processes at the same territory may have different character at different points of time, namely advective, convective or turbulent-diffusive.

Conclusion

Motor transport-induced dust pollution change kinematics in winter period at the territory of Tbilisi is explored in case of western background wind. Diurnal pattern of dust spatial distribution and its propagation peculiarities are studied. Through analysis of wind velocity and concentration fields it is obtained that spatial distribution of heavily dust polluted areas depends on city mains location, on one hand and on dynamic impact of relief, and local circulation systems formed by daily change of thermal regime at underlying surfaces, on the other.

Comparison of dust concentration spatial and time distribution in winter period with that obtained through calculation for summer one, shows not only their similarity, but also points at substantial qualitative difference between them. Distinctions are manifested in different location of high dust pollution areas, maximum dust pollution level occurrence time, and dust vertical and horizontal distribution. Differences between concentration distribution in case of the same background wind can be explained by the seasonal character of local hydrometeorological fields formation. In particular, Tbilisi city relief, in case of western background light wind, generates local cyclonic vortex in winter, and anticyclone vortex in summer. These vortices, due to diurnal change of thermodynamic fields in the surface layer of atmosphere, experience periodic variations of shape, velocity field and location during a day. The mentioned circumstance has an effect on dust transfer-diffusion process and causes differences in spatial distribution of concentration fields between the winter and summer periods.

Acknowledgment. The work is performed with the support of grant project №FR-3667-18 of Shota Rustaveli National Science Foundation of Georgia.

References

- [1] http://air.gov.ge/reports_page.
- [2] Kharchilava J.F., Lomaia O.V., Bukia G.N. The Conditions of Aerosols Formation and Accumulation in Cities, in Proc. 3th Int. Aerosol Conf., Kyoto, Japan, Pergamon, 24-27 September, 1990, v. 2, pp. 986-989.
- [3] Amiranashvili A.G., Gzirishvili T.G. Aerosols and Ice Crystals in the Atmosphere. Tbilisi, Metsniereba, 1991, p. 113 (in Russian).
- [4] Amiranashvili A.G., Chikhladze V.A., Kharchilava J.F., Buachidze N.S, Intskirveli L.N. Variations of the Concentrations of Dust, Nitrogen Oxides, Sulphur Dioxide and Ozone in the Surface Air in Tbilisi in 1981-2003, in Proc. 16th Int. Conf. on Nucleation&Atmospheric Aerosols, Kyoto, Japan, 26-30 July, 2004, pp. 678-681.
- [5] Amiranashvili A., Bliadze T., Tsikhladze V. Photochemical smog in Tbilisi. Monograph. In Proc. of Mikh. Nodia Institute of Geophysics, Part 63, 2012, p. 160 (ISSN 1512-1135).
- [6] Kirkidatze D., Nikiforov G., Chankvetadze A., Chkhaidze G. Some Results of Studies of Atmospheric Aerosols in M. Nodia Institute of Geophysics in the Recent Three Decades. Trans. of Mikheil Nodia Institute of Geophysics (ISSN 1512-1135), v. 66, pp. 178-185, Tbilisi, 2016 (in Russian).
- [7] Intskirveli L., Gigauri N., Surmava A., Kukhalashvili V., Mdivani S. Study of Tbilisi atmospheric air pollution by PM-particles and dust, in Proc. Of Scientific conference “Modern problems of ecology”, Tbilisi-Telavi, Georgia, 26-28 September, 2020, vol. 7, pp. 252-255 (ISSN 1512-1976).
- [8] Surmava A., Intskirveli L.,Kukhalashvili V., Gigauri G. Numerical Investigation of Meso- and Microscale Diffusion of Tbilisi Dust. Annals of Agrarian Science, v. 18, No. 3 2020, pp. 295-302.
- [9] Surmava A., Kukhalashvili V., Gigauri N., Intskirveli L., Kordzakhia G. Numerical Modeling of Dust Propagation in the Atmosphere of a City with Complex Terrain. The Case of Background Eastern Light Air. Journal of Applied Mathematics and Physics, v. 8, No.7, 2020, pp. 1222-1228. <https://doi.org/10.4236/jamp.2020.87092>.
- [10] Kukhalashvili V., Kordzakhia G., Gigauri N., Surmava A., Intskirveli L. Numerical Modelling of Dust Propagation in the Atmosphere of Tbilisi City: The Case of Background Eastern Gentle Breeze. Journal of the Georgian Geophysical Society, v. 23(1), 2020, pp. 46 -50.

[11] Kukhalashvili V., Gigauri N., Surmava A., Demetrashvili D., Intskirveli L. Numerical Modelling of Dust Propagation in the Atmosphere of Tbilisi City: The Case of Background Eastern Fresh Breeze. Journal of the Georgian Geophysical Society, v. 23(1), 2020, pp. 51 -56.

[12] Surmava A., Intskirveli L., Gverdtsiteli L. Numerical modeling of dust distribution in the atmosphere of Tbilisi. I.A case of weak background wind from the west. Proceedings of the M. Nodia Institute of Geophysics, vol. LXXII, 2020, p. 89-96.

ქ. თბილისის ტერიტორიაზე ზამთარში მტვრის გავრცელების რიცხვითი მოდელირება

ა. სურმავა, ლ. გვერდწითელი, ლ. ინჭკირველი, ნ. გიგაური

რეზიუმე

ატმოსფერული პროცესების ევოლუციის 3D რეგიონალური მოდელისა და მინარევების გადატანა - დიფუზიის განტოლების ერთობლივი ინტეგრირებით რიცხობრივად მოდელირებული და გაანალიზებულია მტვრის გავრცელება ქ. თბილისის ტერიტორიაზე ზამთრის პერიოდში დასავლეთის ფონური სუსტი ქარის დროს. ატმოსფეროს დაბინძურების ძირითად წყაროს წარმოადგენს ქალაქის ქუჩებსა და მაგისტრალზე მოძრავი ავტოტრანსპორტი. გამოკვლეულია ძირითადი თავისებურებანი რომლებიც ახასიათებენ რთული რელიეფის პირობებში მტვრის სივრცეში გავრცელების პროცესს. შესწავლილია რთული რელიეფის როლი პასიური მინარევების დიფუზიის პროცესში, დადგენილია ქალაქის მაღალი დამტვერიანების ზონები, აღმოჩენილია განსხვავებები რომლებიც არსებობენ ზამთრისა და ზაფხულის სეზონებში ატმოსფერული ჰაერის დამტვერიანებების სივრცულ განაწილებებს შორის. განსაზღვრულია დროის ინტერვალები, როდესაც ფორმირდება ჰაერის მაღალი დამტვერიანება ან ხდება ჰაერის თვითდასუფთავების პროცესი. შესწავლილია ატმოსფეროს სასაზღვრო ფენის ქვედა ნაწილში მტვრის კონცენტრაციის დროსა და სივრცეში ცვლილება. მიღებული იქნა, რომ 0.8 - 1.5 ზღვრულად დასაშვები კონცენტრაცია (ზღვ) მიიღება დღის 9 და 18 -ზე ქალაქის ცენტრალურ და ჩრდილო -აღმოსავლეთ ნაწილებში მდებარე გლდანისა და თემქის რაიონის ტერიტორიებზე.

Численное Моделирование Распространения Пыли Зимой на территории г. Тбилиси

А. Сурмава, Л. Гвердцители, Л. Инцкирвели, Н. Гигаури

Резюме

Исследовано распространение пыли на территории города Тбилиси в зимний период при фоновом слабом западном ветре с использованием региональной трехмерной модели эволюции атмосферных процессов 3D и интегрирования уравнения переноса-диффузии примесей. В модели автомобильный транспорт рассматривается как основной нестационарный источник загрязнения, от которого в атмосферу выбрасывается пыль. Путем численного моделирования показано, что процесс загрязнения атмосферного воздуха пылью протекает в четыре этапа и зависит от интенсивности автомобильного движения, микрорельефа города и расположения автомобильных магистралей. Исследованы особенности распространения пыли в зимний период года, роль сложного рельефа при диффузии пыли, определены зоны высокого запыленности, выявлены различия между пространственными распределениями пыли в зимний и летный периоды года. Определены интервалы времени, когда формируются зоны высокого запыления воздуха или происходит самоочищения городского воздуха. Получено, что зоны высокой запыленности с концентрациями 0.8-1.5 предельно допустимая концентраций формируются с 9 ч. утра до 6 ч вечера на территориях Глданского и Темкинского районов.

Holiday Climate Index in Kakheti (Georgia)

¹Avtandil G. Amiranashvili, ²Liana G. Kartvelishvili

¹*Mikheil Nodia Institute of Geophysics of Ivane Javakishvili Tbilisi State University, Tbilisi, Georgia
1, M. Alexidze Str., 0160, Tbilisi, Georgia, e-mail: avtandilamiranashvili@gmail.com*

²*National Environmental Agency of Georgia*

ABSTRACT

Data about long-term monthly average values of Holiday Climate Index (HCI) for 12 locations of Kakheti (Akhmeta, Dedoplistskaro, Gombori, Gurjaani, Kvareli, Lagodekhi, Omalo, Sagarejo, Shiraki, Telavi, Tsnori and Udabno) are presented. For 6 stations of this region (Dedoplistskaro, Gurjaani, Kvareli, Lagodekhi, Sagarejo and Telavi) detailed analysis of the monthly, seasonally and annually HCIs values over a 60-year period (1956-2015) are carried out. Comparison of monthly values of HCI and Tourism Climate Index (TCI) for four points of Kakheti (Dedoplistskaro, Kvareli, Sagarejo and Telavi) based on data from 1961 to 2010 are carried out.

Key Words: *Bioclimate, Tourism Climate Index, Holiday Climate Index.*

Introduction

Weather and climate are two factors that in many respects influence on tourism development. Many climate indices for tourism have been applied in past research [1-6]. The most widely known and applied index is the Tourism Climate Index (TCI) proposed by Mieczkowski [7]. In south Caucasus countries, monthly value of TCI be calculated in Georgia, first for Tbilisi [8], then for many other locations of Caucasus (Armenia, Azerbaijan, North Caucasus etc.) [9-16]. For example, the statistical characteristics of the monthly mean, annual and half year values of tourism TCI and its components for four points of Kakheti (Telavi, Dedoplistskaro, Kvareli and Sagarejo) in the period from 1961 through 2010 in [14, 16] are represented. In particular, the changeability of the indicated bioclimatic parameters into 1986÷2010 in comparison with 1961÷1985 is studied, and also the trends of values of TCI for higher enumerated points are investigated.

Despite the TCI's wide application, it has been subject to substantial critiques [17]. The four key deficiencies of the TCI include: (1) the subjective rating and weighting system of climatic variables; (2) it neglects the possibility of an overriding influence of physical climatic parameters (e.g., rain, wind); (3) the low temporal resolution of climate data (i.e., monthly data) has limited relevance for tourist decision-making; and (4) it neglects the varying climatic requirements of major tourism segments and destination types (i.e., beach, urban, winter sports tourism).

To overcome the above noted limitations of the TCI, a Holiday Climate Index (HCI) was developed to more accurately assess the climatic suitability of destinations for tourism. The word "holiday" was chosen to better reflect what the index was designed for (i.e., leisure tourism), since tourism is much broader by definition ("Tourism is a social, cultural and economic phenomenon which entails the movement of people to countries or places outside their usual environment for personal or business/professional purposes" [18-23]).

Results of comparison of the holiday climate index and the tourism climate index in Tbilisi are presented in [24]. Comparison of the values and categories of the Tourism Climate Index and Holiday

Climate Index in Tbilisi shows that the intra-annual variation of both indices is similar and has a bimodal form. However, given that the TCI is calculated for the so-called “average tourist” (regardless of gender, age, physical condition), the value and category of this index is lower than the HCI values and categories. In general, HCI more adequately determines the bioclimatic state of the environment for the development of various types of tourism than TCI [24]. The detailed information on the variability of the monthly values of the Holiday Climate Index (HCI) in Tbilisi in 1956-2015 in [25] are presented. It also presents data on the interval forecast of variability of HCI values in Tbilisi for the next few decades.

This study develops a long-term average of HCI for 12 stations of Kakheti region of Georgia (Akhmeta, Dedoplistskaro, Gombori, Gurjaani, Kvareli, Lagodekhi, Omalo, Sagarejo, Shiraki, Telavi, Tsnori and Udabno), detailed analysis of the monthly, seasonally and annually HCIs values over a 60-year period (1956-2015) for 6 stations of this region (Dedoplistskaro, Gurjaani, Kvareli, Lagodekhi, Sagarejo and Telavi) and comparison of monthly values of HCI and TCI for four points of Kakheti (Dedoplistskaro, Kvareli, Sagarejo and Telavi) based on data from 1961 to 2010.

Study Area, Material and Methods

Study area - Kakheti region of Georgia (below - Kakheti). Kakheti is located in the eastern part of Georgia. Area - 11375 km², population - 314.7 thous. pers., (including of urban - 71.4 thous. pers.), the capital of region - Telavi (population - 19.8 thous. pers.) [www.geostat.ge].

A visit to Kakheti can be a fascinating experience because of its beautiful mountain landscapes, stunning regions, ancient world temples and monasteries, picturesque valleys and rivers and home to amber grapes that grows under the warmth of the sun. Kakheti is not only famous as a tourism destination, but it is also locally recognized as Georgia’s center for winemaking.

Studies for 12 locations of Kakheti (Akhmeta, Dedoplistskaro, Gombori, Gurjaani, Kvareli, Lagodekhi, Omalo, Sagarejo, Shiraki, Telavi, Tsnori and Udabno) are carried out.

Fig. 1 depicts the map of the arrangement of the indicated meteorological stations. Table 1 presents information about coordinates and heights of these 12 meteorological stations, whose data were used in the work. These stations that are located from 223 to 1880 meters above sea level and are open to fresh and pure air because of this.

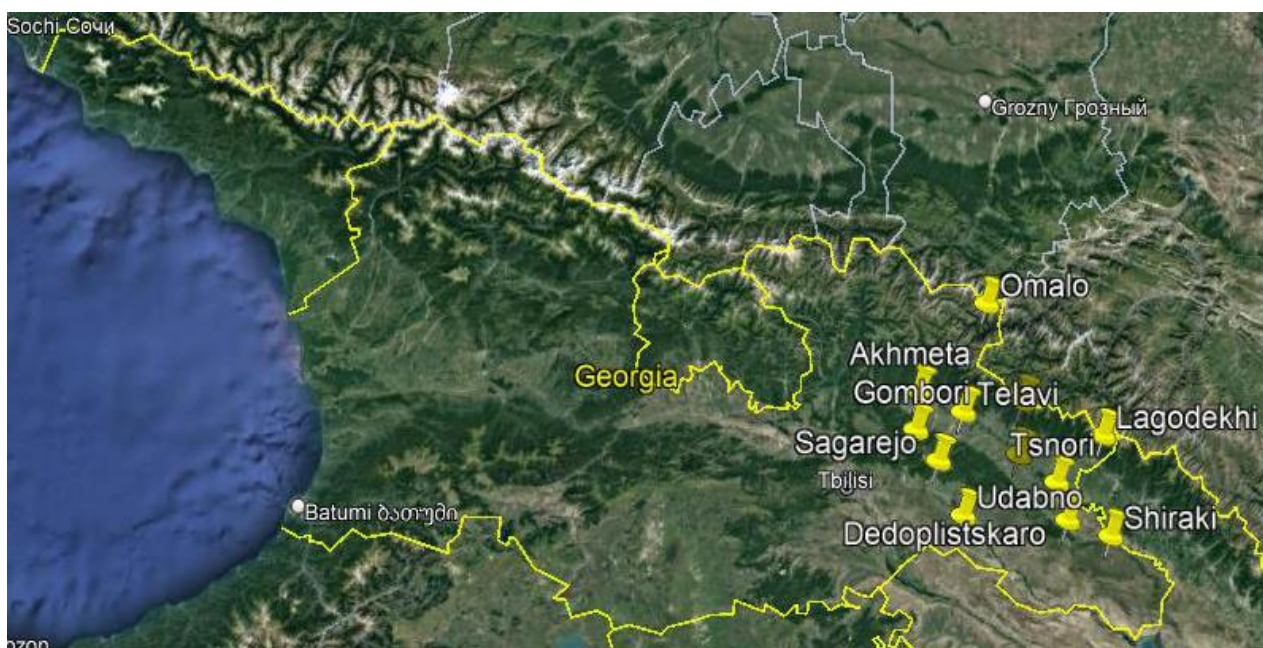


Fig.1. Locations of 12 meteorological stations in Kakheti.

Table 1. Coordinates and heights of the 12 meteorological stations in Kakheti.

Location (Abbreviation)	Latitude, N°	Longitude, E°	Height, m, a.s.l.
Akhmeta (Akhm)	42.02	45.22	567
Dedoplistskaro (Ded)	41.47	46.08	800
Gombori (Gom)	41.86	45.20	1085
Gurjaani (Gur)	41.75	45.80	410
Kvareli (Kvar)	41.97	45.83	449
Lagodekhi (Lagod)	41.82	46.30	362
Omalo	42.38	45.63	1880
Sagarejo (Sag)	41.73	45.33	802
Shiraki (Shir)	41.40	46.33	555
Telavi (Tel)	41.93	45.48	568
Tsnori	41.63	46.02	223
Udabno (Udab)	41.50	45.47	750

In this work the Holiday Climate Index (HCI) is used. The HCI uses five climatic variables related to the three facets essential to tourism (table 2): thermal comfort (TC), aesthetic (A), and physical (P) facet. The five climatic variables used for the HCI input are maximum air temperature and relative humidity (TC), cloud cover (A), precipitation and wind (P) [12].

The HCI score is calculated according to the following formula: $HCI = 4 \cdot T + 2 \cdot A + 3 \cdot R_d + 1 \cdot W$. In tables 2-4 components of Holiday Climate Index, HCI's rating scheme and HCI's category are presented.

Table 2. Components of Holiday Climate Index (HCI).

Facet	Climatic Variable	Index Weighting (%)
Thermal Comfort (TC)	Dry-bulb Temperature (°C): Maximum Temperature (°C)	40%
	Relative Humidity (%): Mean RH	
Aesthetic (A)	Cloud Cover (%)	20%
Physical (P)	Amount of Rain (mm)	30%
	Wind Speed (km/h)	10%

Table 3. HCI's Rating Scheme.

Rating	T - Effective Temperature (°C)	A - Daily Cloud Cover (%)	R _d - Daily Precipitation (mm)	W - Wind Speed (km/h)
10	23÷25	11÷20	0	1÷9
9	20÷22; 26	1÷10; 21÷30	<3	10÷19
8	27÷28	0; 31÷40	3÷5.99	0; 20÷29
7	18÷19; 29÷30	41÷50		
6	15÷17; 31÷32	51÷60		30÷39
5	11÷14; 33÷34	61÷70	6÷8.99	
4	7÷10; 35÷36	71÷80		
3	0÷6	81÷90		40÷49
2	-5÷-1; 37÷39	90÷99	9÷12	
1	<-5	100		
0	>39		>12	50÷70
-1			>25	
-10				>70

Table 4. HCI's Category.

HCI Score	Category (Abbreviation)	HCI Score	Category (Abbreviation)
90÷100	Ideal	40÷49	Marginal (Marg.)
80÷89	Excellent (Excell.)	30÷39	Unfavorable
70÷79	Very Good	20÷29	Very Unfavorable
60÷69	Good	10÷19	Extremely Unfavorable
50÷59	Acceptable (Accept.)	9÷-9; -10÷-20	Impossible

For the 12 indicated localities the long-term monthly average values of HCI with the use data of Georgian National Environmental Agency are calculated. For 6 stations of this region (Dedoplistskaro, Gurjaani, Kvareli, Lagodekhi, Sagarejo and Telavi) detailed analysis of the monthly, seasonally and annually HCIs values over a 60-year period (1956-2015) are carried out. Comparison of monthly HCI and TCI values for four points of Kakheti (Dedoplistskaro, Kvareli, Sagarejo and Telavi) based on data from 1961-2010 are carried out [14,16].

In the work analysis of data is carried out with the use of the standard statistical analysis methods. The following designations will be used below: Mean – average values; Min – minimal values; Max - maximal values; 99%_Low and 99%_Upp - Low and Upper levels of 99% confidence interval of mean values; R² - coefficient of determination; R - coefficient of linear correlation.

Results and discussion

Results in the Table 4-18 and Fig. 2-17 are presented.

1. Basic information about HCI for 12 points of Kakheti.

Data about long-term mean of HCI real values at 12 locations of Kakheti in Fig. 2 are presented.

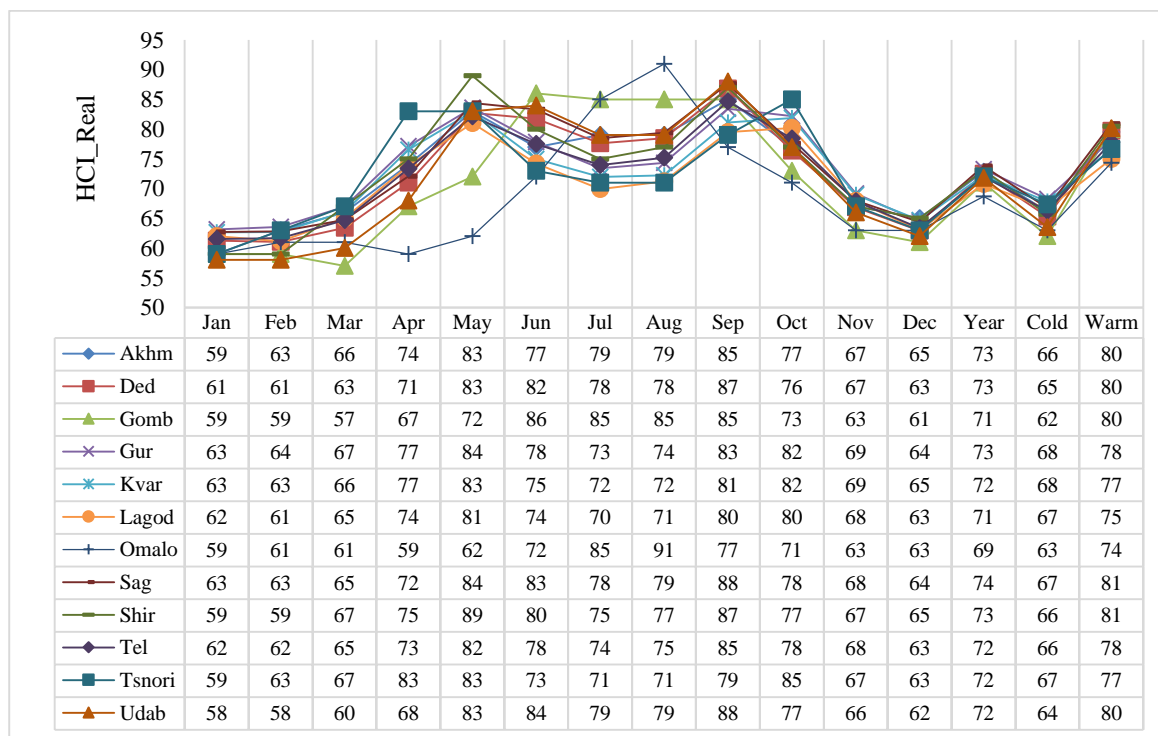


Fig. 2. Mean of HCI real values at 12 locations of Kakheti.

As follows from Fig. 2 mean monthly values of HCI change from 57 (Gombori, March, Acceptable) to 91 (Omalo, August, Ideal). The variability of HCI values for individual items is as follows:

Akhmeta (59, January – 85, September), Dedoplistskaro (61, January, February – 87, September), Gombori (57, March – 86, June), Gurjaani (63, January – 84, May), Kvareli (63, January, February – 83, May), Lagodekhi (61, February – 81, May), Omalo (59, January, April – 91, August), Sagarejo (63, January, February – 88, September), Shiraki (59, January, February – 89, May), Telavi (62, January, February – 85, September), Tsnori (59, January – 85, October, Udabno (58, January, February – 88, September).

Table 5. Linear correlation coefficient between separated stations on the mean monthly values of HCI (R min = 0.60, $\alpha = 0.05$; R min = 0.56, $\alpha = 0.075$; R min = 0.52, $\alpha = 0.10$; R min = 0.47, $\alpha = 0.15$).

Location	Akhm	Ded	Gomb	Gur	Kvar	Lagod	Omalo	Sag	Shir	Tel	Tsnori	Udab
Akhm	1	0.97	0.87	0.92	0.89	0.90	0.64	0.97	0.97	0.97	0.82	0.96
Ded	0.97	1	0.91	0.91	0.87	0.89	0.64	1.00	0.96	0.98	0.75	1.00
Gomb	0.87	0.91	1	0.71	0.64	0.66	0.85	0.90	0.78	0.82	0.52	0.93
Gur	0.92	0.91	0.71	1	0.99	0.99	0.38	0.92	0.95	0.97	0.94	0.89
Kvar	0.89	0.87	0.64	0.99	1	1.00	0.32	0.88	0.93	0.95	0.96	0.85
Lagod	0.90	0.89	0.66	0.99	1.00	1	0.35	0.90	0.94	0.96	0.94	0.87
Omalo	0.64	0.64	0.85	0.38	0.32	0.35	1	0.62	0.48	0.53	0.22	0.67
Sag	0.97	1.00	0.90	0.92	0.88	0.90	0.62	1	0.96	0.98	0.76	0.99
Shir	0.97	0.96	0.78	0.95	0.93	0.94	0.48	0.96	1	0.98	0.84	0.94
Tel	0.97	0.98	0.82	0.97	0.95	0.96	0.53	0.98	0.98	1	0.85	0.97
Tsnori	0.82	0.75	0.52	0.94	0.96	0.94	0.22	0.76	0.84	0.85	1	0.72
Udab	0.96	1.00	0.93	0.89	0.85	0.87	0.67	0.99	0.94	0.97	0.72	1

Linear correlation coefficient between separated station on the mean monthly values of HCI changes as follows (Table 5). Akhmeta: 0.64 (Omalo) - 0.97 (Dedoplistskaro, Sagarejo, Shiraki, Telavi); Dedoplistskaro: 0.64 (Omalo) – 1.00 (Sagarejo, Udabno); Gombori: 0.52 (Tsnori) – 0.93 (Udabno); Gurjaani: 0.38 (non sign, Omalo) – 0.99 (Kvareli, Lagodekhi); Kvareli: 0.32 (non sign, Omalo) – 1.00 (Lagodekhi); Lagodekhi: 0.35 (non sign, Omalo) – 1.00 (Kvareli); Omalo: 0.22 (non sign, Tsnori) – 0.85 (Gombori); Sagarejo: 0.62 (Omalo) – 1.00 (Dedoplistskaro); Shiraki: 0.48 (Omalo) – 0.98 (Telavi); Telavi: 0.53 (Omalo) – 0.98 (Dedoplistskaro, Sagarejo, Shiraki); Tsnori: 0.22 (non sign, Omalo) – 0.96 (Kvareli); Udabno: 0.67 (Omalo) – 1.00 (Dedoplistskaro).

Distribution of mean monthly values of TCI for 12 locations of Kakhети by a ninth power of polynomial ($R^2 \geq 0.978$) are described. Coefficients of the equation of the regression of the intra-annual motion of mean monthly values of HCI for these points in Table 6 are presented.

Table 6. Coefficients of the equation of the regression of the intra-annual motion of mean monthly values of HCI for 12 points of Kakhети.

Equation of regress., coefficients	$HCI = a \cdot X^9 + b \cdot X^8 + c \cdot X^7 + d \cdot X^6 + e \cdot X^5 + f \cdot X^4 + g \cdot X^3 + h \cdot X^2 + i \cdot X + j, (X\text{-Month})$										
	a	b	c	d	e	f	g	h	i	j	R ²
Akhm	-7.05E-05	3.81E-03	-8.54E-02	1.02E+00	-7.08E+00	2.82E+01	-6.12E+01	6.11E+01	-9.50E+00	4.65E+01	0.978
Ded	-2.35E-04	1.35E-02	-3.26E-01	4.37E+00	-3.52E+01	1.76E+02	-5.40E+02	9.72E+02	-9.25E+02	4.08E+02	0.996
Gom	-6.56E-05	3.78E-03	-9.19E-02	1.23E+00	-9.81E+00	4.79E+01	-1.41E+02	2.38E+02	-2.10E+02	1.33E+02	0.979
Gur	4.62E-05	2.70E-03	6.75E-02	9.38E-01	7.94E+00	4.22E+01	1.39E+02	2.69E+02	2.74E+02	1.09E+02	0.996
Kvar	2.38E-06	-2.91E-04	1.31E-02	-2.93E-01	3.67E+00	-2.67E+01	1.12E+02	-2.56E+02	2.88E+02	-5.82E+01	0.990
Lagod	-4.37E-05	2.44E-03	-5.61E-02	6.80E-01	-4.68E+00	1.81E+01	-3.67E+01	3.36E+01	-7.02E+00	5.81E+01	0.993
Omalo	1.57E-04	-9.07E-03	2.22E-01	-3.00E+00	2.44E+01	-1.23E+02	3.82E+02	-6.99E+02	6.78E+02	-2.00E+02	0.990
Sag	4.92E-05	2.88E-03	7.19E-02	9.99E-01	8.46E+00	4.49E+01	1.48E+02	2.87E+02	2.92E+02	1.16E+02	0.997
Shir	1.22E-04	7.16E-03	1.79E-01	2.49E+00	2.11E+01	1.12E+02	3.68E+02	7.14E+02	7.27E+02	2.89E+02	0.984
Tel	-1.46E-04	8.30E-03	-1.98E-01	2.59E+00	-2.02E+01	9.72E+01	-2.85E+02	4.89E+02	-4.45E+02	2.23E+02	0.993
Tsnori	2.22E-04	-1.31E-02	3.31E-01	-4.69E+00	4.06E+01	-2.20E+02	7.40E+02	-1.45E+03	1.49E+03	-5.31E+02	0.996
Udab	-2.79E-04	1.61E-02	-3.93E-01	5.30E+00	-4.31E+01	2.17E+02	-6.71E+02	1.22E+03	-1.16E+03	4.95E+02	0.997

In Table 7 information about distribution types of mean monthly values of HCI at 12 locations of Kakheti are provided.

As follows from this Table in general bimodal distribution type of HCI is observed (10 locations from 12). For stations Akhmeta, Dedoplistskaro, Sagarejo, Shiraki and Telavi the first and second extremum in HCI distribution fall on May and September respectively; for stations Gurjaani, Kvareli and Lagodekhi - on May and September-October; for Tsnori - on April-May and October and for Udabno - on May-June and September.

For Gombori unimodal distribution type of HCI with plateau in June-September is observed; for Omalo - unimodal distribution type with maximum in August.

Table 7. Distribution types of mean monthly values of HCI at 12 locations of Kakheti.

Location	Distribution type	First extremum	Second extremum	Location	Distribution type	First extremum	Second extremum
Akhmeta	Bimodal	May	Sep	Omalo	Unimodal	Aug	
Dedoplistskaro				Sagarejo	Bimodal	May	Sep
Gombori	Unimodal, plateau	June-Sept		Shiraki			
Gurjaani	Bimodal	May	Sep-Oct	Telavi			
Kvareli				Tsnori			
Lagodekhi				Udabno	May-Jun	Sep	

In Table 8 and 9 data about categories of mean monthly and seasonal values of HCI at 12 locations of Kakheti in cold and warm period are presented.

Table 8. Categories of mean monthly and seasonal values of HCI at 12 locations of Kakheti in cold period.

Location	Jan	Feb	Mar	Oct	Nov	Dec	Cold	Year
Akhmeta	Accept.	Good	Good	Very Good	Good	Good	Good	Very Good
Dedoplistskaro	Good							
Gombori	Accept.	Accept.	Accept.					
Gurjaani	Good	Good	Good	Excell.				
Kvareli								
Lagodekhi								
Omalo	Accept.	Good	Good	Very Good				
Sagarejo	Good							
Shiraki	Accept.			Accept.				
Telavi	Good	Good	Excell.					
Tsnori	Accept.	Accept.	Very Good					
Udabno								

Table 9. Categories of mean monthly and seasonal values of HCI at 12 locations of Kakheti in warm period.

Location	Apr	May	Jun	Jul	Aug	Sep	Warm
Akhmeta	Very Good	Excell.	Very Good	Very Good	Very Good	Excell.	Excell.
Dedoplistskaro			Excell.				
Gombori	Good	Very Good	Very Good	Excell.	Excell.		
Gurjaani	Very Good	Excell.		Very Good	Very Good	Very Good	
Kvareli							
Lagodekhi	Accept.	Good	Very Good	Excell.	Ideal	Very Good	
Omalo							
Sagarejo	Very Good	Excell.	Excell.	Very Good	Very Good	Excell.	
Shiraki							
Telavi							
Tsnori	Excell.	Excell.	Very Good	Very Good	Very Good	Very Good	
Udabno	Good		Excell.				

As follows from these Tables, categories of mean monthly and seasonal values of HCI at 12 locations of Kakheti changes from Acceptable to Ideal.

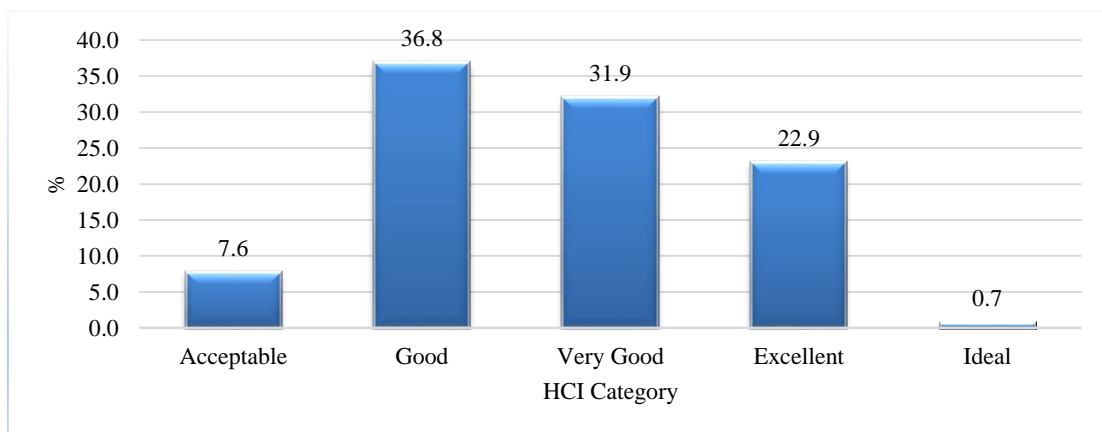


Fig. 3. Repetition of mean monthly HCI category at 12 locations of Kakheti.

In Fig. 3 information about repetition of mean monthly HCI category at 12 locations of Kakheti is presented. So, as follows from Tables 8, 9 and Fig. 3 in Kakheti there are favorable conditions for the development of tourism and resorts throughout the year.

2. Analysis of HCI and HCI components at 6 points of Kakheti in 1956-2015.

Detailed analysis of HCI and HCI components at 6 points of Kakheti from 1956 to 2015 in Fig. 4-15 and Tables 10-15 are presented.

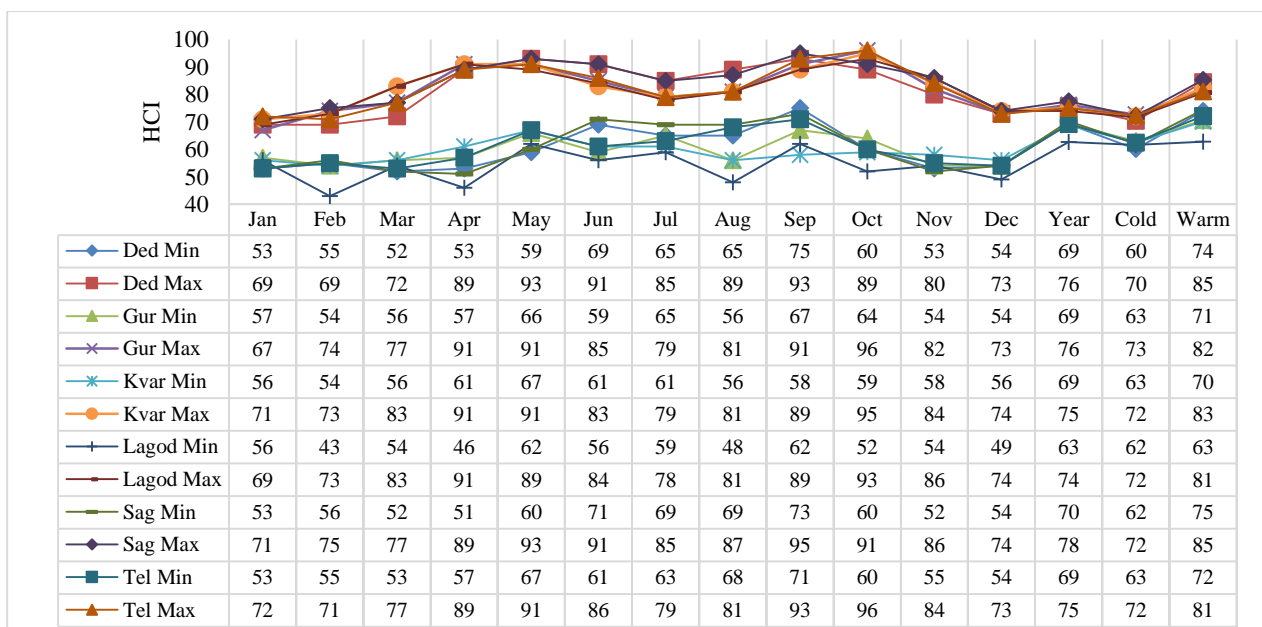


Fig. 4. Min and Max values of HCI at 6 locations of Kakheti in different months and season in 1956-2015.

In Fig. 4 and Tables 10,11 data about Min and Max values of HCI and these categories at 6 locations of Kakheti in different months and season in 1956-2015 are presented. As follows from these Fig. and Tables values of HCI in 6 locations of Kakheti changes from 43 (Lagodekhi, February, Marginal) to 96 (Telavi, October, Ideal).

Table 10. Categories of Min and Max values of HCI at 6 locations of Kakheti in cold period in 1956-2015.

Location	Parameter	Jan	Feb	Mar	Oct	Nov	Dec	Cold	Year	
Ded	Min	Accept. Good	Accept.	Accept. Very Good	Good	Accept. Excell.	Accept. Very Good	Good Very Good	Good Very Good	
Ded	Max		Good		Excell.					
Gur	Min		Accept. Very Good		Marg.					Good
Gur	Max									Ideal
Kvar	Min	Accept.	Marg.	Accept. Excell.	Accept. Ideal	Accept. Excell.	Marg.	Very Good		
Kvar	Max	Very Good							Very Good	
Lagod	Min	Accept.	Marg.	Accept. Excell.	Accept. Ideal	Accept. Excell.	Marg.	Very Good		
Lagod	Max	Good							Very Good	
Sag	Min	Accept. Very Good	Accept. Very Good	Accept. Very Good	Good Ideal	Accept. Excell.	Accept. Very Good	Good Very Good	Very Good	
Sag	Max								Good	
Tel	Min								Good	
Tel	Max								Very Good	

Table 11. Category of Min and Max values of HCI at 6 locations of Kakheti in warm period in 1956-2015.

Location	Parameter	Apr	May	Jun	Jul	Aug	Sep	Warm
Ded	Min	Accept.	Accept.	Good	Good	Good	Very Good	Very Good Excell.
Ded	Max	Excell.	Ideal	Ideal	Excell.	Excell.	Ideal	
Gur	Min	Accept.	Good Ideal	Accept.	Good Very Good	Accept. Excell.	Good	
Gur	Max	Ideal		Excell.			Ideal	
Kvar	Min	Good		Good			Accept.	
Kvar	Max	Ideal		Excell.			Excell.	
Lagod	Min	Marg.	Good	Accept.	Accept.	Marg.	Good	Good
Lagod	Max	Ideal	Excell.	Excell.	Very Good	Excell.	Excell.	Excell.
Sag	Min	Accept. Excell.	Good Ideal	Very Good	Good	Good Excell.	Very Good Ideal	Very Good Excell.
Sag	Max			Ideal	Excell.			
Tel	Min			Good	Good			
Tel	Max			Excell.	Very Good			

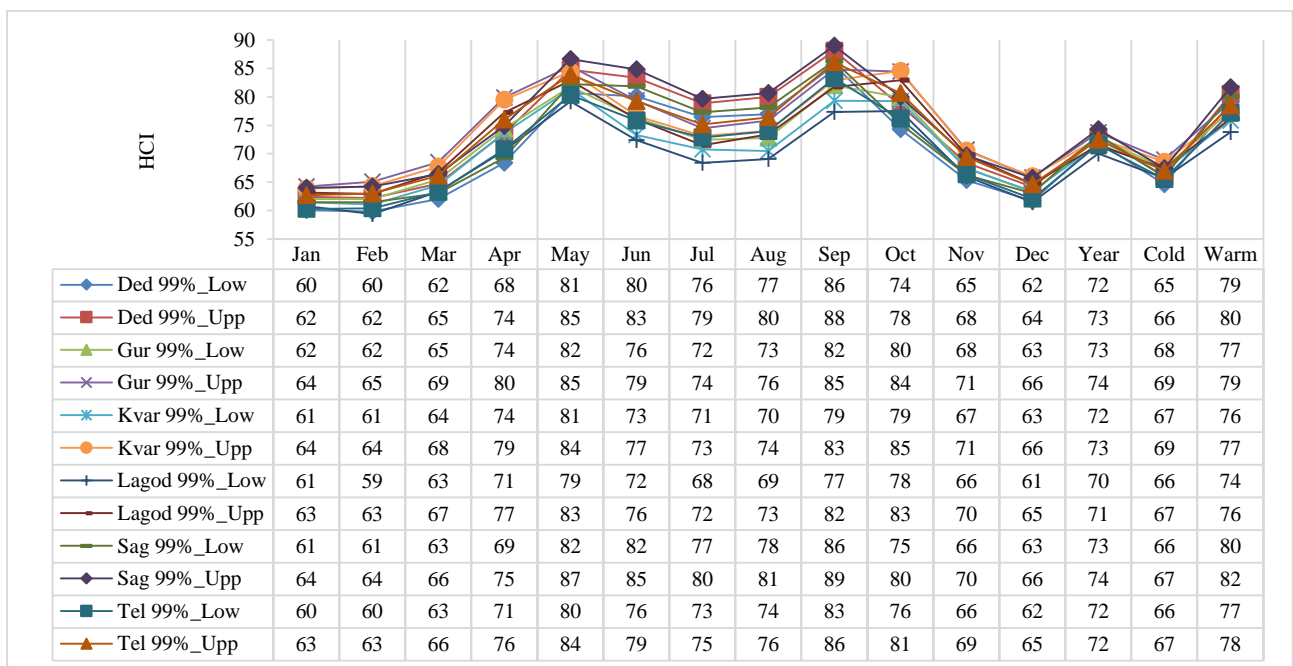


Fig. 5. Low and Upper levels of 99% confidence interval of mean values of HCI at six locations of Kakheti in 1956-2015.

Table 12. Category of Low and Upper levels of 99% confidence interval of mean values of HCI at 6 locations of Kakheti in cold period in 1956-2015.

Location	Parameter	Jan	Feb	Mar	Oct	Nov	Dec	Cold	Year	
Ded	99%_Low	Good	Good	Good	Very Good	Good	Good	Good	Very Good	
Ded	99%_Upp									
Gur	99%_Low				Excell.	Very Good				Good
Gur	99%_Upp									
Kvar	99%_Low				Very Good	Excell.				Good
Kvar	99%_Upp									
Lagod	99%_Low				Accept.	Very Good				Excell.
Lagod	99%_Upp				Good					
Sag	99%_Low				Good	Very Good				Excell.
Sag	99%_Upp									
Tel	99%_Low				Very Good	Excell.				Good
Tel	99%_Upp									

Table 13. Category of Low and Upper levels of 99% confidence interval of mean values of HCI at 6 locations of Kakheti in warm period in 1956-2015.

Location	Parameter	Apr	May	Jun	Jul	Aug	Sep	Warm
Ded	99%_Low	Good	Excell.	Excell.	Very Good	Very Good	Excell.	Very Good
Ded	99%_Upp	Very Good						
Gur	99%_Low	Very Good		Very Good	Good	Good	Very Good	Very Good
Gur	99%_Upp	Excell.						
Kvar	99%_Low	Very Good		Very Good	Good	Very Good	Excell.	Very Good
Kvar	99%_Upp							
Lagod	99%_Low	Very Good		Excell.	Very Good	Very Good	Excell.	Very Good
Lagod	99%_Upp							
Sag	99%_Low	Good		Excell.	Excell.	Excell.	Excell.	Excell.
Sag	99%_Upp	Very Good						
Tel	99%_Low	Very Good		Very Good	Very Good	Very Good	Very Good	Very Good
Tel	99%_Upp							

In Fig. 5 and Tables 12,13 information about 99%_Low and 99%_Upp levels of mean values of HCI and these categories at 6 locations of Kakheti in different months and season in 1956-2015 are presented. As follows from these Fig. and Tables values of HCI in 6 locations of Kakheti changes from 43 (Lagodekhi, February, Marginal) to 96 (Telavi, October, Ideal). As follows from these Fig. and Tables values of HCI in 6 locations of Kakheti changes from 59 (Acceptable) to 89 (Excellent).



Fig. 6. Repetition of monthly values of TCI categories at six locations of Kakheti in 1956-2015.

In Fig. 6 data about repetition of monthly values of TCI category at six locations of Kakheti in 1956-2015 are presented. The variability of this repetition for separate locations of Kakheti is as follows: Dedoplistskaro (2,6%, Ideal – 34.2%, Good), Gurjaani (1,5%, Ideal – 33,8% Good), Kvareli (1,4%, Ideal – 36.3%, Good), Lagodekhi (0,6%, Marginal, Ideal – 35.4%, Good), Sagarejo (5,0%, Ideal – 34.2%, Good), Telavi (0,6%, Ideal – 36.7%, Good).

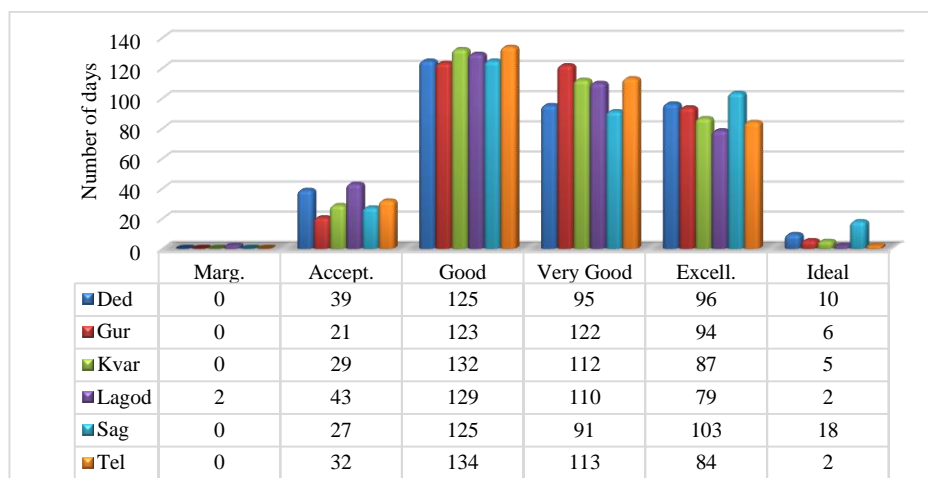


Fig. 7. Number of days in year of various categories of HCI at six locations of Kakheti in 1956-2015.

Respectively the variability of number of days in year of various categories of HCI for separate locations of Kakheti is as follows (Fig. 7): Dedoplistskaro (10, Ideal – 125, Good), Gurjaani (6, Ideal – 123, Good), Kvareli (5, Ideal – 132, Good), Lagodekhi 2, Marginal, Ideal – 129, Good), Sagarejo (18, Ideal – 125, Good), Telavi (2, Ideal – 134, Good).

In Table 14 and Fig. 8-11 data about Min, Max and mean values of HCI components at six locations of Kakheti are presented. As follows from these Table and Fig. the monthly values of HCI components changes from 0 (Rd, Lagodekhi) to 10 (all components, all locations). The monthly mean values of HCI components changes from 3.3 (T, Lagodekhi) to 10 (W, all locations besides Lagodekhi).

Table 14. Data of Min and Max values of HCI components at six locations of Kakheti.

Parameter	All data		Mean monthly 1956-2015		All data		Mean monthly 1956-2015	
	Min	Max	Min	Max	Min	Max	Min	Max
	Dedoplistskaro				Lagodekhi			
T	2	10	3.3	9.2	2	10	3.6	9.3
A	2	10	4.8	7.0	2	10	4.6	7.0
Rd	5	10	8.4	9.0	0	10	7.7	9.0
W	9	10	9.8	10	8	10	9.6	9.9
	Gurjaani				Sagarejo			
T	2	10	3.6	9.4	2	10	3.4	9.5
A	3	10	5.2	7.4	3	10	5.1	7.4
Rd	2	10	8.1	9.0	2	10	8.3	9.0
W	10	10	10	10	9	10	9.9	10
	Kvareli				Telavi			
T	2	10	3.6	9.6	2	10	3.4	9.6
A	3	10	4.8	7.2	1	10	4.5	6.5
Rd	2	10	7.9	9.0	5	10	8.0	9.0
W	10	10	10	10	9	10	9.9	10

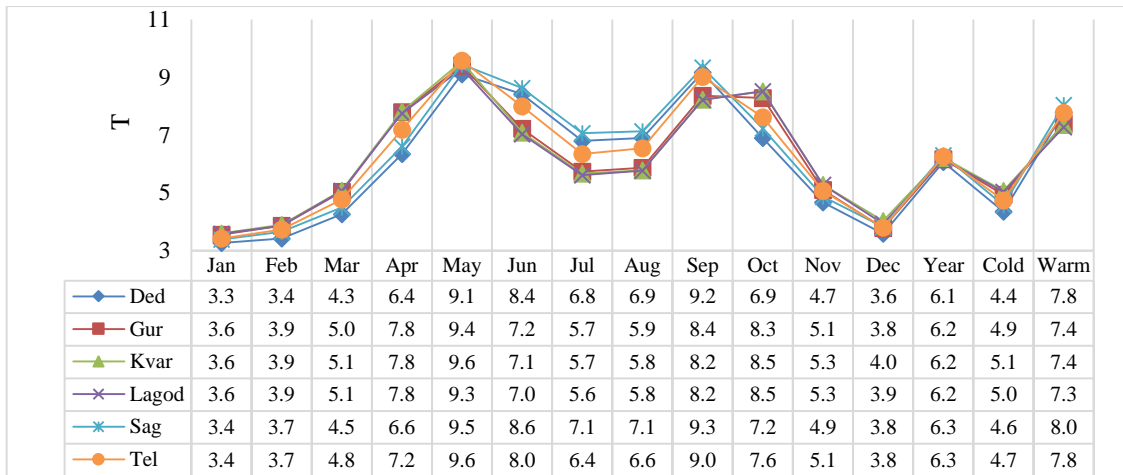


Fig. 8. Mean monthly and seasonal values of T component of HCI at six locations of Kakheti in 1956-2015.

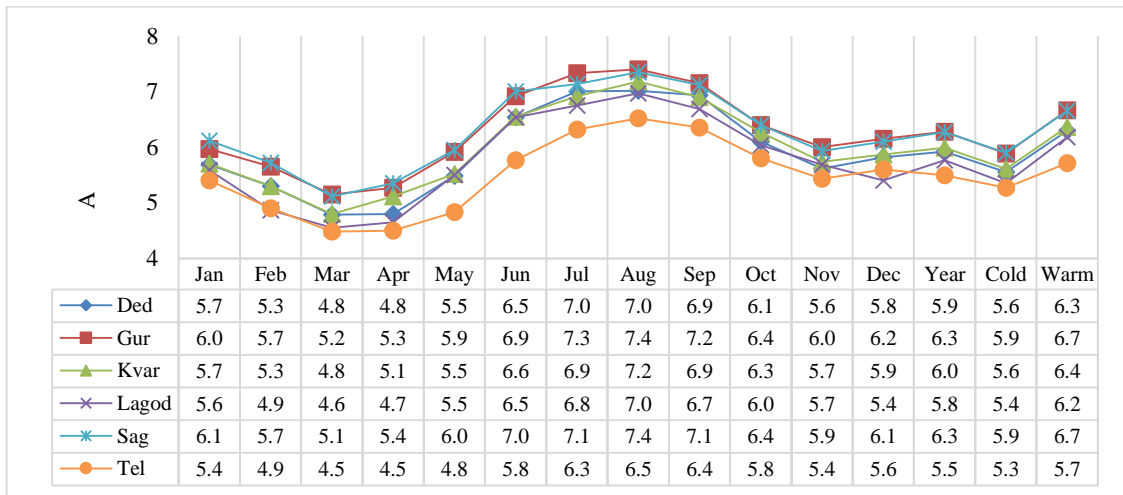


Fig. 9. Mean monthly and seasonal values of A component of HCI at six locations of Kakheti in 1956-2015.

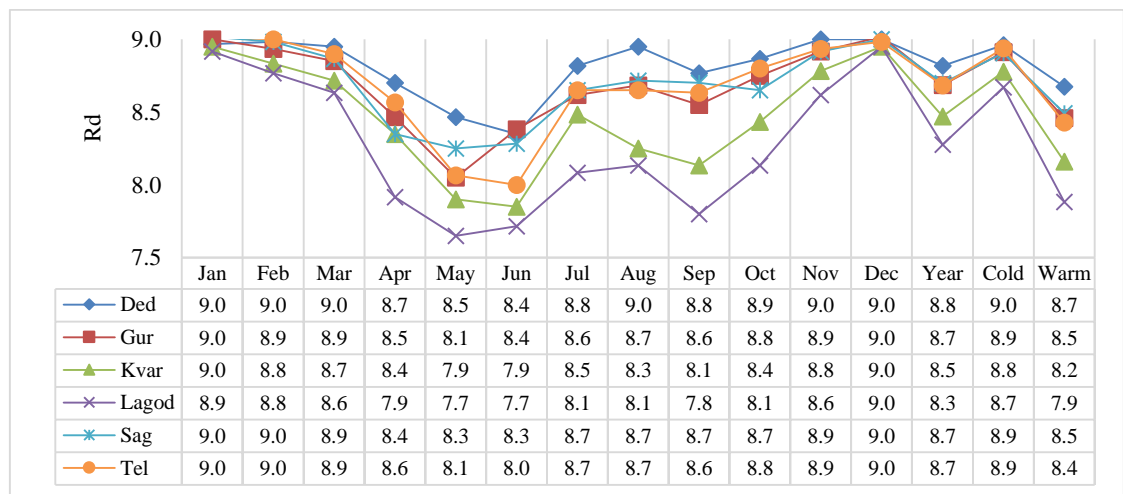


Fig. 10. Mean monthly and seasonal values of Rd component of HCI at six locations of Kakheti in 1956-2015.

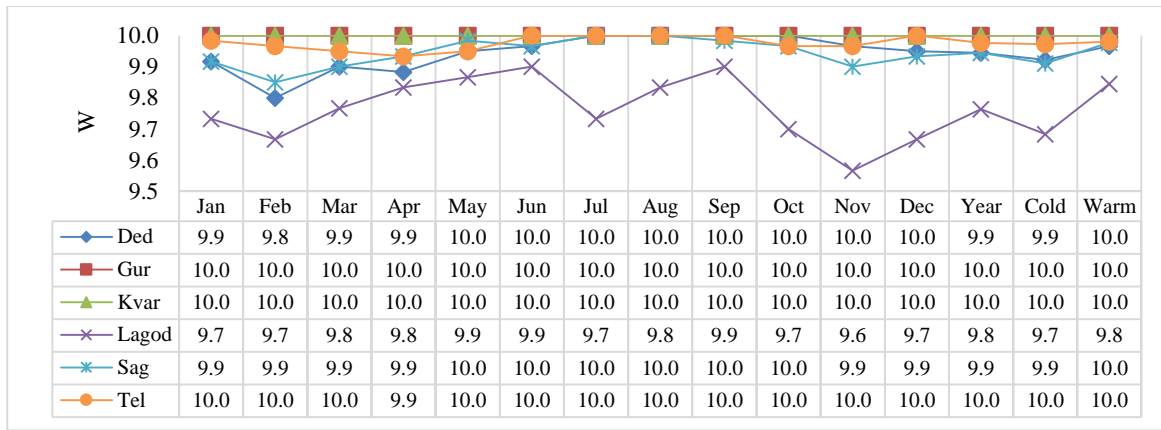


Fig. 11. Mean monthly and seasonal values of W component of HCI at six locations of Kakheti in 1956-2015.

Intra-annual distribution of mean monthly values of T component of HCI at six locations of Kakheti (Fig. 8), as well HCI (Fig. 2, Table 7), bimodal distribution type of this parameter is observed. The first extremum of T component for all locations in May is observed (range of change: 9.1, Dedoplistskaro – 9.6, Kvareli, Telavi). The second extremum of T component for Dedoplistskaro, Sagarejo and Telavi in September is observed, for Gurjaani, Kvareli and Lagodekhi – in September-October. Range of change for Dedoplistskaro, Sagarejo and Telavi is as follow: 9.0, Telavi – 9.3, Sagarejo. Range of change for Gurjaani, Kvareli and Lagodekhi is as follow: 8.2, Kvareli, Lagodekhi – 8.5, Kvareli, Lagodekhi.

Intra-annual distribution of mean monthly values of A component of HCI at six locations of Kakheti (Fig. 9) unimodal distribution type of this parameter is observed. The Max values of A component for all locations, besides Dedoplistskaro, in August is observed (range of change: 6.5, Telavi – 7.4, Gurjaani, Sagarejo). In Dedoplistskaro unimodal distribution type with plateau of A value is observed (7.0, July-August).

Intra-annual variation of mean monthly values of Rd component of HCI at six locations of Kakheti (Fig. 10) is not very significant. Range of change for separate locations is as follow. Dedoplistskaro: 8.4, June – 9.0, November-March, August; Gurjaani: 8.1, May – 9.0, December, January; Kvareli: 7.9, May, June – 9.0, December, January; Lagodekhi: 7.7, May, June – 9.0, December; Sagarejo: 8.3, May, June – 9.0, December – February; Telavi: 8.0, June – 9.0, December – February.

Intra-annual variation of mean monthly values of W component of HCI at six locations of Kakheti (Fig. 11) is not significant - range of change for all stations – from 9.6 to 10.0.

In Table 15 and Fig. 12-15 data about Min, Max and mean values of share of HCI components in HCI values at six locations of Kakheti are presented. As follows from these Table and Fig. the monthly values of share of HCI components changes from 0% (Share Rd, Lagodekhi) to 60.6% (Share T, Gurjaani, Lagodekhi). The monthly mean values of share of HCI components changes from 11.4% (Share W, Sagarejo) to 46.8% (Share T, Telavi).

Intra-annual distribution of mean monthly values of Share of T component of HCI value at six locations of Kakheti (Fig. 12), as well HCI and T (Fig. 2, 8, Table 7), bimodal distribution type of this parameter is observed. Range of change of Share of T component of HCI value for all stations (first extremum) is as follow: 44.0%, Dedoplistskaro – 46.8%, Telavi. Range of change of Share of T component of HCI value (second extremum) for Dedoplistskaro, Sagarejo and Telavi is as follow: 42.2%, Dedoplistskaro – 42.6%, Sagarejo. Range of change for Gurjaani, Kvareli and Lagodekhi is as follow: 40.1, Gurjaani – 42.5%, Lagodekhi.

Table 15. Data of Min and Max values of share of HCI components in HCI value at six locations of Kakheti, (%)

Parameter	All data		Mean monthly 1956-2015		All data		Mean monthly 1956-2015	
	Min	Max	Min	Max	Min	Max	Min	Max
	Dedoplistskaro				Lagodekhi			
Share T	12.3	53.3	21.2	44.0	13.1	60.6	23.1	46.1
Share A	7.5	30.8	13.2	18.5	7.4	29.2	12.5	19.7
Share Rd	19.0	50.9	30.3	44.3	0	50.9	28.1	43.3
Share W	10.0	19.2	11.5	16.2	9.0	23.3	12.2	15.9
	Gurjaani				Sagarejo			
Share T	13.1	60.6	22.4	45.1	15.1	60.0	21.5	44.9
Share A	8.5	27.4	13.6	20.0	9.4	26.9	14.1	19.4
Share Rd	9.1	47.4	28.7	42.9	10.0	50.9	29.2	43.3
Share W	10.4	18.5	12.0	15.9	10.0	19.6	11.4	15.9
	Kvareli				Telavi			
Share T	12.1	60.0	22.9	46.1	14.5	54.8	22.1	46.8
Share A	7.5	27.3	13.3	19.9	3.6	27.4	11.7	17.6
Share Rd	10.0	49.1	28.5	43.0	20.5	50.9	29.3	44.1
Share W	10.5	18.5	12.1	16.0	10.2	18.9	11.8	16.3

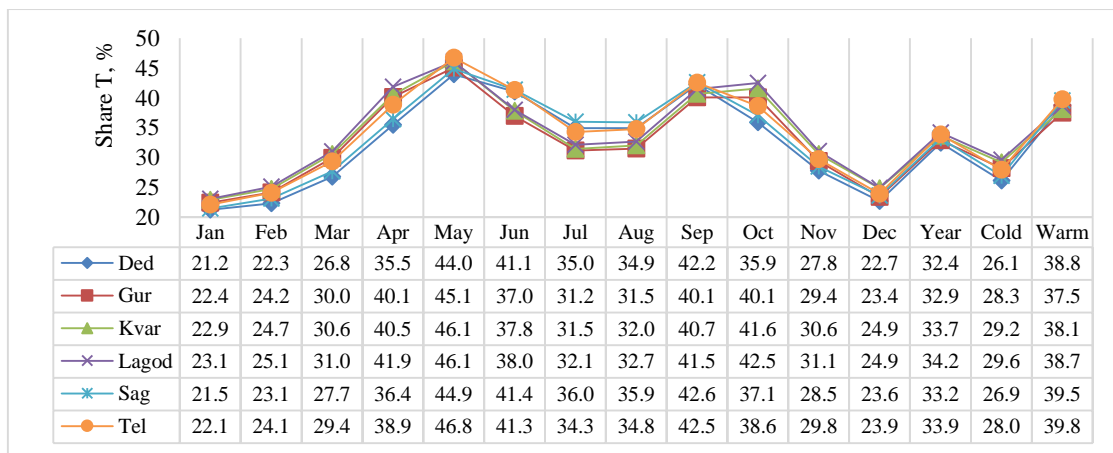


Fig. 12. Share of mean monthly and seasonal values of T component in HCI value at six locations of Kakheti in 1956-2015.

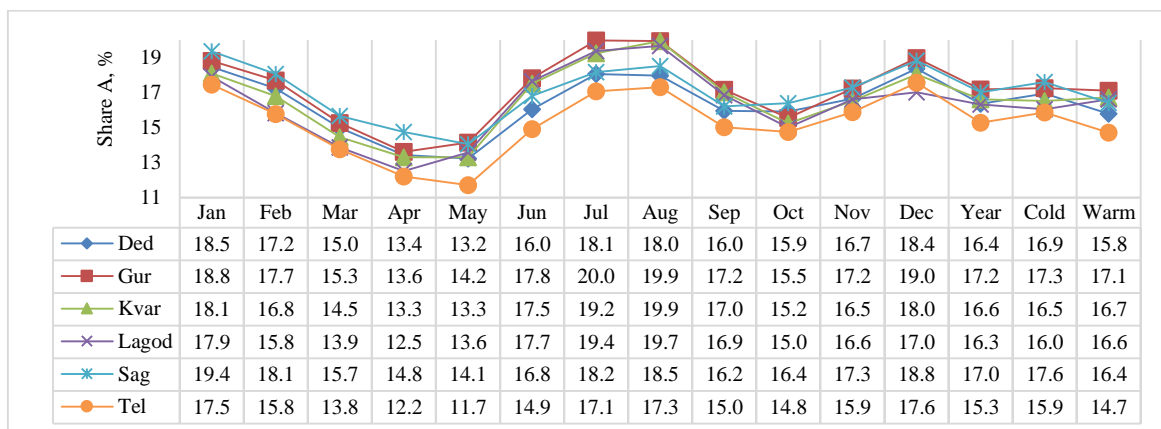


Fig. 13. Share mean monthly and seasonal values of A component in HCI value in six locations of Kakheti in 1956-2015.

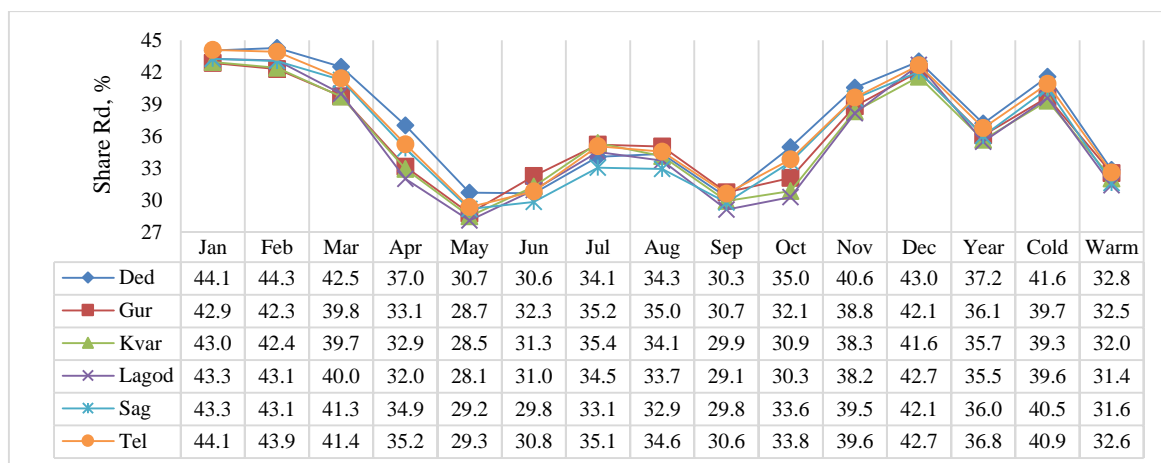


Fig. 14. Share mean monthly and seasonal values of Rd component in HCI value at six locations of Kakheti in 1956-2015.

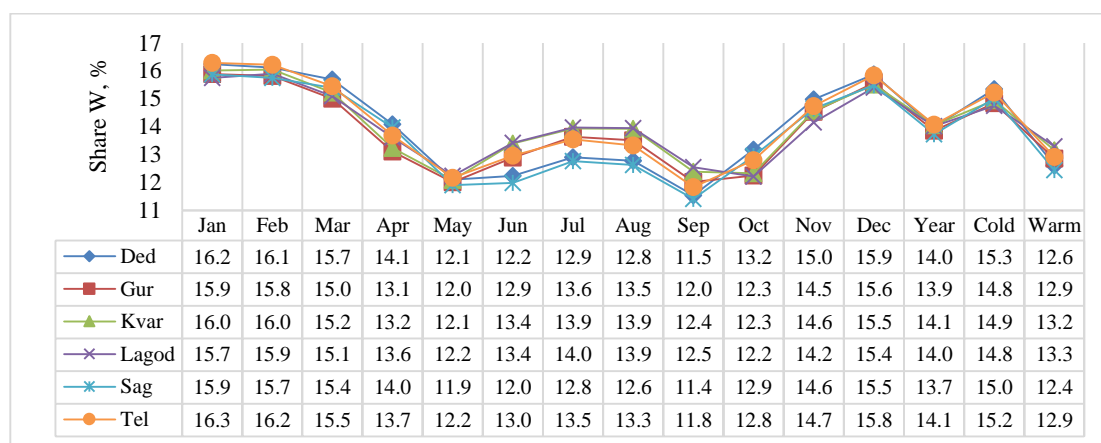


Fig. 15. Share of mean monthly and seasonal values of W component in HCI value at six locations of Kakheti in 1956-2015.

Intra-annual distribution of mean monthly values of Share of A components of HCI value at six locations of Kakheti (Fig. 13), as well HCI, T and Share T (Fig. 2, 8, 12, Table 7), have a bimodal type. For Dedoplistskaro, Gurjaani, Kvareli and Telavi first extremum of Share of A distribution in January and December is observed, for Lagodekhi and Sagarejo – in January. The second extremum of this distribution for all stations in July-August is observed.

Range of change of Share of A component of HCI value for all stations (first extremum) is as follow: 17.5%, Telavi – 19.0%, Gurjaani; Lagodekhi and Sagarejo – 17.9% and 19.4% respectively. Range of change of Share of A component of HCI value for all stations (second extremum) is as follow: 17.1%, Telavi – 20.0%, Gurjaani.

Intra-annual distribution of mean monthly values of Share of Rd and Share of W components of HCI value at all locations of Kakheti (Fig. 14, 15) have a bimodal type. Whereinto for both distributions, the first and second extremums coincide in time: December-February and July-August respectively. It should be noted that in both cases the summer extremum is significantly lower than the winter one.

Range of change of Share of Rd component of HCI value for all stations (first extremum) is as follow: 41.6%, Kvareli – 44.3%, Dedoplistskaro. Range of change of Share of Rd component of HCI value for all stations (second extremum) is as follow: 32.9%, Sagarejo – 35.4%, Kvareli.

Range of change of Share of W component of HCI value for all stations (first extremum) is as follow: 15.4%, Lagodekhi – 16.3%, Telavi. Range of change of Share of W component of HCI value for all stations (second extremum) is as follow: 12.6%, Sagarejo – 14.0%, Lagodekhi.

Range of change of different components of HCI for all station in year, cold and warm seasons is as follow respectively.

T component (Fig. 8). Year: 6.1, Dedoplistskaro – 6.3, Sagarejo, Telavi; cold season: 4.4, Dedoplistskaro – 5.1, Kvareli; warm season: 7.3, Lagodekhi – 8.0, Sagarejo.

A component (Fig. 9). Year: 5.5, Telavi – 6.3, Gurjaani, Sagarejo; cold season: 5.3, Telavi – 5.9, Gurjaani, Sagarejo; warm season: 5.7, Telavi – 6.7, Gurjaani, Sagarejo.

Rd component (Fig. 10). Year: 8.3, Lagodekhi – 8.8, Dedoplistskaro; cold season: 8.7, Lagodekhi – 9.0, Dedoplistskaro; warm season: 7.9, Lagodekhi – 8.7, Dedoplistskaro.

W component (Fig. 11). Year: 9.8, Lagodekhi – 10.0, Gurjaani, Kvareli, Telavi; cold season: 9.7, Lagodekhi – 10.0, Gurjaani, Kvareli, Telavi; warm season: 9.8, Lagodekhi – 10.0, all other locations.

Share of T component (Fig. 12, %). Year: 32.4, Dedoplistskaro – 34.2, Lagodekhi; cold season: 26.1, Dedoplistskaro – 29.6, Lagodekhi; warm season: 37.5, Gurjaani – 39.8, Telavi.

Share of A component (Fig. 13, %). Year: 15.3, Telavi – 17.2, Gurjaani; cold season: 15.9, Telavi – 17.6, Sagarejo; warm season: 14.7, Telavi – 17.1, Gurjaani.

Share of Rd component (Fig. 14, %). Year: 35.5, Lagodekhi – 37.2, Dedoplistskaro; cold season: 39.3, Kvareli – 41.6, Dedoplistskaro; warm season: 31.4, Lagodekhi – 32.8, Dedoplistskaro.

Share of W component (Fig. 15, %). Year: 13.7, Sagarejo – 14.1, Kvareli, Telavi; cold season: 14.8, Gurjaani, Lagodekhi – 15.3, Dedoplistskaro; warm season: 12.4, Sagarejo – 13.3, Lagodekhi.

3. Comparison of TCI and HCI at four location of Kakheti in 1961-2010.

Comparison of TCI and HCI values in many investigations are provided [17, 18, 20, 22, 23]. In Georgia, an analogous study conducted for Tbilisi [24].

Comparison of TCI and HCI at four location of Kakheti (Dedoplistskaro, Kvareli, Sagarejo and Telavi) in 1961-2010 is provided below (Fig. 16, 17, Table 16-18).

In Fig. 16 data about mean monthly and seasonal values of TCI and HCI at four locations of Kakheti in 1961-2010 are presented.

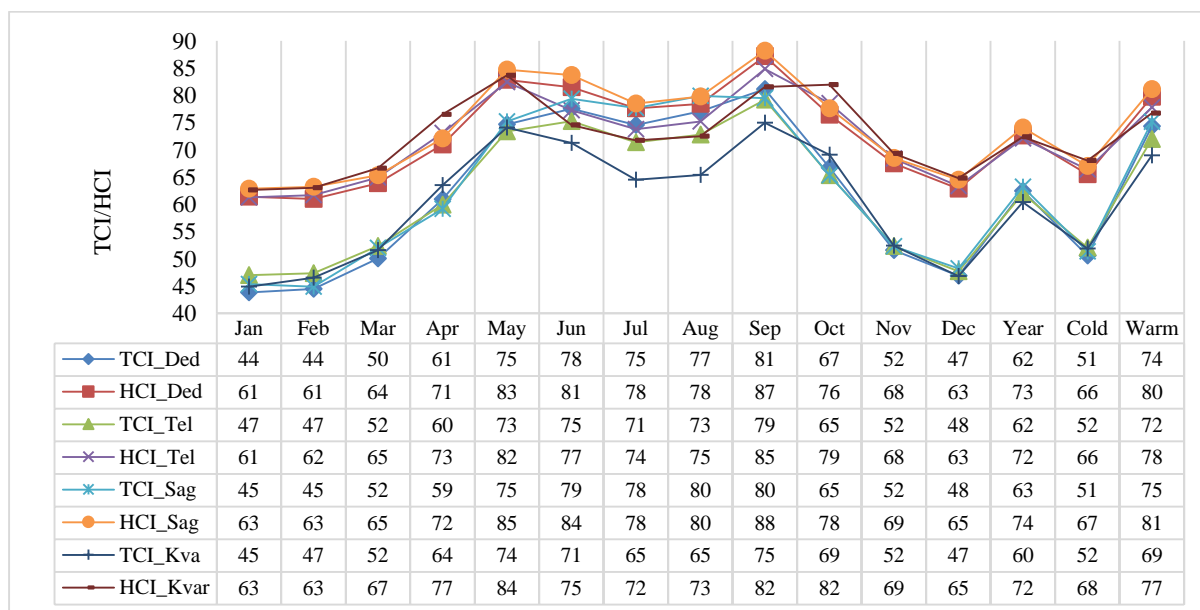


Fig. 16. Comparison of mean monthly and seasonal values of TCI and HCI at four location of Kakheti in 1961-2010.

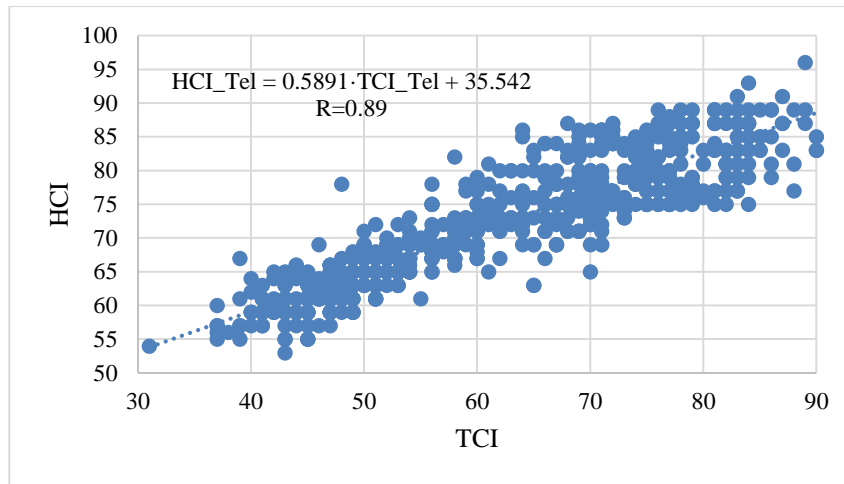


Fig. 17. Linear correlation and regression between monthly values of TCI and HCI in Telavi in 1961-2010.

$$\begin{aligned} \text{HCI_Ded} &= 0.5886 \cdot \text{TCI_Ded} + 35.869, (R=0.93); \text{HCI_Kvar} = 0.584 \cdot \text{TCI_Kvar} + 37.127, (R=0.88); \\ \text{HCI_Sag} &= 0.587 \cdot \text{TCI_Sag} + 36.969, (R=0.92); \text{For all values of } R, \alpha < 0.005. \end{aligned}$$

Comparison of the values and categories of the Tourism Climate Index [7] and Holiday Climate Index (Fig. 16, 17, Table 16-18) shows that the intra-annual variation of both indices is similar and has a bimodal form. Linear correlation between monthly values of TCI and HCI is high (Fig. 17, example for Telavi) and change from 0.88 (Kvareli) to 0.93 (Dedoplistskaro).

Table 16. Comparison of TCI and HCI categories at four location of Kakheti in cold period of year in 1961-2010.

Month	Jan	Feb	Mar	Oct	Nov	Dec	Cold	Year
TCI_Ded	Marg. Good	Marg. Good	Accept. Good	Good Very Good	Accept. Good	Marg. Good	Accept. Good	Good Very Good
HCI_Ded								
TCI_Tel								
HCI_Tel								
TCI_Sag				Good				
HCI_Sag								
TCI_Kvar								
HCI_Kvar				Excell.				

Table 17. Comparison of TCI and HCI categories at four location of Kakheti in warm period of year in 1961-2010.

Month	Apr	May	Jun	Jul	Aug	Sep	Warm		
TCI_Ded	Good	Very Good Excell.	Very Good	Very Good	Very Good	Excell.	Very Good		
HCI_Ded	Very Good		Excell.						
TCI_Tel	Accept. Very Good		Very Good						
HCI_Tel			Very Good						
TCI_Sag			Very Good						
HCI_Sag	Good		Excell.			Good	Good	Very Good Excell.	Good
TCI_Kvar			Very Good						
HCI_Kvar			Very Good						

Comparison of TCI and HCI categories shows, that in cold months, season and year HCI categories on 1-2 step higher than TCI categories (Table 16). Difference on 2 step in the following cases are observed: TCI_Marginal → HCI_Good, in January, February, December for all stations; TCI_Good →

HCI_Excellent, in October, for Kvareli. Difference on 1 step: TCI_Acceptable → HCI_Good, in March, November, Cold season, for all stations; TCI_Good → HCI_Very Good in October for Dedoplistskaro, Telavi, Sagarejo and in year, for all stations.

Comparison of TCI and HCI categories in warm months and season shows, that HCI categories either the same than TCI categories, or 1-2 steps higher (Table 17).

Difference on 2 step in the following case is observed: TCI_Acceptable → HCI_Very Good, in April for Telavi and Sagarejo.

Difference on 1 step: TCI_Good → HCI_Very Good, in April, for Dedoplistskaro and Kvareli; in July, August and Warm season, for Kvareli; TCI_Very Good → HCI_Excellent in May for all stations, in June for Dedoplistskaro and Sagarejo; in September, for Telavi, Sagarejo and Kvareli; in Warm season, for Sagarejo.

The same categories: Very Good, in June, for Telavi and Kvareli; in July and August, for Dedoplistskaro, Telavi and Sagarejo; in Warm season, for Dedoplistskaro and Telavi. Excellent, in September, for Dedoplistskaro.

Table 18. Comparison of Min and Max values of TCI and HCI and its category for four location of Kakheti in 1961-2010.

Parameter		Value	Location/Month/Category
TCI	Min	44	Dedoplistskaro, January, February, Marginal
	Max	81	Dedoplistskaro, September, Excellent
	Mean on Station	62	Good
HCI	Min	61	Good, Dedoplistskaro, January, February; Telavi, January
	Max	88	Sagarejo, September, Excellent
	Mean on Station	73	Very Good

In Table 18 comparison of Min and Max values of TCI and HCI and its category for four location of Kakheti is provided. As follows from this Table TCI value change from 44 (Dedoplistskaro, January, February, Marginal) to 81 (Dedoplistskaro, September, Excellent) and HCI value – from 61 (Good, Dedoplistskaro, January, February; Telavi, January) to 88 (Sagarejo, September, Excellent). Mean on the one station for TCI is 62 (Good) and for HCI - 73 (Very Good).

However, given that the TCI is calculated for the so-called “average tourist” (regardless of gender, age, physical condition), the value and category of this index is lower than the HCI values and categories. In general, HCI more adequately determines the bioclimatic state of the environment for the development of certain types of tourism (winter tourism, extreme tourism, etc.) than TCI.

Conclusion

In the future, we are planning a more detailed study of the climatic resources of this and other regions of Georgia for tourism, recreation and treatment (mapping the territory on HCI and TCI, long-term forecasting of HCI and TCI, determining other modern climatic and bioclimatic indicators for tourism, recreation and treatment, assessing the adequacy of bioclimatic indicators scales to human health, etc.).

References

- [1] Abegg B. Klimaänderung und Tourismus. Zurich: Schlussbericht NFP 31. vdfHochschulverlag AG ander ETH, 1996.
- [2] Matzarakis A. Weather - and Climate-Related Information for Tourism. Tourism and Hospitality Planning & Development, August, 2006, vol. 3, No. 2, pp. 99–115.
- [3] Matzarakis A., Cheval S., Lin T.-P., Potchter, O. Challenges in Applied Human Biometeorology. Atmosphere 2021, 12, 296. <https://doi.org/10.3390/atmos12030296>
- [4] Amiranashvili A.G., Chikhladze V.A. Saakashvili N.M., Tabidze M.Sh., Tarkhan-Mouravi I.D. Bioclimatic Characteristics of Recreational Zones – Important Component of the Passport of the Health

- Resort – Tourist Potential of Georgia. Trans. of the Institute of Hydrometeorology at the Georgian Technical University, vol. 117, ISSN 1512-0902, 2011c, pp. 89-92.
- [5] Kartvelishvili L., Matzarakis A., Amiranashvili A., Kutaladze N. Assessment of Touristical-Recreation Potential of Georgia on Background Regional Climate Change. Proc. of IIST Int. Scientific-Practical Conference “Tourism: Economics and Business”, June 4-5, Batumi, Georgia, 2011, pp. 250-252.
- [6] Rutty M., Steiger R., O. Demiroglu O.C., Perkins D.R.. Tourism Climatology: Past, Present, and Future. Int. Journ. of. Biometeorology. Published online: 08 January 2021. <https://doi.org/10.1007/s00484-020-02070-0>
- [7] Mieczkowski Z. The Tourism Climate Index: A Method for Evaluating World Climates for Tourism. The Canadian Geographer 1985, N 29, pp. 220-233.
- [8] Amiranashvili A., Matzarakis A., Kartvelishvili L. Tourism Climate Index in Tbilisi. Trans. of the Institute of Hydrometeorology, ISSN 1512-0902, Tbilisi, 18 – 19 November, 2008, vol. 115, pp. 27 - 30.
- [9] Amiranashvili, A., Chargazia, Kh., Matzarakis, A. Comparative Characteristics of the Tourism Climate Index in the South Caucasus Countries Capitals (Baku, Tbilisi, Yerevan). Journal of the Georgian Geophysical Society, ISSN: 1512-1127, Issue (B). Physics of Atmosphere, Ocean, and Space Plasma, 2014, vol.17B, pp. 14-25.
- [10] Amiranashvili A., Chargazia Kh., Matzarakis A., Kartvelishvili L. Tourism Climate Index in the Coastal and Mountain Locality of Adjara, Georgia. International Scientific Conference “Sustainable Mountain Regions: Make Them Work”. Proceedings, Borovets, Bulgaria, ISBN 978-954-411-220-2, 14-16 May, 2015, pp.238-244, http://geography.bg/MountainRegions_Sofia2015
- [11] Rybak O. O., Rybak E. A. Application of Climatic Indices for Evaluation of Regional Differences in Tourist Attractiveness. Nauchnyy zhurnal KubGAU, №121(07), 2016, 24 p., <http://ej.kubagro.ru/2016/07/pdf/16.pdf>
- [12] Amiranashvili A.G., Japaridze N.D., Kartvelishvili L.G., Khazaradze K.R., Matzarakis A., Povolotskaya N.P., Senik I.A. Tourism Climate Index of in the Some Regions of Georgia and North Caucasus. Journal of the Georgian Geophysical Society, ISSN: 1512-1127, Issue (B). Physics of Atmosphere, Ocean, and Space Plasma, 2017, vol.20B, pp. 43-64.
- [13] Amiranashvili A.G., Japaridze N.D., Kartvelishvili L.G., Khazaradze K.R., Kurdashvili L.R. Tourism Climate Index in Kutaisi (Georgia). International Scientific Conference „Modern Problems of Ecology“, Proceedings, ISSN 1512-1976, v. 6, Kutaisi, Georgia, 21-22 September, 2018, pp. 227-230.
- [14] Amiranashvili A.G., Kartvelishvili L.G., Matzarakis A., Megrelidze L.D. The Statistical Characteristics of Tourism Climate Index in Kakheti (Georgia). Journal of the Georgian Geophysical Society, ISSN: 1512-1127, Physics of Solid Earth, Atmosphere, Ocean and Space Plasma, v. 21(2), Tbilisi, 2018, pp. 95-112.
- [15] Amiranashvili A., Kartvelishvili L. Statistical Characteristics of the Monthly Mean Values of Tourism Climate Index in Mestia (Georgia) in 1961-2010. Journal of the Georgian Geophysical Society, ISSN: 1512-1127, Physics of Solid Earth, Atmosphere, Ocean and Space Plasma, v. 22(2), 2019, pp. 68–79.
- [16] Kartvelishvili L., Amiranashvili A., Megrelidze L., Kurdashvili L. Turistul Rekreativiuli Resursebis Shefaseba Klimatis Cvlilebebis Fonze. Publish Hous "Mtsignobari", ISBN 978-9941-485-01-5, Tbilisi, 2019, 161 p., (in Georgian). <http://217.147.235.82/bitstream/1234/293074/1/turistulRekreaciuliResursebisShefasebaKlimatisCvlilebebisFonze.pdf>.
- [17] Scott D., Rutty M., Amelung B., Tang M. An Inter-Comparison of the Holiday Climate Index (HCI) and the Tourism Climate Index (TCI) in Europe. Atmosphere 7, 80, 2016, 17 p., doi:10.3390/atmos7060080www.
- [18] Javan K. Comparison of Holiday Climate Index (HCI) and Tourism Climate Index (TCI) in Urmia. Physical Geography Research Quarteli. vol. 49, iss. 3, 2017, pp. 423-439.
- [19] Öztürk A., Göral R. Climatic Suitability in Destination Marketing and Holiday Climate Index. Global Journal of Emerging Trends in e-Business, Marketing and Consumer Psychology (GJETeMCP). An Online International Research Journal (ISSN: 2311-3170), Vol: 4 Issue: 1, 2018, pp. 619-629.
- [20] Yu D. D., Rutty M., Scott D., Li S. A Comparison of the Holiday Climate Index: Beach and the Tourism Climate Index Across Coastal Destinations in China International Journal of Biometeorology, 2020, <https://doi.org/10.1007/s00484-020-01979-w>, 8 p.
- [21] Demiroglu O., Saygili-Araci F., Pacal A., Hall C., Kurnaz M. Future Holiday Climate Index (HCI) Performance of Urban and Beach Destinations in the Mediterranean. Atmosphere, 11, 911, 2020, doi:10.3390/atmos11090911, 30 p.

- [22] Rutty M., Scott D., Matthews L., Burrowes R., Trotman A., Mahon R., Charles A. An Inter-Comparison of the Holiday Climate Index (HCI: Beach) and the Tourism Climate Index (TCI) to Explain Canadian Tourism Arrivals to the Caribbean. *Atmosphere* 2020, 11, 412.
- [23] Hejazizadeh Z., Karbalaee A., Hosseini S.A., Tabatabaei S.A. Comparison of the Holiday Climate Index (HCI) and the Tourism Climate Index (TCI) in Desert Regions and Makran Coasts of Iran. *Arab. J. Geosci.* 12, 803, 2019, <https://doi.org/10.1007/s12517-019-4997-5>
- [24] Amiranashvili A., Kartvelishvili L., Matzarakis A. Comparison of the Holiday Climate Index (HCI) and the Tourism Climate Index (TCI) in Tbilisi. *Int. Sc. Conf. „Modern Problems of Ecology“*, Proc., ISSN 1512-1976, v. 7, Tbilisi-Telavi, Georgia, 26-28 September, 2020, pp. 424-427.
- [25] Amiranashvili A., Kartvelishvili L., Matzarakis A. Changeability of the Holiday Climate Index (HCI) in Tbilisi. *Transactions of Mikheil Nodia Institute of Geophysics*, ISSN 1512-1135, vol. LXXII, 2020, pp. 131-139.

დასვენების კლიმატური ინდექსი კახეთში

ა.ამირანაშვილი, ლ.ქართველიშვილი

რეზიუმე

წარმოდგენილია მონაცემები დასვენების კლიმატური ინდექსის (დკი) მრავალწლიური საშუალო თვიური მნიშვნელობების შესახებ კახეთის 12 ადგილისთვის (ახმეტა, დედოფლისწყარო, გომბორი, გურჯაანი, ყვარელი, ლაგოდეხი, ომლო, საგარეჯო, შირაქი, თელავი, წნორი და უდაბნო). ამ რეგიონის 6 სადგურისთვის (დედოფლისწყარო, გურჯაანი, ყვარელი, ლაგოდეხი, საგარეჯო და თელავი) ჩატარებულია ყოველთვიური, სეზონური და წლიური დკი-ს მნიშვნელობების დეტალური ანალიზი 60 წლიანი პერიოდის განმავლობაში (1956-2015 წწ.). ჩატარებულია დკი-ს და ტურიზმის კლიმატური ინდექსის ყოველთვიური მნიშვნელობების შედარება კახეთის ოთხ პუნქტზე (დედოფლისწყარო, ყვარელი, საგარეჯო და თელავი) 1961-2010 წწ. მონაცემების საფუძველზე.

Климатический Индекс Отдыха в Кахетии (Грузия)

А.Г. Амиранашвили, Л. Г. Картвелишвили

Резюме

Представлены данные о многолетних среднемесячных значениях Климата Индекса Отдыха (КИО) для 12 населенных пунктов Кахетии (Ахмета, Дедоплисцкаро, Гомбори, Гурджаани, Кварели, Лагодехи, Омало, Сагареджо, Шираки, Телави, Цнори и Удабно). Для 6 станций этого региона (Дедоплисцкаро, Гурджаани, Кварели, Лагодехи, Сагареджо и Телави) проведен детальный анализ месячных, сезонных и годовых значений КИО за 60-летний период времени (1956-2015 гг.). Проведено сравнение месячных значений КИО и Климатического Индекса Туризма для четырех пунктов Кахетии (Дедоплисцкаро, Кварели, Сагареджо и Телави) по данным с 1961 по 2010 гг.

Connection of Holiday Climate Index with Public Health (on Example of Tbilisi and Kakheti Region, Georgia)

¹Avtandil G. Amiranashvili, ²Aza A. Revishvili, ^{2,3}Ketevan R. Khazaradze, ^{3,4}Nino D. Japaridze

¹Mikheil Nodia Institute of Geophysics of Ivane Javakhishvili Tbilisi State University, Tbilisi, Georgia
avtandilamiranashvili@gmail.com

²Georgian State Teaching University of Physical Education and Sport, Tbilisi, Georgia

³Ministry of Internally Displaced Persons from Occupied Territories, Labour, Health and Social Affairs of Georgia, Tbilisi, Georgia

⁴Tbilisi State Medical University, Tbilisi, Georgia

ABSTRACT

Study on connection of Holiday Climate Index (HCI) and its components and components ratings with public health (on Example of Tbilisi - mortality by cardiovascular diseases and Kakheti Region of Georgia - ambulance calls, hospitalization, total mortality are presented).

It is shown that, in general, all indicators adequately correspond to the degree of bioclimatic comfort of the living environment for people. In particular, the relationship between HCI and mortality in Tbilisi has the form of a second-degree polynomial, and in Kakheti - a third-degree polynomial. At the same time, with an increase in the degree of bioclimatic comfort to the category "Very good", there is a tendency for a decrease in mortality. With the transition to the "Excellent" category, there is a slight increase in mortality. A similar result for some components of the HCI and the ratings of these components was obtained.

It is proposed to make adjustments in determining the degree of comfort of bioclimatic index scales for the population, taking into account local social and climatic conditions.

Key words: *Holiday Climate Index (HCI), public health, ambulance calls, hospitalization, mortality.*

Introduction

Studies of weather conditions, climate change, atmospheric air quality, and also of different heliogeophysical and space factors for the human organism, are conducted in many countries of the world [1-8].

A significant number of works are devoted to the study of the impact on human health of individual meteorological and heliogeophysical elements, space weather parameters, etc. In particular, these elements include: air temperature [5, 8-11], wind speed [11], humidity [12], atmospheric pressure [13], solar activity (Wolf's number) [12-15], the geomagnetic fields [16-18], solar radiation [9], the cosmic rays [9, 19], light ions [9, 20, 21], aerosols [9, 22], ozone [9, 23], other air toxic admixtures and etc. [9, 24].

It is well known that the effects of a significant increase in the mortality of population with the strong cold and the extreme heat [25-31]. For determining the extent of comfort or discomfort of the human living environment for her health (so-called "average person") frequently are used different simple and complex thermal indices [32-36].

Simple thermal indices involve more than one meteorological parameter and consider the combined effects on human organism (air equivalent- effective temperature - EET, Wet-bulb-globe temperature - derived from energy budget models. Such indices are popular in recent years, for example: Physiologically Equivalent Temperature (PET), Standard Effective Temperature (SET), Physiological Subjective Temperature and Subjective Temperature (MENEX), the Universal Thermal Climate Index (UTCI) etc. [37-41]).

Action on the human organism by the higher indicated factors have different scales - from minute, hour, day, decade and month to the seasonal and annual [5,7,9,14,17,19,23,30,31,33,36,42-44]. For example, in the works [43,44] is obtained that the dependence of mortality on EET takes the classical form - the decrease of mortality from the gradation "Sharply Coldly" to "Comfortably" with further increase to the gradation "Warmly". It is found in the works [30,31] that the relationship between the average monthly air temperature in Kutaisi and Kakheti region and such indices of the health of population as the total number of emergency medical calls, cases of hospitalizations and deaths has the form of a third power polynomial. In general, in the warm months there is a decrease of the total number of emergency medical calls, cases of hospitalizations and deaths. In the hot months, there is a worsening in these indicators of health, comparable to the cold months of the year (increase of the emergency medical calls, cases of hospitalizations and deaths).

For the bioclimatic zoning of territories (including for evaluating the bioclimatic potential of health resort- tourist industry) frequently is used the mean monthly values of simple thermal indices [33,42,45,46].

The standard scale and categories of the majority of these indices is usually used in this case for describing the real (hour or day) bioclimatic situation. In the latter case, as a rule, with the monthly averaging of meteorological data occurs range reduction of the scale of thermal indices and decrease of its sensitivity for evaluating the degree of the bioclimatic comfort of environment for the people. Therefore, the numerical values of the standard scale of thermal indices always cannot coincide with the verbal description of the categories of these indices. The results of investigating the connection of eight simple thermal indices and Tourism Climate Index with the monthly mortality of the population of Tbilisi city apropos of the cardiovascular diseases, which made it possible to estimate the representativeness of the standard scales and categories of the indicated indices as the bioclimatic indicator in monthly time scale, were represented in [36].

In this work, comparative analysis of the connection of eight simple thermal indices and Tourism Climate Index (TCI) [47] with the monthly mortality of the population of Tbilisi city apropos of cardiovascular diseases is represented. The values of simple thermal indices were calculated with the use of mean monthly and mean monthly for 13 hours data of meteorological elements. Between all studied simple thermal indices practically direct functional connection with the coefficient of linear correlation not lower than 0.86 is observed. The connection of simple thermal indices with the TCI is nonlinear and takes the form of third power polynomial. The possibility of using the standard scales and categories of the indicated indices as the bioclimatic indicator in monthly time scale is studied. As a whole, all indices adequately correspond to the degree of the bioclimatic comfort of environment for the people - with an increase in the level of comfort the mortality diminishes. Most representative for this purpose is Missenard air effective temperature in 13 hours.

This work is a continuation of the study [36]. Results of investigation on connection of Holiday Climate Index (HCI) and its components and components ratings [48-51] with public health (on Example of Tbilisi - mortality by cardiovascular deseases and Kakheti Region of Georgia - ambulance calls, hospitalization, total mortality) are presented below.

Study Area, Material and Methods

Study area – Tbilisi and Kakheti region of Georgia (below - Kakheti).

In this work the Holiday Climate Index (HCI) is used. The HCI uses five climatic variables related to the three facets essential to tourism (table 2): thermal comfort (TC), aesthetic (A), and physical (P) facet. The five climatic variables used for the HCI input are maximum air temperature and relative humidity (TC), cloud cover (A), precipitation and wind (P) [48].

The HCI score is calculated according to the following formula [48]:

$$HCI = 4 \cdot T + 2 \cdot A + 3 \cdot R_d + 1 \cdot W.$$

This formula uses the following bioclimatic indicators and ratings of these indicators. Thermal Comfort (TC) - effective temperature (°C) [47] – combination of maximum dry-bulb temperature and mean values of relative humidity (%), rating of TC is T; CC - daily cloud cover (%), rating of CC is A; DP - daily precipitation (mm), rating of DP is R_d ; WS - wind speed (m/sec or km/hour), rating of WS is W. Values of all HCI components rating change from 0 to 10 [48-51]. Note that below the physical dimensions of the studied parameters are omitted.

In this work the data of HCI for Tbilisi is used [49]. For Kakheti, these data were averaged over six locations of this region (Dedoplistskaro, Gurjaani, Kvareli, Lagodekhi, Sagarejo and Telavi) [51].

The population health indicators are as follows. Mean monthly decade mortality by cardiovascular diseases – M; mean monthly decade values of ambulance calls - AC, hospitalization – H; total mortality – TM.

In the work analysis of data is carried out with the use of the standard statistical analysis methods. The following designations will be used below: Mean – average values; Min – minimal values; Max - maximal values; Range: Max – Min; St Dev – standard deviation; Cv, % - coefficient of variation: $Cv = 100 \cdot St\ Dev / Mean$; R^2 - coefficient of determination; R - coefficient of linear correlation; α - level of significance.

$R(M)$, ... etc. - coefficient of linear correlation between mortality, ... etc. with mean monthly values of HCI componets and these ration respectively; $\alpha (R(M))$, ... etc. - the corresponding values of the significance level of the linear correlation coefficient.

Results

Results in Table 1-6 and Fig. 1-13 are presented.

Table 1. Statistical characteristics of mean monthly values of HCI and mean monthly decade mortality by cardiovascular diseases in Tbilisi in 1980-1990.

Variable	HCI	HCI Category	Mortality
Mean	73	Very Good	105
Min	53	Acceptable	70
Max	89	Excellent	168
Range	36		98
St Dev	8.7		17.3
Cv, %	11.9		24.7

In Table 1 statistical characteristics of mean monthly values of HCI and mean monthly decade mortality by cardiovascular diseases in Tbilisi in 1980-1990 is presented. As follows from this Table values of HCI change from 53 (category - Acceptable) to 89 (category - Excellent). Values of M change from 70 to 168.

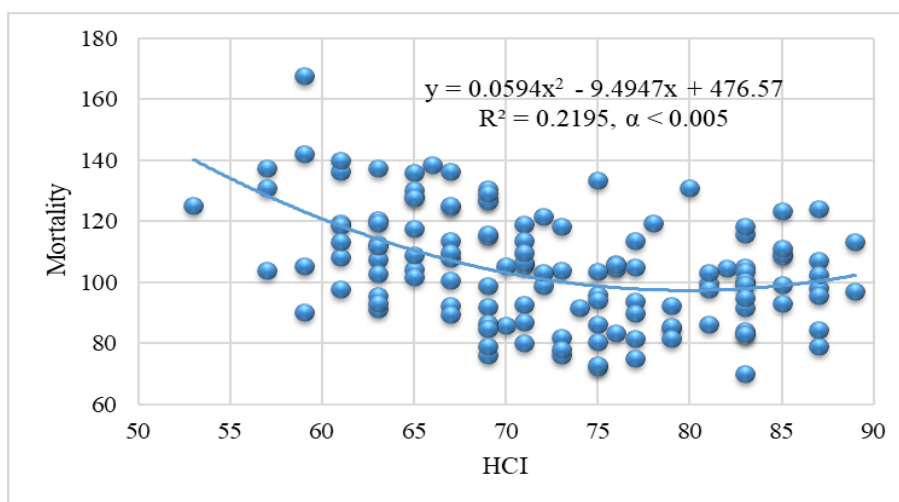


Fig.1. Connection between mean monthly values of HCI with mean monthly decade mortality by cardiovascular diseases in Tbilisi in 1980-1990.

Connection between mean monthly values of HCI with mean monthly decade mortality by cardiovascular diseases in Tbilisi (Fig. 1) has the form of a second degree polynomial. It should be noted that with an increase of the HCI values, in general, a decrease of mortality is observed. At the same time, at high values of HCI, mortality tends to increase.

This fact is also clearly demonstrated in Fig. 2. As follows from Fig. 2 in the range of HCI categories Acceptable – Very Good, mortality by cardiovascular diseases in Tbilisi decreases. With the transition of HCI categories from Very Good to Excellent - mortality increases. And although this growth is statistically insignificant, the some trend is still observed.

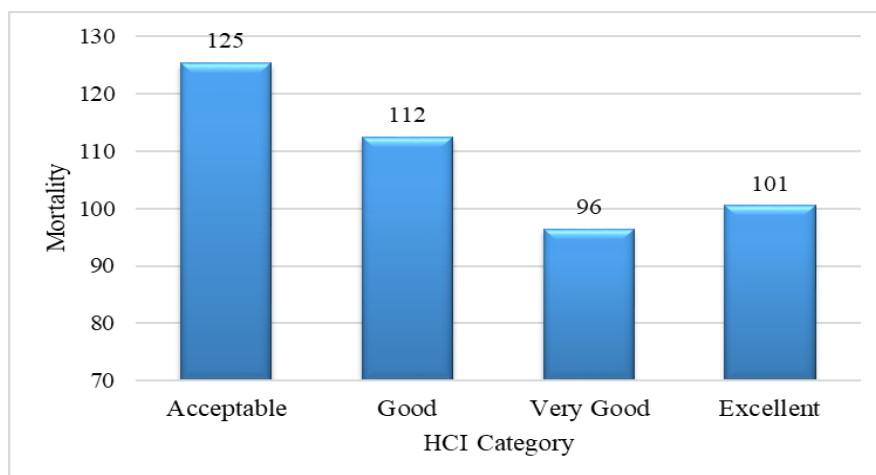


Fig.2. Mean monthly decade mortality by cardiovascular diseases at different HCI categories in Tbilisi in 1980-1990.

It should be noted that in our early studies [36, 42] it was found that the relationship between the Tourism Climate Index (TCI) and mortality by cardiovascular diseases in Tbilisi has an inverse linear relationship. At the same time, for both TCI categories Very good and Excellent, the mortality rate in Tbilisi was the same.

Table 2. Statistical characteristics of HCI components and ratings of these components in Tbilisi in 1980-1990.

Variable	TC	T	CC	A	DP	R _d	WS	W
Mean	19.2	6.2	5.9	5.6	40.1	8.9	0.9	10.0
Min	1.6	3	2.8	3	0.0	8	0.1	8
Max	34.1	10	9.0	9	158.7	10	1.9	10
Range	32.5	7.0	6.2	6.0	158.7	2	1.8	2
St Dev	9.4	2.1	1.2	1.2	33.4	0.3	0.4	0.2
Cv, %	48.8	33.3	19.5	21.8	83.4	3.6	44.2	2.5

In Table 2 statistical characteristics of HCI components and ratings of these components in Tbilisi in 1980-1990 are presented. As follows from this Table the variability of the investigated parameters is as follows. TC: – mean – 19.2, range of change – 1.6÷34.1; T: mean – 6.2, range of change – 3÷10; CC: mean – 5.9, range of change – 2.8÷9.0; A: mean – 5.6, range of change – 3÷9; DP: mean – 40.1, range of change – 0.0÷158.7; R_d: mean – 8.9, range of change – 8÷10; WS: mean – 0.9, range of change – 0.1÷1.9; W: mean – 10.0, range of change – 8÷10.

Table 3. Linear correlation coefficient between HCI components and ratings of these components with mean monthly decade mortality by cardiovascular diseases in Tbilisi in 1980-1990.

Variable	TC	T	CC	A	DP	R _d	WS	W
R(M)	-0.65	-0.36	0.25	-0.23	-0.14	0.09	-0.18	0.00
α (R(M))	<0.005	<0.005	0.005	0.01	0.10	no sign	0.05	no sign

In Table 3 information about linear correlation coefficient between HCI components and ratings of these components with mean monthly decade mortality by cardiovascular diseases in Tbilisi is presented. The largest value of the linear correlation coefficient between the values of M and the HCI components is -

0.65 (with TC), the smallest is -0.14 (with DP). The largest value of the linear correlation coefficient between the M values and the ratings of the HCI components is -0.36 (with T), the smallest is 0.0 (with W). Thus, the TC and T parameters make the main contribution to the mortality by cardiovascular diseases variability.

For example, in Fig. 3 shows a graph of the linear correlation between the values of TC and M.

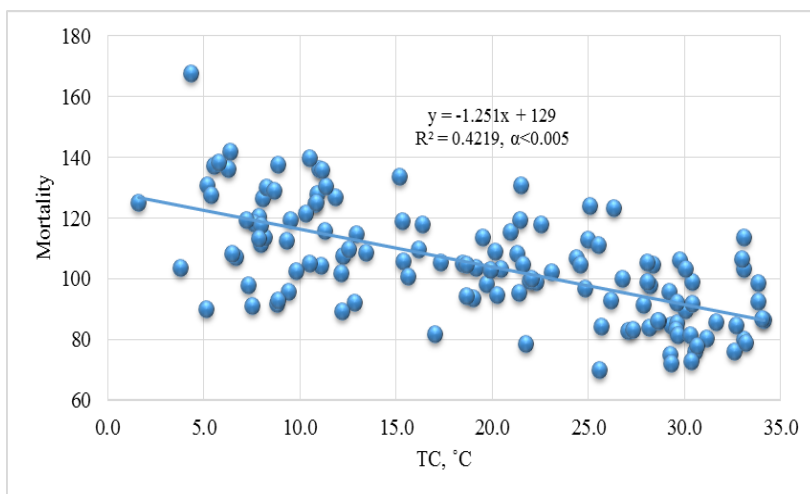


Fig.3. Connection between mean monthly values of TC with mean monthly decade mortality by cardiovascular diseases in Tbilisi in 1980-1990.

Connection between mean monthly values of T with mean monthly decade mortality by cardiovascular diseases in Tbilisi in Fig. 4 is presented. As follows from this Fig. connection between indicated parameters has the form of a second degree polynomial ($R^2 = 0.3078, \alpha < 0.005$).

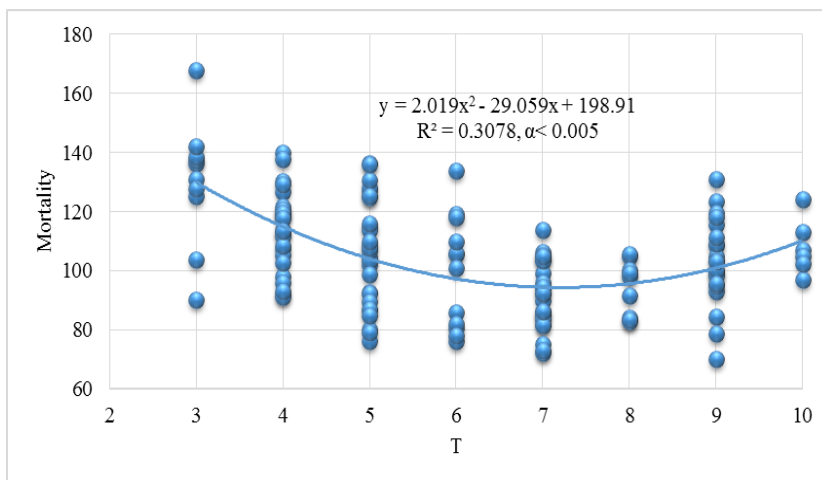


Fig.4. Connection between mean monthly values of T with mean monthly decade mortality by cardiovascular diseases in Tbilisi in 1980-1990.

Note that the second degree polynomial describes the relationship between M and T better than linear regression ($R = -0.36$, or $R^2 = 0.1239, \alpha < 0.005$, Table 3). Fig. 4 also implies that, in general, with an increase in the values of T, mortality decreases, with the exception of values of $T > 9$ (some increase in mortality). This fact determines the connection between HCI and M in the case of the HCI category "Excellent" (some increase in mortality, Fig.2).

The statistical characteristics of mean monthly values of HCI and mean monthly decade values of ambulance calls, hospitalization and total mortality in Kakheti in 2013 and 2015 in Table 4 are presented.

Table 4. Statistical characteristics of mean monthly values of HCI and mean monthly decade values of ambulance calls, hospitalization and total mortality in Kakheti in 2013 and 2015.

Variable	HCI	Ambulance Calls	Hospitalization	Total mortality
Mean	71	1753	435	15
Min	62	1184	296	8
Max	84	2265	623	26
Range	22	1081	326	18
St Dev	6.8	335.8	79.0	4.2
Cv, %	9.6	19.1	18.2	28.0

As follows from this Table values of HCI change from 62 (category Good) to 84 (category Excellent). Values of ambulance calls change from 1184 to 2265, hospitalization - from 296 to 623, total mortality - from 8 to 26.

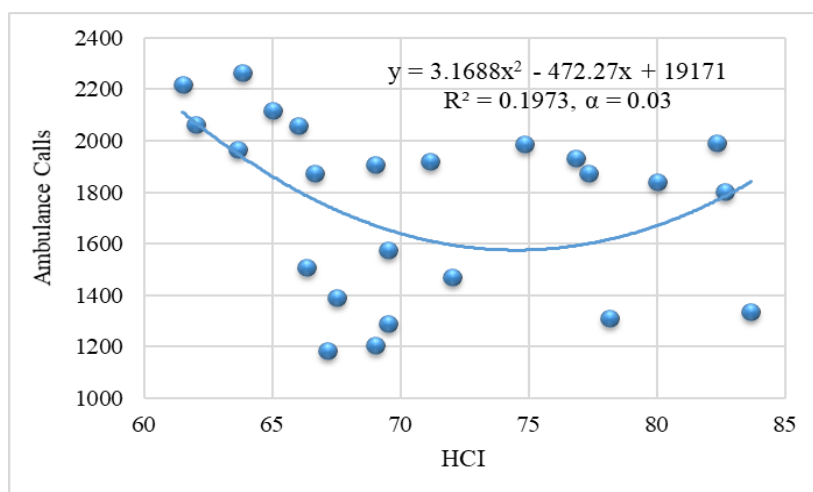


Fig.5. Connection between mean monthly values of HCI with mean monthly decade value of ambulance calls in Kakheti in 2013 and 2015.

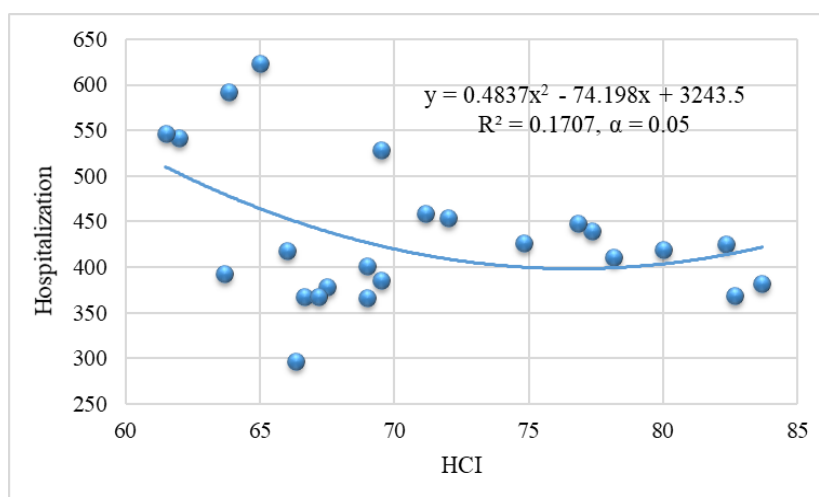


Fig.6. Connection between mean monthly values of HCI with mean monthly decade value of hospitalization in Kakheti in 2013 and 2015.

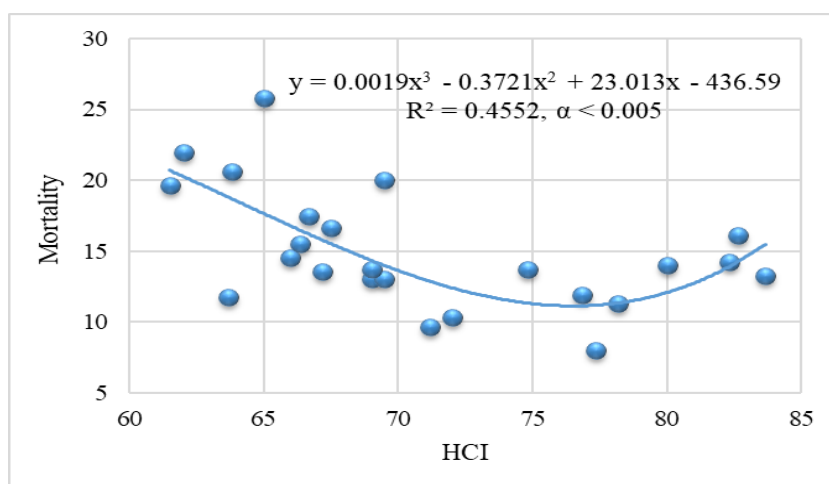


Fig.7. Connection between mean monthly values of HCI with mean monthly decade value of total mortality in Kakheti in 2013 and 2015.

In Fig. 5-7 connection between mean monthly values of HCI with mean monthly decade value of ambulance calls, hospitalization and total mortality in Kakheti are presented. As follows from these Fig. all the indicated connections have the form of a second degree polynomial (as in the case of Tbilisi, Fig.1). That is, in general with an increase of HCI values, there is a decrease of the values of AC, H and TM, and with the HCI category Excellent), their slight increase.

Table 5. Statistical characteristics of HCI components and ratings of these components in Kakheti in 2013 and 2015.

Variable	TC	T	CC	A	DP	R _d	WS	W
Mean	19.8	6.0	5.8	5.8	73.6	8.5	1.0	10
Min	6.3	3.3	3.8	3.8	17.0	6.5	0.8	10
Max	33.8	9.8	8.0	8.0	202.1	9.0	1.3	10
Range	27.5	6.5	4.2	4.2	185.1	2.5	0.5	0
St Dev	9.4	1.8	1.2	1.2	48.8	0.6	0.1	0
Cv, %	47.3	30.3	20.7	20.7	66.3	7.6	11.7	0.0

In Table 5. statistical characteristics of HCI components and ratings of these components in Kakheti in 2013 and 2015 are presented. As follows from this Table the variability of the investigated parameters is as follows. TC: – mean – 19.8, range of change – 6.3÷33.8; T: mean – 6.0, range of change – 3.3÷9.8; CC: mean – 5.8, range of change – 3.8÷8.0; A: mean – 5.8, range of change – 3.8÷8.0; DP: mean – 73.6, range of change – 17.0÷202.1; R_d: mean – 8.5, range of change – 6.5÷9.0; WS: mean – 1.0, range of change – 0.8÷1.3; W: mean – 10, range of change – 10÷10.

Table 6. Linear correlation coefficient between HCI components and ratings of these components with mean monthly decade values of ambulance calls, hospitalization and total mortality in Kakheti in 2013 and 2015.

Variable	TC	T	CC	A	DP	R _d	WS	W
R(AC)	-0.37	-0.20	-0.30	-0.30	-0.28	0.28	0.05	0.00
α (R(AC))	0.075	0.35	0.15	0.15	0.2	0.2	no sigh	no sigh
R(H)	-0.06	-0.29	-0.24	-0.24	-0.17	0.22	-0.22	0.00
α (R(H))	no sign	0.15	0.26	0.26	no sign	0.30	0.30	no sign
R(TM)	-0.57	-0.46	-0.22	-0.22	-0.25	0.26	0.35	0.00
α (R(TM))	0.005	0.02	0.30	0.30	0.24	0.22	0.1	no sigh

In Table 6 data about linear correlation coefficient between HCI components and ratings of these components with mean monthly decade values of ambulance colls, hospitalization and total mortality in Kakheti are presented.

The largest value of the linear correlation coefficient between the values of AC and the HCI components is -0.37 (with TC), the smallest is 0.05 (with WS). The largest value of the linear correlation coefficient between the AC values and the ratings of the HCI components is -0.30 (with A), the smallest is 0.0 (with W).

The largest value of the linear correlation coefficient between the values of H and the HCI components is -0.24 (with CC), the smallest is -0.06 (with TC). The largest value of the linear correlation coefficient between the H values and the ratings of the HCI components is -0.29 (with T), the smallest is 0.0 (with W).

The largest value of the linear correlation coefficient between the values of TM and the HCI components is -0.57 (with TC), the smallest is -0.22 (with CC). The largest value of the linear correlation coefficient between the TM values and the ratings of the HCI components is -0.46 (with T), the smallest is 0.0 (with W).

For example, in Fig. 8-13 shows a graphs of the connections between the values of TC and T with values of ambulance calls, hospitalization and total mortality in Kakheti.

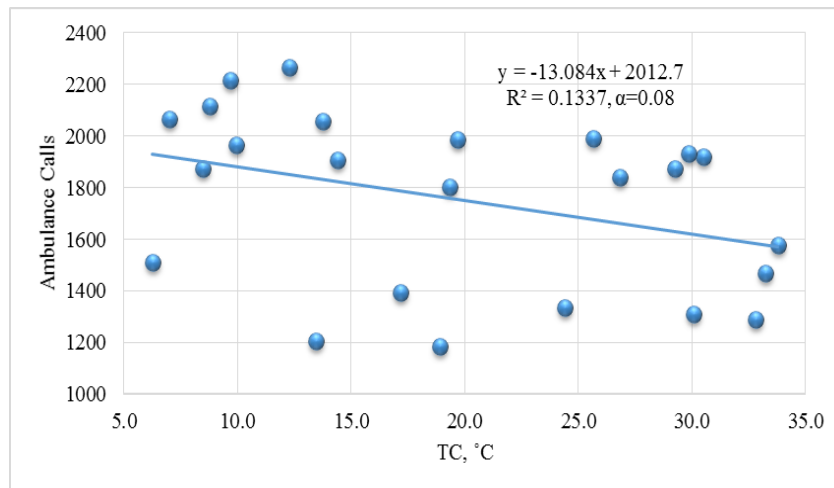


Fig.8. Connection between mean monthly values of TC with mean monthly decade value of ambulance calls in Kakheti in 2013 and 2015

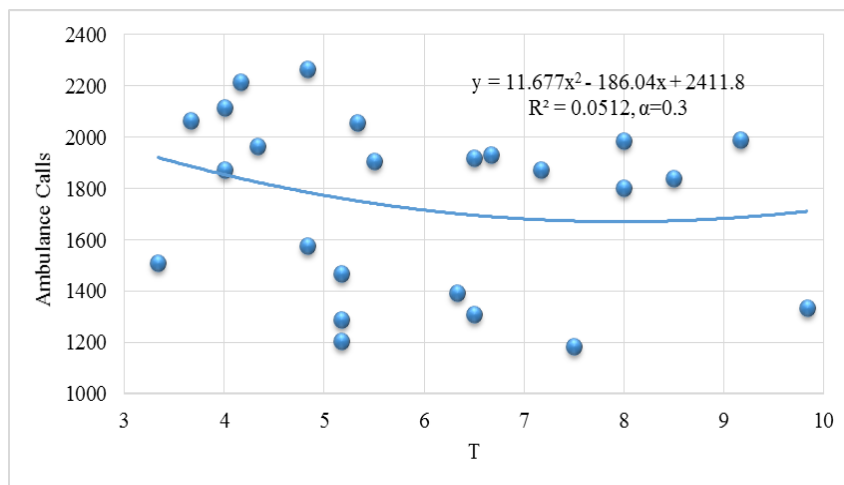


Fig.9. Connection between mean monthly values of T with mean monthly decade value of ambulance calls in Kakheti in 2013 and 2015.

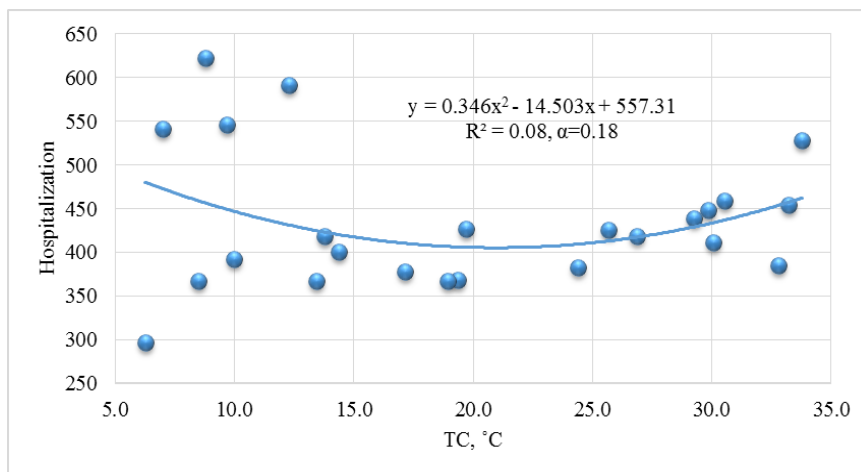


Fig.10. Connection between mean monthly values of TC with mean monthly decade value of hospitalization in Kakheti in 2013 and 2015.

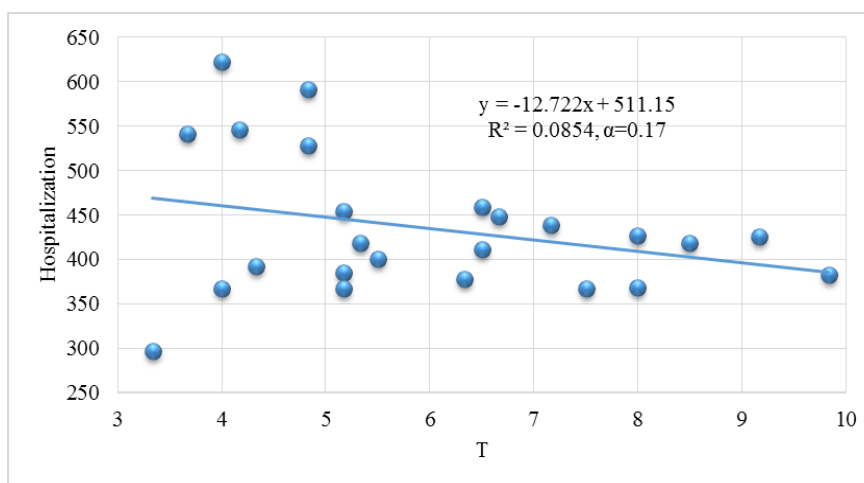


Fig.11. Connection between mean monthly values of T with mean monthly decade value of hospitalization in Kakheti in 2013 and 2015.

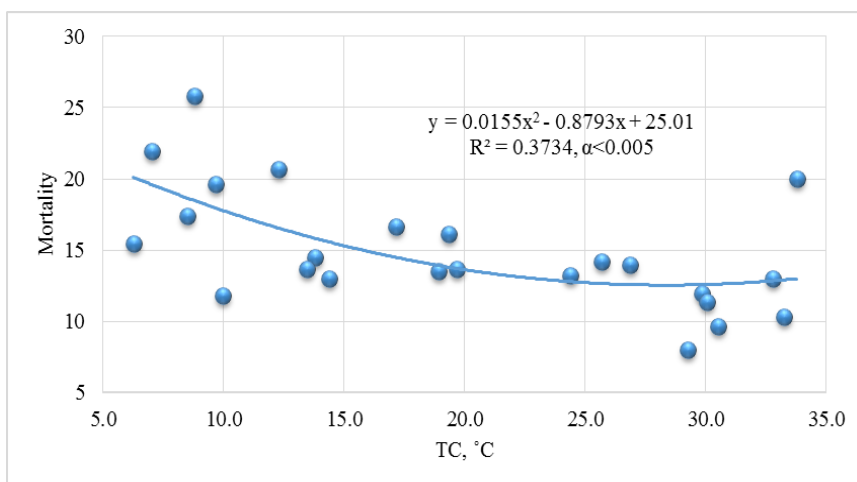


Fig.12. Connection between mean monthly values of TC with mean monthly decade value of total mortality in Kakheti in 2013 and 2015.

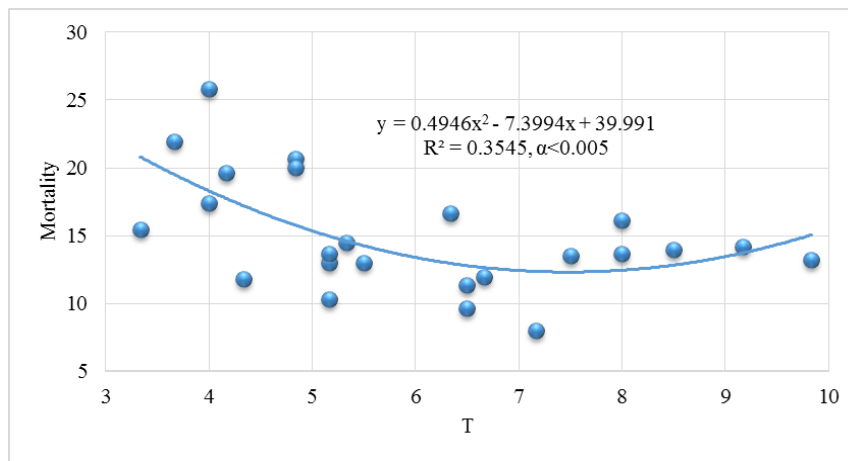


Fig.13. Connection between mean monthly values of T with mean monthly decade value of total mortality in Kakheti in 2013 and 2015.

So, the relationship between the values of TC and AC is inverse linear, and between T and AC - a polynomial of the second degree (Fig. 8 and 9); the relationship between the values of TC and H is a polynomial of the second degree, and between T and H - inverse linear (Fig. 10 and 11); the relationship between the values of TC and TM and T and TM are a polynomial of the second degree (Fig. 11 and 13).

Thus, the relationship between the HCI and the health indicators of the population in Tbilisi and Kakheti is in general similar to the classical form of mortality distribution according to the scale of thermal indices (a decrease in mortality from gradations with low uncomfortable values of the scale to comfortable ones, and then there is an increase in mortality in gradations with high uncomfortable values scale) [36].

Therefore, in our opinion, it is desirable to make adjustments in determining the degree of comfort of the HCI scale (as well as scales of other bioclimatic indices), taking into account local social and climatic conditions.

Conclusion.

In the future, we are planning to envisage conducting similar studies for other regions of Georgia.

References

- [1] Kalkstein L.S. Biometeorology – Looking at the Links between Weather, Climate and Health. WMO. Bulletin 2, 2001, v. 50, pp. 1–6.
- [2] Mc Michael A.J., Woodruff R.E., Hales S. Climate Change and Human Health: Present and Future Risks. Lancet, 367, 2006, pp. 859-868.
- [3] Golitzyn G.S., Granberg I.G., Efimenco N.P., Povolotzkaya N.P. Atmosphere and Health. Zemlya I vseleennaya, ISSN: 0044-3948, № 3, 2009, pp. 27-36, (in Russian).
- [4] Povolotzkaya N.P., Trubina M.A., Engelgardt L.T. A.L. Chizhevsky – Founder of Cosmic Ecology. International Scientific Conference „Modern Problems of Ecology“, Proceedings, ISSN 1512-1976, v. 6, Kutaisi, Georgia, 21-22 September, 2018, pp. 25-29.
- [5] Amiranashvili A., Chikhladze V., Kartvelishvili L., Khazaradze K. Expected Change of the Extremal Air Temperature and its Influence on the Mortality (Based on the Example to Tbilisi City), International Cooperation Network for East European and Central Asian Countries: EECA Conference - October 7-8, 2010, Yerevan, Armenia, <http://be.sci.am/>.
- [6] Perevedentsev Yu.P., Zandi Rahman, Aukhadeev T.R., Shantalinskii K.M. Assessment of Climate Influence on a Man in Droughty Conditions of Southwest Iran. Vestnik Udmurtskogo Universiteta, Biologia. Nauki o Zemle, T. 25, Vip. 1, 2015, pp.104-113, (in Russian).

- [7] Japaridze N., Khazaradze K. Studies in the Field of the Influence of Natural and Anthropogenic Environmental Factors on Human Health in Georgia: Current Status and Planned Works. International Scientific Conference “Natural Disasters in Georgia: Monitoring, Prevention, Mitigation”. Proceedings, ISBN 978-9941-13-899-7, Publish House of Iv. Javakhishvili Tbilisi State University, December 12-14, Tbilisi, 2019, pp. 201-204.
- [8] Matzarakis A., Cheval S., Lin T.-P., Potchter, O. Challenges in Applied Human Biometeorology. *Atmosphere* 2021, 12, 296. <https://doi.org/10.3390/atmos12030296>
- [9] Amiranashvili A., Bliadze T., Chikhladze V. Photochemical smog in Tbilisi. Monograph, Trans. of Mikheil Nodia institute of Geophysics, ISSN 1512-1135, vol. 63, Tb., 2012, 160 p., (in Georgian).
- [10] Vasin V.A., Yefimenko N.V., Granberg I.G., Povolotskaya N.P., Golitsyn G.S., Ginzburg A.S., Mkrtchyan R.I., Zherlitsina L.I., Kortunova Z.V., Maksimenkov L.O., Pogarskiy F.A., Savinykh V.V., Senik I.A., Sklyar A.P., Rubinshteyn K.G. Nekotoryye osobennosti izucheniya svyazi serdechno-sosudistykh zabolevaniy s ekologicheskimi i meteorologicheskimi faktorami na nizkogornykh kurortakh Rossii. *Vrach skoroy pomoshchi*, ISSN: 2074-742X, № 5, 2009, pp. 61-62, (in Russian).
- [11] Amiranashvili A.G., Gogua R.A., Matiashvili T.G., Kirkitadze D.D., Nodia A.G., Khazaradze K.R., Kharchilava J.F., Khurodze T.V., Chikhladze V.A. The Estimation of the Risk of Some Astro-Meteo-Geophysical Factors for the Health of the Population of the City of Tbilisi, Int. Conference “Near-Earth Astronomy 2007” Abstract, Terskol, Russia, 3-7 September 2007.
- [12] Davis R.E., Gregor G.R., Enfield K.B. Humidity: A Review and Primer on Atmospheric Moisture and Human Health. *Environmental Research*, v. 144, Part A, January 2016, pp. 106-116.
- [13] Azcárate T., Mendoza B., Levi J.R. Influence of Geomagnetic Activity and Atmospheric Pressure on Human Arterial Pressure during the Solar Cycle 24. *Advances in Space Research*, v. 58, iss. 10, 2016, pp. 2116-2125
- [14] Amiranashvili A., Amiranashvili V., Kartvelishvili L., Nodia Kh., Khurodze T. Influence of Air Effective Temperature and Geomagnetic Storms on the Population of Tbilisi City. *Trans. of the Institute of Hydrometeorology*, v. No 115, ISSN 1512-0902, Tbilisi, 2008, pp. 434 – 437, (in Russian).
- [15] Zenchenko T.A., Dimitrova S., Stoilova I., Breus T.K. Individual Responses of Arterial Pressure to Geomagnetic Activity in Practically Healthy Subjects. *Klin. Med.*, v. 87(4), 2009, pp.18–24.
- [16] Palmer S., Rycroft M., Cermack M. Solar and Geomagnetic Activity, Extremely Low Frequency Magnetic and Electric Fields and Human Health at the Earth’s Surface. *Surv. Geophys.*, v. 27, 2006, pp. 557–595. doi:10.1007/s10712-006-9010-7
- [17] Amiranashvili A.G., Cornélissen G., Amiranashvili V., Gheonjian L., Chikhladze V.A., Gogua R.A., Matiashvili T.G., Paatashvili T., Kopytenko Yu.A., Siegelova J., Dusek J., Halberg F. Circannual and circadecennian stages in mortality from cardiovascular causes in Tbilisi, Republic of Georgia (1980-1992). *Scripta medica (Brno)*, 2002, 75 pp. 255-260.
- [18] Shaposhnikov D., Revich B., Gurfinkel Yu., Naumova E. The Influence of Meteorological and Geomagnetic Factors on Acute Myocardial Infarction and Brain Stroke in Moscow, Russia. *Int. J. of Biometeorology*, v. 58, iss. 5, 2014, pp. 799–808.
- [19] Amiranashvili A.G., Bakradze T. S., Berianidze N.T., Japaridze N.D., Khazaradze K.R. Effect of Mean Annual Changeability of Air Temperature, Surface Ozone Concentration and Galactic Cosmic Rays Intensity on the Mortality of Tbilisi City Population. *Journal of the Georgian Geophysical Society, Issue B. Physics of Atmosphere, Ocean and Space Plasma*, v.19B, Tbilisi, 2016, pp. 135-143.
- [20] Slepikh V.V., Povolotskaya N.P., Korshunova Z.V., Terre N.I., Fedorov V.A. Ionization Background of the Trees and Plants of Kislovodsk Park. *Voprosy kurortologii, fizioterapii i lechnoy fizicheskoy kul'tury*, ISSN: 0042-8787, eISSN: 2309-1355, N 3, 2006, pp. 37-39, (in Russian).
- [21] Kudrinskaya T. V., Kupovykh G. V., Redin A. A. Studying the Ionization of Atmospheric Surface Layer in Different Geophysical Conditions. *Russian Meteorology and Hydrology*, April 2018, Vol. 43, Issue 4, pp. 258–263.
- [22] Hori A., Hashizume M., Tsuda Y., Tsukahara T., Nomiya T. Effects of Weather Variability and Air Pollutants on Emergency Admissions for Cardiovascular and Cerebrovascular Diseases. *Int. J. Environ Health Res.*, v. 22(5), 2012, pp.416–430. doi:10.1080/09603123.2011.650155
- [23] Amiranashvili A., Khurodze T., Shavishvili P., Beriashvili R., Iremashvili I. Dynamics of the Mortality of the Population of Tbilisi City and its Connection with the Surface Ozone Concentration. *Journ. of*

- Georgian Geophysical Soc., Iss. (B), Physics of Atmosphere, Ocean and Space Plasma, vol.16b, Tbilisi, 2013, pp. 31-38.
- [24] Lagidze L., Matchavariani L., Tsvitshivadze N., Khidasheli N., Paichadze N., Motsonelidze N., Vakhtangishvili M. Medical Aspects of Atmosphere Pollution in Tbilisi, Georgia. *Journal of Environmental Biology*, Vol.36, Special Issue, 2015, pp. 101-106.
- [25] Tkachuk S.V. The Indexes of Weather Comfort Conditions Review and their Relation to Mortality. *Proceedings of Hydrometcentre of Russia*, Vol. 347, 2012, pp. 223–245, (in Russian).
- [26] Muthers S., Laschewski G., Matzarakis A. The Summers 2003 and 2015 in South-West Germany: Heat Waves and Heat-Related Mortality in the Context of Climate Change. *Atmosphere*, November 2017, 13 p., DOI: 10.3390/atmos8110224, <https://www.researchgate.net/publication/321085363>
- [27] Ruuhela R., Jylha K., Lanki T., Tiittanen P., Matzarakis A. Biometeorological Assessment of Mortality Related to Extreme Temperatures in Helsinki Region, Finland, 1972-2014. *Int. Journ. Of Environmental Research and Public Health*, vol. 14, iss. 8., 2017, 19 p.
- [28] Ruuhela R., Votsis A., Kukkonen, J., Jylhä K., Kankaanpää S., Perrels A. Temperature-Related Mortality in Helsinki Compared to Its Surrounding Region Over Two Decades, with Special Emphasis on Intensive Heatwaves. *Atmosphere*, 12, 46, 2021, [CrossRef]
- [29] Rustemeyer N., Howells M. Excess Mortality in England during the 2019 Summer Heatwaves. *Climate*, 9, 14, 2021. <https://doi.org/10.3390/cli9010014>
- [30] Amiranashvili A.G., Japaridze N.D., Kartvelishvili L.G., Khazaradze K.R., Khazaradze R.R. Effects of Variations of the Monthly Mean Air Temperature on the Population Health of Imereti Region of Georgia. *International Scientific Conference „Modern Problems of Ecology“*, Proceedings, ISSN 1512-1976, v. 6, Kutaisi, Georgia, 21-22 September, 2018, pp. 38-41.33
- [31] Khazaradze K.R., Chkhitunidze M.S., Japaridze N.D. Effects of Variations of the Monthly Mean Max Air Temperature on the Population Health of Kakheti Region of Georgia. *Int. Sc. Conf. „Modern Problems of Ecology“*, Proc., ISSN 1512-1976, v. 7, Tbilisi-Telavi, Georgia, 26-28 September, 2020, pp. 356-359.
- [32] Steadman R.G. Norms of Apparent Temperature in Australia. *Aust. Met. Mag.*, Vol. 43, 1994, pp. 1-16.
- [33] Landsberg H.E. The Assessment of Human Bioclimate. A Limited Review of Physical Parameters. *Technical Note No 123, WMO, No 331, 1972, 37 p.*
- [34] BSR/ASHRAE Standard 55P, Thermal Environmental Conditions for Human Occupancy 2/24/03 Most Current Draft Standard, 2003, 50 p.
- [35] Tkachuk S.V. Comparative Analysis of Bioclimatic Indexes for Prediction Using a Mesoscale Model. *Uchenie Zapiski Rossiiskogo Gosudarstvennogo Gidrometeorologicheskogo Universiteta*, No 20, 2011, pp. 109-118, (in Russian), http://weatherlab.ru/sites/default/files/library/Sravn_ind.pdf
- [36] Amiranashvili A.G., Japaridze N.D., Khazaradze K.R. On the Connection of Monthly Mean of Some Simple Thermal Indices and Tourism Climate Index with the Mortality of the Population of Tbilisi City Apropos of Cardiovascular Diseases. *Journal of the Georgian Geophysical Society*, ISSN: 1512-1127, *Physics of Solid Earth, Atmosphere, Ocean and Space Plasma*, v. 21(1), Tbilisi, 2018, pp.48 -62.
- [37] Farajzadeh H., Saligheh M., Alijani B., Matzarakis A. Comparison of selected thermal indices in the northwest of Iran. *Natural Environment Change*, v. 1, N 1, 2015, pp. 1- 20.
- [38] Urban A., Kysely J. Comparison of UTCI with Other Thermal Indices in the Assessment of Heat and Cold Effects on Cardiovascular Mortality in the Czech Republic. *Int. J. Environ. Res. Public Health*, vol. 11, 2014, pp. 952-967.
- [39] Urban A., Di Napoli C., Cloke H. L., Kysely J., Pappenberger F., Sera F., Schneider R., Vicedo-Cabrera A. M., Acquaotta F., Ragetti M. S., Iniguez C., Tobias A., Indermitte E., Orru H., Jaakkola J. J. K., Ryti N. R. I., Pascal M., Huber V., Schneider A., de Donato F. et al. Evaluation of the ERA5 Reanalysis-Based Universal Thermal Climate Index on Mortality Data in Europe. *Environmental Research*, ISSN 0013-9351, vol. 198, 2021, DOI: 10.1016/j.envres.2021.111227
- [40] Shahraki F., Esmaelnejad M., Bostani M. K. Determining the Climate Calendar of Tourism in SistanBaluchestan Province, Iran. *Romanian Review of Regional Studies*, ISSN: 1841-1576, el ISSN: 2344-3707, vol. 10, Iss. 2, 2014, pp. 87-94.
- [41] Roshan G., Yousefi R., Kovács A., Matzarakis A. A Comprehensive Analysis of Physiologically Equivalent Temperature Changes of Iranian Selected Stations for the Last Half Century. *Theor. Appl. Climatol.*, ISSN: 0177-798X, eISSN 1434-4483, 2016, <https://doi.org/10.1007/s00704-016-1950-3>

- [42] Amiranashvili A.G., Japaridze N.D., Kartvelishvili L.G., Khazaradze K.R., Matzarakis A., Povolotskaya N.P., Senik I.A. Tourism Climate Index of in the Some Regions of Georgia and North Caucasus. Journal of the Georgian Geophysical Society, Issue B. Physics of Atmosphere, Ocean and Space Plasma, v. 20B, 2017, pp. 43–64.
- [43] Amiranashvili A., Danelia R., Mirianashvili K., Nodia Kh., Khazaradze K., Khurodze T., Chikhladze V. On the Applicability of the Scale of Air Equivalent- Effective Temperature in the Conditions of Tbilisi City. Trans. of Mikheil Nodia Institute of Geophysics, ISSN 1512-1135, vol. 62, Tbilisi, 2010, pp. 216-220, (in Russian).
- [44] Khazaradze K. R. Comparative Analysis of Mean-Daily Value of Air Equivalent-Effective Temperature in Tbilisi and Kojori. Journal of the Georgian Geophysical Society, Issue B. Physics of Atmosphere, Ocean and Space Plasma, v. 20B, 2017, pp. 65–72.
- [45] Amiranashvili A., Mirianashvili K., Fedorova N., Levit V., Fabiana Medeiros Carnaúba, Aliton Oliveira da Silva. Comparative Analysis of Air Equivalent - Effective Temperature in Some Cities of Georgia and Brazil, Proc. of Int. Conf. “Environment and Global Warming”, Dedicated to the 100th Birthday Anniversary of Academician F. Davitaya, Collected Papers New Series, N 3(82), ISSN 2333-3347, Tbilisi, 2011, pp. 105-110
- [46] Amiranashvili A.G., Chikhladze V.A. Saakashvili N.M., Tabidze M.Sh., Tarkhan-Mouravi I.D. Bioclimatic Characteristics of Recreational Zones – Important Component of the Passport of the Health Resort – Tourist Potential of Georgia, Pressing Problems in Hydrometeorology and Ecology, Papers of the Int. Conf. Dedicated to the 90th Anniversary of Academician G. Svanidze, September 27-29, Tbilisi, 2011, Trans. Of the Institute of Hydrometeorology at the Georgian Technical University, vol. 117, ISSN 1512-0902, Tbilisi, 2011, pp. 89-92.
- [47] Mieczkowski Z. The Tourism Climate Index: A Method for Evaluating World Climates for Tourism. The Canadian Geographer 1985, N 29, pp. 220-233.
- [48] Scott D., Ruddy M., Amelung B., Tang M. An Inter-Comparison of the Holiday Climate Index (HCI) and the Tourism Climate Index (TCI) in Europe. Atmosphere 7, 80, 2016, 17 p., doi:10.3390/atmos7060080www
- [49] Amiranashvili A., Kartvelishvili L., Matzarakis A. Comparison of the Holiday Climate Index (HCI) and the Tourism Climate Index (TCI) in Tbilisi. Int. Sc. Conf. „Modern Problems of Ecology“, Proc., ISSN 1512-1976, v. 7, Tbilisi-Telavi, Georgia, 26-28 September, 2020, pp. 424-427.
- [50] Amiranashvili A., Kartvelishvili L., Matzarakis A. Changeability of the Holiday Climate Index (HCI) in Tbilisi. Transactions of Mikheil Nodia Institute of Geophysics, ISSN 1512-1135, vol. LXXII, 2020, pp. 131-139.
- [51] Amiranashvili A.G., Kartvelishvili L.G. Holiday Climate Index in Kakheti (Georgia). Journal of the Georgian Geophysical Society, e-ISSN: 2667-9973, p-ISSN: 1512-1127, Physics of Solid Earth, Atmosphere, Ocean and Space Plasma, v. 24(1), 2021, pp. 44–62.

დასვენების კლიმატური ინდექსის კავშირი ადამიანის ჯანმრთელობასთან (თბილისის და კახეთის რეგიონის მაგალითზე)

ა. ამირანაშვილი, ა.რევიშვილი, ქ. ხაზარაძე, ნ. ჯაფარიძე

რეზიუმე

წარმოდგენილია დასვენების კლიმატური ინდექსის (დკი), მისი კომპონენტებისა და ამ კომპონენტების რეიტინგების ადამიანის ჯანმრთელობასთან კავშირის გამოკვლევა (თბილისის მაგალითზე - სიკვდილიანობა გულსისხლძარღვთა დაავადებების გამო და კახეთის რეგიონში - სასწრაფო სამედიცინო დახმარების გამოძახების, ჰოსპიტალიზაციის, საერთო სიკვდილიანობის რაოდენობა).

ნაჩვენებია, რომ მთლიანობაში ყველა მაჩვენებელი ადეკვატურად შეესაბამება ადამიანების საარსებო გარემოს კომფორტულობის ხარისხს. კერძოდ, დკი-ს კავშირს სიკვდილიანობასთან თბილისში აქვს მეორე ხარისხის პოლინომის, ხოლო კახეთში- მესამე ხარისხის პოლინომის სახე. ამავე დროს, ბიოკლიმატური კომფორტულობის ხარისხის ზრდისას დონემდე “მალთან კარგი” დაიკვირვება სიკვდილიანობის შემცირების ტენდენცია. კომფორტულობის დონეზე “შესანიშნავი” გადასვლისას აღინიშნება სიკვდილიანობის გარკვეული მატება. ანალოგიური შედეგია მიღებული დკი-ს ზოგიერთი კომპონენტისა და მათი რეიტინგებისათვის. შემოთავაზებულია მოსახლეობის ბიოკლიმატური კომფორტულობის ხარისხის სკალის განსაზღვრის კორექტირების ჩატარება ადგილობრივი სოციალური და კლიმატური პირობების გათვალისწინებით.

Связь климатического индекса отдыха со здоровьем людей (на примере Тбилиси и Кахетинского региона, Грузия)

А.Г. Амиранашвили, А. А. Ревিশвили, К. Р. Хазарадзе, Н. Д. Джапаридзе

Резюме

Представлено исследование связи Климатического Индекса Отдыха (КИО), его компонент и рейтингов этих компонент на здоровье людей (на примере Тбилиси - смертность от сердечно-сосудистых заболеваний и Кахетинского региона Грузии – количество вызовов скорой медицинской помощи, госпитализаций, общей смертности).

Показано, что в целом все показатели адекватно соответствуют степени биоклиматической комфортности среды обитания для людей. В частности, связь КИО со смертностью в Тбилиси имеет вид полинома второй степени, а в Кахетии – полинома третьей степени. При этом, с повышением степени биоклиматической комфортности до категории “Очень хорошая” наблюдается тенденция уменьшения смертности. С переходом на уровень комфортности с категорией “Превосходная” отмечается некоторый рост смертности. Аналогичный результат получен для некоторых компонент КИО и рейтингов этих компонент.

Предлагается внести корректировки в определение степени комфортности шкал биоклиматических индексов для населения с учетом местных социальных и климатических условий.

Information for contributors

Papers intended for the Journal should be submitted in two copies to the Editor-in-Chief. Papers from countries that have a member on the Editorial Board should normally be submitted through that member. The address will be found on the inside front cover.

1. Papers should be written in the concise form. Occasionally long papers, particularly those of a review nature (not exceeding 16 printed pages), will be accepted. Short reports should be written in the most concise form not exceeding 6 printed pages. It is desirable to submit a copy of paper on a diskette.
2. A brief, concise abstract in English is required at the beginning of all papers in Russian and in Georgian at the end of them.
3. Line drawings should include all relevant details. All lettering, graph lines and points on graphs should be sufficiently large and bold to permit reproduction when the diagram has been reduced to a size suitable for inclusion in the Journal.
4. Each figure must be provided with an adequate caption.
5. Figure Captions and table headings should be provided on a separate sheet.
6. Page should be 20 x 28 cm. Large or long tables should be typed on continuing sheets.
7. References should be given in the standard form to be found in this Journal.
8. All copy (including tables, references and figure captions) must be double spaced with wide margins, and all pages must be numbered consecutively.
9. Both System of units in GGS and SI are permitted in manuscript
10. Each manuscript should include the components, which should be presented in the order following as follows:
Title, name, affiliation and complete postal address of each author and dateline.
The text should be divided into sections, each with a separate heading or numbered consecutively.
Acknowledgements. Appendix. Reference.
11. The editors will supply the date of receipt of the manuscript.

CONTENTS - სარჩევი

T. Kiria, A. Esakia, M. Nikolaishvili, E. Lomadze - Magnetic Declination Statistics (Dusheti 1935-1989) and Deep Self-Learning Model თ. ქირია, ა. ესაკია, მ. ნიკოლაიშვილი, ე. ლომადე -მაგნიტური (დუშეთი 1935-1989 წ.წ.) მიხრილობის სტატისტიკა და ღრმა თვითსწავლებადი მოდელი	5 – 10
Z. Zerakidze, J. Kiria, T. Kiria - The Sharle Statistical Structure ზ. ზერაკიძე, ჯ. ქირია, თ. ქირია - შარლეს სტატისტიკური სტრუქტურები	11 –17
Z. Zerakidze, J. Kiria, T. Kiria -The Consistent Criteria for Hypotheses Testing for Sharle Statistical Structure ზ. ზერაკიძე, ჯ. ქირია, თ. ქირია - ჰიპოთეზათა შემოწმების ძალდებული კრიტერიუმები შარლეს სტატისტიკური სტრუქტურებისათვის	18 – 24
A.Amiranashvili, N. Bolashvili, Z. Gulashvili, N. Jamrshvili, N. Suknidze, Kh. Tavidashvili - Modeling the Distribution of Hailstones by Mean Max Sizes on the Territory of Kakheti (Georgia) using Data of the Freezing Level in the Atmosphere and Radar Measurements ა. ამირანაშვილი, ნ. ბოლაშვილი, ზ. გულაშვილი, ნ. ჯამრიშვილი, ნ. სუქნიძე, ხ. თავიდაშვილი - კახეთის ტერიტორიაზე საშუალო მაქსიმალური ზომების მიხედვით სეტყვის განაწილების მოდელირება ატმოსფეროში გაყინვის დონისა და რადარის გაზომვების მონაცემების გამოყენებით	25 – 36
A.Surmava, L. Gverdsiteli, L. Intskirveli, N. Gigauri - Numerical Simulation of Dust Distribution in City Tbilisi Territory in the Winter Period ა. სურმავა, ლ. გვერდწითელი, ლ. ინჭკირველი, ნ. გიგაური - ქ. თბილისის ტერიტორიაზე ზამთარში მტვრის გავრცელების რიცხვითი მოდელირება	37 – 43
A.Amiranashvili, L. Kartvelishvili - Holiday Climate Index in Kakheti (Georgia) ა.ამირანაშვილი, ლ.ქართველიშვილი - დასვენების კლიმატური ინდექსი კახეთში	44 – 62
A.Amiranashvili, A. Revishvili, K. Khazaradze, N. Japaridze - Connection of Holiday Climate Index with Public Health (on Example of Tbilisi and Kakheti Region, Georgia) ა. ამირანაშვილი, ა.რევიშვილი, კ. ხაზარაძე, ნ. ჯაფარიძე - დასვენების კლიმატური ინდექსის კავშირი ადამიანის ჯანმრთელობასთან (თბილისის და კახეთის რეგიონის მაგალითზე)	63 – 76
Information for contributors ავტორთა საყურადღებო	77 – 77

საქართველოს გეოფიზიკური საზოგადოების ჟურნალი

მყარი დედამიწის, ატმოსფეროს, ოკეანისა და კოსმოსური პლაზმის ფიზიკა

ტომი 24, № 1

ჟურნალი იბეჭდება საქართველოს გეოფიზიკური საზოგადოების პრეზიდიუმის დადგენილების
საფუძველზე

ტირაჟი 30 ცალი

JOURNAL OF THE GEORGIAN GEOPHYSICAL SOCIETY

Physics of Solid Earth, Atmosphere, Ocean and Space Plasma

Vol. 24, № 1

Printed by the decision of the Georgian Geophysical Society Board

Circulation 30 copies

ЖУРНАЛ ГРУЗИНСКОГО ГЕОФИЗИЧЕСКОГО ОБЩЕСТВА

Физика Твердой Земли, Атмосферы, Океана и Космической Плазмы

Том 24, № 1

Журнал печатается по постановлению президиума Грузинского геофизического общества

Тираж 30 экз

Tbilisi-თბილისი-Тбилиси

2021